Further re-analyses looking for effects of phylogenetic diversity on community biomass and stability

Bradley J. Cardinale*1, Patrick Venail1,2, Kevin Gross3, Todd H. Oakley4, Anita Narwani1,5, Eric Allan6, Pedro Flombaum7, Jasmin Joshi8,9, Peter B. Reich10,11, David Tilman12,13, Jasper van Ruijven14

1School of Natural Resources and Environment, University of Michigan, 440 Church Street, Ann Arbor, Michigan 48109, USA.
2Section of Earth and Environmental Sciences, Institute F.-A. Forel, University of Geneva, Versoix, Switzerland.
3Statistics Department, North Carolina State University, 2311 Stinson Drive, Raleigh, North Carolina 27695-8203, USA.
4Department of Ecology, Evolution and Marine Biology, University of California, Santa Barbara, California 93106-9620, USA.
5Aquatic Ecology, Eawag (Swiss Federal Institute of Aquatic Science and Technology), Dübendorf 8600, Switzerland.
6Institute of Plant Sciences, University of Bern, Altenbergrain 21, Bern, Switzerland.
7Centro de Investigaciones del Mar y la Atmósfera, Conicet/Universidad de Buenos Aires, C1428EGA Buenos Aires, Argentina.
8Institute of Biochemistry and Biology, Biodiversity Research/Systematic Botany, University of Potsdam, Maulbeerallee 1, 14469 Potsdam, Germany.
9Berlin-Brandenburg Institute of Advanced Biodiversity Research (BBIB), Altensteinstr 6, 14195 Berlin, Germany.
10Department of Forest Resources, University of Minnesota, 1530 Cleveland Avenue N, Saint Paul, Minnesota, USA.
11Hawkesbury Institute for the Environment, University of Western Sydney, Penrith NSW 2751, Australia.
12College of Biological Sciences, University of Minnesota, 1987 Upper Buford Circle, Saint Paul, Minnesota 55108, USA.
13Bren School of Environmental Science and Management, University of California, Santa Barbara, CA 93106, USA.
14Department of Nature Conservation and Plant Ecology, Wageningen University, Droevendaalsesteeg 3, 6708PB Wageningen University, The Netherlands.

*Corresponding author: E-mail: bradcard@umich.edu
Species richness (SR) and phylogenetic diversity (PD) are highly correlated measures of plant diversity. Each, by itself, is significantly associated with plant community biomass in biodiversity experiments. As presented by Cadotte (2015) and as we present below, reasonable but alternative analyses that attempt to control for this correlation in different ways provide contradictory or inconclusive support for the hypothesis that PD is superior to SR as a predictor of community biomass.

In Venail et al. (2015), we re-analyzed data from 16 experimental manipulations of grassland SR to look at how SR and PD influence variation of plant community biomass through time. Using four types of analyses, we showed that, after statistically controlling for variation in SR, PD was not related to community biomass or to the temporal stability of biomass. We did, however, find that SR tends to increase the biomass production of plant communities after controlling for PD.

In his comment, Cadotte expressed two concerns about our analyses. One is that we used non-random subsets of experiments, rather than the full dataset, for some of our analyses (Types 2, 3). We were clear in stating these analyses were based on non-random subsets that were specifically chosen to minimize the SR-PD correlation and avoid problems associated with multi-collinearity. We acknowledge that our tests are conservative, a cost of which is that they sacrifice statistical power while, at the same time, minimizing the chance of drawing an incorrect conclusion. But we disagree with Cadotte’s suggestion that our use of non-random data subsets led to ‘biased’ conclusions.

Cadotte’s second concern was that our analyses did not account for differences in biomass across studies. This is an important criticism to consider; we made a mistake by not controlling for variation in biomass. To address this issue, Cadotte used mixed models where study was included as a random effect, and ran analyses that standardized biomass among sites. Collectively, these led Cadotte to conclude “All analyses strongly support previous literature claims about the value of PD and I further show that: i) PD provides a more powerful explanation of variation in biomass production than species richness; ii) PD explains variation in biomass production after controlling for richness; and iii) the use of data subsets inadvertently biased the conclusions.”

We have two concerns with Cadotte’s re-analysis. First, Cadotte’s approach largely ignores the concerns we raised about multi-collinearity. When two or more predictors exhibit a high degree of correlation, each predictor contains little unique information. As a result, it is difficult (if not impossible) to estimate their independent effects using statistical methods like multivariate or partial regression (Dormann et al. 2013). The consequences of multi-collinearity include inflated error estimates that can alter conclusions about what predictors are significant or not, as well as unstable parameter estimates that can change in sign and magnitude with minor alterations to analyses (Graham 2003; Zuur, Ieno & Elphick 2010). Multi-collinearity is a concern for the dataset of Venail et al. (2015) because PD and SR are correlated with \( r = 0.90 \). We were concerned about drawing inferences from predictors that have little unique information, which is why we performed analyses that all attempted to hold one of the two predictors constant while examining the impact of the other. In contrast, Cadotte performed model selection using the full dataset where the SR-PD correlation was \( r = 0.90 \). We remain skeptical of this approach because of the difficulties generating reliable estimates for strongly correlated predictors.

A second issue with Cadotte’s analyses, which we are guilty of for some analyses in our paper, is the assumption that the relationship between biodiversity (PD or SR) and community biomass is linear.
Most studies included in the Venail et al. dataset have shown that the effect of biodiversity on community biomass is positive, but non-linear and decelerating. For example, Cardinale et al. (2011) summarized the form of diversity-biomass relationships for 433 experimental manipulations of primary producer richness and concluded "Of the studies that have shown a positive effect of producer diversity on producer biomass, 79% were best fit by some form of a positive but decelerating curve (log, power, or M-M functions, Fig. 5A)." In contrast, only 13% of studies to date are best fit by linear relationships.

We re-ran Cadotte’s analyses after accounting for non-linear relationships and found that most of his conclusions did not hold. Our modified analyses (provided in accompanying R-code) re-run the same analyses of Cadotte, which account for variation in biomass among studies, but using In-transformed predictors to also account for positive, decelerating relationships. Cadotte’s first set of analyses modeled biomass in experimental plots as linear functions of SR and/or PD with study included as a random effect to account for differences in biomass among sites. These produced an AIC of 10216 and 10194 for SR and PD, respectively, and an AIC of 10196 for a model including both SR and PD as predictors. In contrast, the best model in our modified analyses included both In-transformed SR and PD with an AIC of 10184. This represents an improved fit to data compared to Cadotte’s analyses, and confirms that failure to account for non-linear relationships led to inferior models.

After confirming that relationships between PD, SR and community biomass are better described by non-linear models, we re-ran Cadotte’s partial regression analyses which found that PD explains a significant fraction of the residual variation in biomass after accounting for effects of SR (F = 4.09, P = 0.04), but SR did not explain residual variation after accounting for effects of PD (F = 0.09, P = 0.77). Using In-transformed predictors where the SR-PD correlation was lower (r = 0.70), we found that ln(PD) explained 0.05% of the variation unaccounted for by ln(SR) (F = 3.79, P = 0.052, R² = 0.005). Yet, ln(SR) explained 1.4% of the residual variation in community biomass unaccounted for by ln(PD) (F = 12, P < 0.01, R² = 0.014).

Cadotte also re-ran our structural equations model (SEM), but used the full dataset where the PD-SR correlation was r = 0.90. He accounted for variation among studies by scaling biomass to have a mean = 0 and SD = 1. Cadotte’s SEM (reproduced in Fig. 1A) shows that PD explains a significant fraction of variation in scaled biomass and SD through time. In contrast, SR did not explain variation in either. We re-ran the same SEM on the full dataset, but using In-transformed predictors to account for non-linear relationships. The modified SEM was a significantly improved fit over the linear version (compare χ², P-values, and AIC for Figures 1A, B), and led to conclusions that were consistent with those from our original paper (Venail et al. 2015) where we found SR impacts community biomass, but PD does not. In contrast, PD affects the SD of biomass through time, but SR does not.

In his final analysis, Cadotte tried to assess whether the five experiments included in our SEM were a ‘biased’ representation of the full set of 16 experiments. He chose 1000 random subsets of five experiments and, for each subset, ran two mixed effects models – one modeling biomass as a function of PD, and one modeling biomass as a function of SR. He then calculated the difference in AIC for the two models. If ΔAIC was < 0 (> 0), this indicated PD (SR) was a better predictor of biomass for that random subset. The frequency distribution of ΔAIC values (Fig. 3 of his comment) is reproduced in Figure 1C. The mean of this distribution was significantly < 0, suggesting PD is a better predictor of biomass than SR in most random subsets of 5 experiments. In addition, the subset of 5 experiments used for our SEM was different than the overall distribution, suggesting biased selection. But conclusions about the ‘representativeness’ of the 5 experiments are overturned when we repeat the
same analyses using ln-transformed predictor variables. Indeed, the balance of evidence favored ln(SR) as the better model (Figure 1D) with the distribution of ΔAIC values being significantly > 0 (mean = +5.64, t = 12.06, P < 0.01). The value of ΔAIC for the subset of 5-experiments used in our SEM is near the center of the distribution, indicating it was not a biased subset.

So where do we stand in this exchange? Cadotte et al. (2008) found that PD was not only a significant predictor of community biomass in grassland biodiversity experiments, it explained ~2% more variation than SR. We suggested that synthesis did not control for multi-collinearity among predictors. When we (Venail et al., 2015) controlled for multi-collinearity (but failed to account for biomass differences among studies), we found PD was not a significant predictor of community biomass or stability, whereas SR was. Cadotte argued in his comment that our new analyses were incorrect because we did not account for variation in biomass among studies, and were biased by our use of data subsets to control for multi-collinearity. Cadotte’s re-analyses led him to conclude that PD is not only significant, but is again a better predictor of community biomass than SR. We responded by pointing out that multi-collinearity continues to be a concern about Cadotte’s analyses, and his conclusions don’t hold after accounting for nonlinear relationships between biodiversity and ecosystem functioning. Whether using the statistical approaches from our original paper (Venail et al. 2015) or model selection favored by Cadotte, we are led to two conclusions: (1) either SR or PD can explain most of the variation in community biomass and stability on their own because they share so much information. However, (2) when we examine their effects after statistically controlling for the other, there is little evidence that PD is a better predictor of ecological function than SR. SR is usually a significant predictor of community biomass and stability after controlling for variation in PD, whereas PD is often (though not always) non-significant after controlling for variation in SR. We would caution against interpreting these results as evidence that PD does not matter for ecosystem functioning. Cadotte is correct that experiments analyzed to date have not been explicitly designed to test hypotheses about PD, and therefore, we will need studies that orthogonally manipulate PD and SR to fully resolve their relative importance. On the other hand, given the existing data and analyses, we think it is important that researchers refrain from claiming that phylogenetic diversity is a ‘strong’ predictor of ecosystem functioning, or a ‘better’ predictor than plant richness in grasslands. Such claims are not supported at this time.


Fig 1: A comparison of results from Cadotte’s analyses, which assumed linear relationships between biodiversity and community biomass, and our analyses that account for non-linear relationships among variables. (A) Cadotte’s Structural Equation Model (SEM) suggested that phylogenetic diversity (PD) was the sole predictor of community biomass as well as the variation in biomass through time. In contrast, (B) our modified SEM using ln-transformed predictors led to a significantly improved fit to the data, and to the conclusions that species richness (SR) was a significant predictor of biomass whereas PD was not, (the same conclusions as our original paper, see Figure 4 in Venail et al. 2015). (C) Cadotte’s comparison of mixed effects models from 1000 randomly chosen subsets of 5 experiments suggests that PD is the best predictor of community biomass. In contrast, (D) our results for 1000 randomly chosen subsets of 5 experiments suggests that when ln-transformed variables are used in the mixed models to account for non-linear relationships, SR is more often the best predictor of community biomass.