- 1 Title: The population effectiveness of opportunistic chlamydia testing in primary care: a
- 2 cluster randomised controlled trial.
- 3 Professor Jane S Hocking¹ PhD, Melbourne School of Population and Global Health,
- 4 University of Melbourne, Victoria, 3010, Australia
- 5 Professor Meredith Temple-Smith DHSc, Department of General Practice, University of
- 6 Melbourne, Victoria, 3010, Australia
- Professor Rebecca Guy PhD, Kirby Institute, University of New South Wales, Sydney, New
 South Wales, 2052, Australia
- 9 Professor Basil Donovan PhD, Kirby Institute, University of New South Wales, Sydney, New
 10 South Wales, 2052, Australia
- 11 Ms Sabine Braat MSc, Melbourne School of Population and Global Health, University of
- 12 Melbourne, 3010, Victoria, Australia
- Professor Matthew Law PhD, Kirby Institute, University of New South Wales, Sydney, New
 South Wales, 2052, Australia
- Professor Jane Gunn PhD, Department of General Practice, University of Melbourne,
 Victoria, 3010, Australia
- 17 Associate Professor David Regan PhD, Kirby Institute, University of New South Wales,
- 18 Sydney, New South Wales, 2052, Australia
- Ms Alaina Vaisey MPH, Department of General Practice, University of Melbourne, Victoria,3010, Australia
- Ms Liliana Bulfone MSc, Deakin University, 221 Burwood Highway, Burwood, Victoria, 3125,
 Australia
- Professor John Kaldor PhD, Kirby Institute, University of New South Wales, Sydney, NewSouth Wales, 2052, Australia
- 25 Professor Christopher K Fairley² PhD, Central Clinical School, Monash University, 99
- 26 Commercial Rd, Melbourne, Victoria, 3004, Australia
- 27 Professor Nicola Low² MD, Institute of Social and Preventive Medicine, University of Bern,
- 28 Finkenhubelweg 11, CH-3012 Bern, Switzerland
- 29 On behalf of the ACCEPt Consortium.
- 30 1. Corresponding author
- 31 Jane Hocking, Melbourne School of Population and Global Health, University of Melbourne,
- 32 Victoria, 3010, Australia
- 33 Phone: +61 3 8344 0762
- 34 Email: jhocking@unimelb.edu.au
- 35
- 36 2. Joint senior authors
- 37 Abstract word count: 404; Main text word count: 4500

38 Abstract

39 Background

Screening young sexually active adults for genital *Chlamydia trachomatis* is promoted, but
its population effectiveness is debated. The Australian Chlamydia Control Effectiveness Pilot
(ACCEPt) investigated the effects of opportunistic chlamydia testing in primary care on
chlamydia prevalence, pelvic inflammatory disease (PID) and epididymitis in the population.
Our hypothesis was that if chlamydia testing rates increased sufficiently, the prevalence of
chlamydia in the population would decrease.

46

47 Methods

48 We conducted a cluster randomised controlled trial. Clusters were rural towns with a 49 minimum of 500 women and men aged 16-29 years resident and no more than six primary 50 care clinics. We randomly allocated each cluster using a computer-generated-minimisation 51 algorithm to receive a clinic-based chlamydia testing intervention or continue usual care. 52 Clinic staff were aware of the allocation, and posters and information cards in the waiting 53 room informed patients that the clinic was in a trial of chlamydia testing but did not specify 54 whether the clinic was intervention or control. The intervention included computerised 55 reminders, an education package, payments for chlamydia testing and feedback on testing 56 rates. The primary outcome was chlamydia prevalence, estimated before randomisation and 57 at trial end in patients aged 16-29 years attending clinics. Secondary outcomes included 58 chlamydia testing and the incidence of PID (diagnosed in clinics and hospitals) and 59 epididymitis (in clinics). Analyses were intention to treat. (Australian Clinical Trial Register 60 ACTRN12610000297022).

61

62 Findings

Between July 2010 and December 2012, we randomly assigned 26 clusters (63 clinics) to
receive a chlamydia testing intervention and 26 clusters (67 clinics) to continue with usual
care. Overall, 93,828 16 to 29 year olds attended intervention and 86,527 attended control
clinics over a mean of 3·1 years. We collected data for the final outcome measurements
between July 2014 and December 2015. The estimated chlamydia prevalence in
intervention clusters decreased from 5·0% (92/1833) to 3·4% (76/2237, difference -1·6%;
95%CI -2·9 to -0·3), and in control, from 4·6% (88/1925) to 3·4% (589/1716, difference

Page 2 of 37

- 70 -1.1%; 95%Cl -2.7 to 0.5). The odds ratio for the difference between intervention and
- 71 control clusters was 0.9 (95%CI 0.5 to 1.5). No adverse events were reported by clinics, clinic
- 72 staff or patients.
- 73

74 Interpretation

- 75 The ACCEPt results, in conjunction with evidence about the feasibility of sustained uptake of
- 76 opportunistic testing in primary care, indicate that sizeable reductions in chlamydia
- 77 prevalence or chlamydia-associated complications might not be achievable.
- 78

79 Funding

- 80 Commonwealth Department of Health, National Health and Medical Research Council,
- 81 Victorian Department of Health and NSW Ministry of Health.

82 Research in context

83 Evidence before this study

84 A systematic review of chlamydia screening interventions identified six RCTs published up to 85 the 14th February 2016, four of which investigated the effect on the incidence of PID of a 86 single offer of a chlamydia screening test and two which investigated the effect of multiple 87 rounds of chlamydia screening on chlamydia prevalence. In a meta-analysis, the incidence of 88 PID was lower in intervention than control groups (risk ratio, RR 0.68; 95%Cl 0.49 to 0.94; I^2 =8%). However, methodological limitations of the trials could have resulted in an over-89 90 estimation of the protective effects of a single chlamydia screening test. A cluster-RCT in 91 women and men in the general population in the Netherlands found no change in chlamydia 92 positivity among those tested after three rounds of screening (RR 0.96, 95% CI 0.84 to 1.09). 93 However, screening uptake was low, with only 16% screened in the first round, falling to 94 10% in the third round. A cluster RCT of a multifaceted intervention that included syndromic 95 management for sexually transmitted infections (STIs) among young adults in the 96 community and STI screening in female sex workers in Peru found no difference in 97 chlamydia prevalence after four years among young adults but in secondary analyses, found 98 a reduction among female sex workers (adjusted RR 0.72; 95% CI 0.54 to 0.98). None of the 99 trials investigated the impact of multiple rounds of testing on both chlamydia prevalence 100 and the incidence of PID. We searched PubMed from January 1 2016 to February 28 2018 101 with the terms "chlamydia" and ("randomised controlled trial" or "randomised clinical trial" or "trial" or "randomly") and restricted the search to clinical trials in English only. No further 102 103 completed trials were identified.

104

105 Added value of this study

The Australian Chlamydia Control Effectiveness Pilot (ACCEPt) is, to our knowledge, the first
RCT to investigate the impact of repeated rounds of testing on the multiple biological
outcomes of chlamydia prevalence, PID and epididymitis. It was a pragmatic trial that
reflected the situation that would occur if an opportunistic chlamydia screening programme
was rolled out. Our chlamydia testing intervention was successfully implemented in 63
clinics, reaching over 90,000 men and women aged 16 to 29 years and increasing absolute
testing rates by 11.9% (from 8.2% to 20.1%), double that achieved in the only other trial of

- 113 multiple rounds of testing in a high-income country. While it did not have a measurable
- effect on estimated prevalence of chlamydia in the population, the incidence of PID
- requiring hospitalisation decreased by 13.7 per 10,000 (95%CI: -26.9 to -0.5). However,
- there was no change in the incidence of PID or epididymitis diagnosed in the clinic.
- 117

118 Implications of all the available evidence

- 119 In high-income countries, the ACCEPt trial results, in conjunction with evidence about the
- 120 feasibility of sustained uptake of opportunistic testing in primary care clinics and evidence
- 121 from previous trials, indicate that sizeable reductions in chlamydia prevalence or chlamydia-
- 122 associated complications might not be achievable.
- 123
- 124
- 125

126 Introduction

127 Screening among young sexually active adults is widely promoted to control transmission of 128 *Chlamydia trachomatis*,¹⁻³ which causes chlamydia, the most commonly diagnosed bacterial sexually transmitted infection (STI) in high income countries.^{1,3,4} Chlamydia is usually 129 130 asymptomatic and common in young sexually active women and men, with an estimated 131 population-based prevalence of 3.1% among women and 2.3% among men aged 16 to 24 years in the UK⁵ and 4.7% among women aged 14 to 24 years in the USA.⁶ Screening for 132 asymptomatic infection aims to prevent pelvic inflammatory disease (PID), which can cause 133 134 tubal factor infertility and ectopic pregnancy. Treatment of a sufficiently large number of 135 infections should limit transmission and reduce prevalence.

136

137 The effectiveness and cost-effectiveness of chlamydia screening interventions are debated.^{7,8} In randomised controlled trials (RCTs), a single round of screening and treatment 138 139 reduced the incidence of PID, although the effect might have been overestimated.⁷ Multiple 140 rounds of screening did not reduce population chlamydia prevalence in young adults, but 141 screening coverage in one trial was very low.⁹ Given the uncertainty about the effects of 142 multiple rounds of screening on both PID and on chlamydia prevalence, additional evidence 143 is needed. The objective of the Australian Chlamydia Control Effectiveness Pilot (ACCEPt) 144 was to investigate the effect of opportunistic testing in primary care clinics on *C* trachomatis 145 prevalence, PID and epididymitis in the population. Our hypothesis was that if chlamydia 146 testing rates increased sufficiently, the population-based prevalence of chlamydia would 147 decrease.

148

149 Methods

- 150 The trial protocol trial can be found at <u>http://www.thelancet.com/protocol-reviews/12PRT-</u>
- 151 <u>9010</u>.¹⁰ We report the trial according to the CONSORT extension for cluster RCTs.

152

153 Study design

- 154 ACCEPt was a pragmatic cluster RCT targeting sexually active women and men aged 16 to 29
- 155 years for annual chlamydia testing at primary care clinics in the Australian states of New
- 156 South Wales, Victoria, South Australia and Queensland. Rural towns were chosen as clusters
- 157 because they are geographically separated to minimise contamination and all clinics in a

care clinics included general practices and Aboriginal medical services. General practitioners
(GPs) in Australia provide most primary care, including over 80% of testing for STIs. Most
general practices in Australia are businesses with varying numbers of GPs, nurses and
support staff. About 5% of towns in ACCEPt had an Aboriginal medical service. We enrolled
clinics between July 2010 and December 2012 for baseline measurements, and collected
final outcome data between July 2014 and December 2015.
The Royal Australian College of General Practitioners (RACGP) National Research and

town could be enrolled. About 25% of Australia's population lives in rural areas. Primary

- 167 Evaluation Ethics Committee and the Aboriginal Medical Research Council approved the
- 168 trial. The trial is registered with the Australian Clinical Trial Register
- 169 (<u>http://www.anzctr.org.au</u>), number ACTRN12610000297022.
- 170

158

171 *Clusters, primary care clinics, GPs and patients*

- 172 Clusters were rural towns, defined by postcode boundaries (remoteness area 2-5 in the
 173 Australian Statistical Geography Standard), with a minimum of 500 men and women aged
 174 16 to 29 years old at the 2006 census and no more than six clinics. Towns were ineligible if
- their primary economic activity was mining, military or tourism.
- 176
- 177 Within each cluster we invited all primary care clinics. If a clinic refused to participate, the 178 cluster was ineligible. GPs gave written consent to implement the intervention or control at 179 the clinic level and to provide access to their consultation and chlamydia testing data. New 180 clinics in each cluster and new GPs during the trial were recruited.
- 181
- 182 Men and women attending clinics for any consultation during the intervention period were 183 eligible for one chlamydia test per year, regardless of symptoms or contact history, if they 184 were aged 16 to 29 years and had ever had vaginal or anal intercourse. We collected de-185 identified data about consultations from electronic patient records.
- 186

187 Randomisation and masking

- 188 A statistician, located at a site away from the clusters, generated the randomisation
- 189 sequence using a computer-generated minimisation algorithm. The algorithm aimed to

190 achieve balance for three pre-specified variables: i) estimated chlamydia prevalence (<3% 191 vs. ≥3%, based on an earlier study);¹¹ ii) chlamydia testing rate (<6% vs. ≥6% based on data 192 from rural primary care clinics in 2008);¹² and iii) number of 16 to 29 year olds in the town 193 (<2000 vs. \geq 2000, the 75th percentile of population size in eligible postcodes). The 194 statistician ran the randomisation algorithm after the baseline estimation of chlamydia prevalence in each cluster and informed research staff of the allocation who in turn notified 195 196 clinics. GPs and clinic staff were aware of the allocation. Posters and information cards in 197 waiting rooms informed patients that the clinic was taking part in a trial of chlamydia testing 198 but did not specify whether the clinic was in the intervention or control. Pathology 199 laboratories conducting the testing were not informed of the allocation.

200

201 *Procedures*

202 We developed a multifaceted intervention package to encourage staff to offer annual 203 chlamydia testing to all eligible patients, based on evidence from systematic reviews and 204 considerations of long-term feasibility in Australian primary care.¹³ Guided by normalisation 205 process theory, the research team worked with each clinic to tailor the intervention to the 206 clinic and embed it into routine practice.¹⁴ The intervention package (see protocol) 207 included:¹⁰ an education package for GPs and nurses about strategies for offering testing for 208 chlamydia, management of infection and partner notification; clinical criteria for PID and 209 epididymitis diagnosis; payments for GPs (AUD\$5, \$7 or \$8 per test where <20%, 20-40% or 210 >40% of eligible patients were tested, respectively); payments for nurses (AUD\$10 for each 211 test); quarterly written feedback to GPs on their testing rates discussed in a face-to-face meeting between a research staff member and GPs; a computer alert prompting testing 212 213 with eligible patients; support to develop a reminder system to recall patients after 12 214 months if chlamydia negative or after three months if treated for chlamydia; and partner 215 notification information and resources, including access to <u>www.letthemknow.org.au</u>. 216

Research staff instructed control clinics to test for and manage chlamydia according to their usual practice. GPs received a minimal education package with clinical criteria for PID and epididymitis diagnosis to minimise measurement bias for these conditions, but they did not receive any other elements of the intervention package.

- In both intervention and control clinics, we instructed staff to test for chlamydia using
- 223 patient-collected urine specimens or vaginal swabs. Diagnosis was based on nucleic acid
- amplification tests (NAAT) conducted by the clinic's usual pathology provider. Clinics were
- instructed to treat and manage chlamydia according to Australian guidelines,¹⁵ with a single
- 226 one gram dose of azithromycin, and to initiate partner notification.
- 227

228 Trial outcomes

229 The primary outcome was chlamydia prevalence in the population served by the clinics in 230 each cluster. We assumed that the proportion of positive chlamydia tests among people 231 attending the clinic would provide an estimate of the population prevalence. We used the 232 same method to estimate the prevalence of *C* trachomatis before randomisation (survey 1) 233 and at the end of the intervention period (survey 2). See Yeung et al. for further detail on prevalence survey methods,¹⁶ but in brief, a member of the research team invited 234 235 consecutive patients in the clinic waiting room to participate, assessed eligibility, and 236 obtained written consent. People attending clinics for a consultation were eligible if they 237 were aged 16 to 29 years and had ever had vaginal or anal intercourse. The research 238 member was based at each clinic for up to six weeks to enrol the target number of 239 participants. Researchers recorded non-participants' age and gender. Participants provided 240 a self-collected specimen for testing.

241

Outcome definitions are described in the protocol.¹⁰ The primary outcome, *C trachomatis* 242 prevalence, was estimated as the proportion of women and men aged 16 to 29 years in the 243 244 surveys with a positive chlamydia test result. Secondary outcomes were measured at the 245 cluster level: incidence of PID in clinics or hospitals or epididymitis in clinics, yearly 246 chlamydia test uptake, chlamydia positivity among those tested, chlamydia re-testing 10-15 247 months after a negative test or re-testing six weeks to six months after a positive test, and repeat chlamydia diagnosis. We installed data extraction software (GRHANITE[™] 248 249 www.grhanite.com) on computers, which extracted anonymised patient information from 250 12 months before recruitment until trial end. 251

We collected PID data from two sources: participating clinics and hospitals. These data could not be combined because they were measured in different ways. At clinics, we measured 254 the incidence proportion (cumulative incidence) of PID diagnosed among women aged 16 to 255 33 years over the intervention period. The denominator was the number of female patients 256 aged 16 to 33 years with at least one consultation during the intervention period. Given the uncertainty about when PID occurs following chlamydia infection,¹⁷ the upper age limit 257 allowed for infection among women aged 26 to 29 years at start of the trial, who might have 258 259 developed PID during the intervention period. GPs were advised to use the United States Centers for Disease Control and Prevention clinical criteria to diagnose PID.¹⁸ In hospitals, 260 261 we used ICD10 codes to measure the incidence proportion of PID among women living in each cluster,¹⁹ obtained from each State health department, based on primary diagnoses of 262 263 PID. The age range for the numerator and denominator for hospitalisation data was 15 to 34 264 years because these data were only available aggregated in 5-year age groups.

265

We measured the incidence proportion of epididymitis among men aged 16 to 29 years attending clinics during the intervention. The denominator was the number of male patients aged 16 to 29 years with at least one consultation at the clinic during the intervention period. The upper age limit was 29 years because epididymitis is usually an acute condition associated with a current chlamydia infection.²⁰ The diagnosis of epididymitis was based on signs and symptoms as defined in Australian guidelines.¹⁵

272

273 Definitions of all other secondary outcomes are listed in the protocol.¹⁰

274

275 Statistical analysis

276 We based the sample size on an assumption of an absolute difference of 2% in estimated 277 chlamydia prevalence at the end of the trial (4% in control and 2% intervention clusters).^{11,21} 278 We needed 52 clusters with an average of 80 women and men aged 16 to 29 years tested 279 for chlamydia per cluster to detect a difference of this size with 80% power and 5% significance, based on an intra-class correlation coefficient (ICC) of 0.009.¹¹ The number of 280 281 patients enrolled at each clinic was proportional to the number of 16 to 29 year olds in the clinic database. This sample size allowed us to estimate chlamydia prevalence with precision 282 of ±0.5%. 283

285 We analysed trial results according to intention to treat. We used generalised linear models 286 that accounted for clustering using the generalised estimating equation approach, with 287 robust standard errors. We fitted an unadjusted model as the initial analysis for each 288 outcome. For the primary outcome, estimated chlamydia prevalence, we also fitted a 289 multivariable model to control for imbalances in cluster-level baseline covariates 290 (socioeconomic status) and potential individual level confounding factors such as gender, 291 age group, number of sex partners in the last 12 months, condom use last encounter, clinic 292 attendance for a sexual health reason and antibiotic use in the last three months. For 293 secondary outcomes, the multivariable models included gender, age group, and 294 socioeconomic status of the area served by the clinic.

295

296 We calculated the absolute difference in chlamydia prevalence between survey 1 and survey 297 2 with 95% confidence intervals (CI) for intervention and control clusters using the binomial 298 error distribution with the identity link function. We calculated the absolute difference for 299 the prevalence treatment effect, PID, and epididymitis using the identity link function. The 300 relative difference between intervention and control clusters was estimated as an odds ratio 301 (OR, prevalence, testing, retesting, repeat infection) using the binomial error distribution 302 with the logit link function or risk ratio (PID, epididymitis) using the Poisson error 303 distribution with the log link function, with 95% CI and p-values from the corresponding 304 hypothesis tests. Statistical significance was taken as a two-sided p-value less than 0.05, 305 with no adjustment for multiple comparisons. Specific models for each secondary outcome are described in the protocol.¹⁰ 306

307

308 We pre-specified exploratory analyses of the effect of the intervention in different 309 subgroups for the primary outcome (gender, age group, estimated baseline chlamydia 310 prevalence and baseline testing rate) and secondary outcomes (gender and age group). We 311 used tests for interaction to examine evidence for heterogeneity of effects between 312 intervention allocation and each subgroup. In addition, we conducted a pre-specified secondary analysis of the primary outcome to explore the effect of adherence to the 313 intervention by excluding intervention clinics that had testing rates below the 25% 314 315 percentile in the final year (10.7%) and control clinics that had testing rates above this 316 cutpoint.

- 317
- 318 We used SAS Version 9.1 for most analyses. We used STATA/SE Version 14.2 to analyse PID
- hospitalisation data and R Version 3.3.1 to obtain the intra-cluster-correlation coefficient.
- 320

321 Deviations from the protocol

There were no deviations in trial implementation. Our statistical analysis plan did not explicitly state the reporting of absolute differences but these are included as good reporting practice.

325

326 *Role of the funding source*

327 Australian Government Department of Health, National Health and Medical Research

- 328 Council, Victorian Department of Health and New South Wales Ministry of Health funded
- 329 ACCEPt. The RACGP and Australian Primary Care Nurse Association provided additional in-
- kind support. The funding bodies had no role in study design, data collection, data analysis,
- data interpretation or writing of this paper. The corresponding author had access to all data
- and responsibility for the decision to submit for publication.
- 333

334 Results

We enrolled 130 clinics in 54 postcode areas and 87% of clinics approached agreed to
participate. We merged four neighbouring towns into two clusters because of close
proximity (<40 kms) and randomised 52 clusters to intervention or control (Figure 1). The
mean distance by road between an intervention and a control cluster was 830 kms (SD 513;
range 40 to 2136 km). The intervention period in each cluster ranged from 2.5 to 4.2 years
with a mean of 3.1 years (SD 0.3). Three clinics in different control clusters (2.3%) withdrew
from the trial.

342

Baseline characteristics of clusters and GPs in intervention and control clusters were
comparable (Table 1). The response rates and characteristics of participants in surveys 1 and

- 2 were comparable between intervention and control and between the two surveys. Given
- the comparability at cluster, clinic and individual level, we report results from the
- 347 unadjusted models. Findings from the adjusted models are shown in the tables.
- 348

We enrolled 3,758 patients into survey 1 and 3,953 into survey 2. In each survey, women were more likely to participate: survey 1, 72.5% women vs 65.7% of men; survey 2, 72.4% of women vs 63.7% of men. The median ages of those who refused and those who participated was 23 years in both surveys. The intra-class correlation for chlamydia prevalence was 0.004 (SE 0.002) for survey 1 and 0.001 (SE 0.002) for survey 2.

354

355 Between surveys 1 and 2, the estimated chlamydia prevalence in intervention clusters 356 decreased from 5.0% (92/1833) to 3.4% (76/2237) (difference -1.6%, 95% CI -2.9% to 357 -0.3%) and, in control, from 4.6% (88/1925) to 3.4% (59/1716) (difference -1.1% 95% CI 358 -2.7% to 0.5%) (Table 2, Figure S1). The absolute difference in treatment effect between 359 intervention and control was estimated as -0.5 (95% CI -2.6 to 1.5; relative difference OR 360 0.9; 95%CI 0.5 to 1.5). In subgroup analyses, there was no evidence of differences in the 361 treatment effect between the intervention and control when stratified by gender, age, 362 Aboriginal status, baseline chlamydia prevalence or baseline chlamydia testing rate 363 (Supplementary Table S2). In a secondary analysis of adherence to the intervention, we 364 estimated an OR of 0.7 (95% CI 0.3 to 1.6) for the relative difference in the treatment effect 365 between intervention and control (Supplementary Table S3).

366

For PID diagnosed in clinics, the incidence estimates in intervention and control clusters
were similar, difference 5·5 per 10,000 (95% CI –13·4 to 24·3) (Table 2, Figure S2). For PID
diagnosed in hospitals, the incidence was lower in intervention (24·2 per 10,000 over 3
years follow-up) than control clusters (37·9 per 10,000 over 3 years follow-up; difference
-13·7 per 10,000, 95% CI –26·9 to –0·5). For epididymitis diagnosed in clinics, the incidence
was similar in intervention and control clusters (difference –1·6 per 10,000, 95% CI –12·4 to
9·1) (Table 2, Figure S2).

374

Among 93,828 patients aged 16 to 29 years who attended intervention clinics at least once
during the trial, 22,769 had at least one chlamydia test (67·8% had one, 20·7% had two and
11·5% had three or more tests). Among 86,527 patients attending control clinics, 14,774 had
at least one chlamydia test (70·7% had one, 19·6% had two and 9·7% had three or more
tests). The proportion of patients tested increased each year: in intervention clusters from
8·2% (95% Cl 7·0 to 9.4) to 20·1% (95% Cl 18·4 to 21·8; difference 11·9%, 95%Cl 10·3 to 13·4)

381 and in control from 8.2% (95% CI 7.2 to 9.2) to 12.9% (95% CI 11.2 to 14.5; difference 4.6%; 382 95% CI 3·3 to 6·0). The increase was greater in intervention than control clusters (OR 1·7, 383 95% CI 1·4 to 2·1, in the final year) (Figure 2A, Supplementary Table S1). Chlamydia testing 384 uptake increased in both sexes and across age groups (Figure 2B, 2C and Supplementary 385 Table S4). Chlamydia positivity decreased in both intervention and control clusters and, in 386 the final year, was lower in intervention clusters (6.8%, 95% Cl 5.6 to 8.0) than control 387 (8.8%, 95% CI 7.7 to 9.9; OR 0.8; 95% CI 0.6 to 1.0) (Supplementary Table S1). Chlamydia 388 positivity was higher in men than women at all time points and highest among 16 to 19 year 389 olds (Supplementary Table S5).

390

391 The proportion of patients who had a repeat test within 10 to 15 months of a negative 392 chlamydia test result was higher in intervention than control clusters throughout the 393 intervention period (Supplementary Table S1). The highest proportion was 10.4% (95% CI 394 8.7 to 12.1) after year one in intervention clusters and declined to below 4.0% in both 395 groups by the final year. Among patients with a positive chlamydia test, the proportion with 396 a repeat test within six weeks to six months was higher in intervention than control 397 throughout the intervention period (Supplementary Table S1). The odds of a repeat positive 398 chlamydia test after retesting was similar in intervention and control clusters during the 399 intervention period (Supplementary Table S1).

400

401 No adverse events were reported by clinics, clinic staff or patients

402

403 Discussion

We implemented a pragmatic multifaceted chlamydia testing intervention in 63 primary
care clinics, reaching over 90,000 men and women aged 16 to 29 years, with an absolute
increase in testing of 11.9% (to 20.1%) over a mean of 3.1 years follow-up. The estimated
prevalence of chlamydia declined in both intervention (absolute difference -1.6%) and
control clusters (absolute difference -1.1%). The incidence of PID-related hospitalisations
was 40% lower (absolute difference -13.7 per 10,000) in intervention than control clusters,
but the incidence of PID and epididymitis diagnosed in clinics were similar.

412 ACCEPt has several strengths. First, it is the only RCT to have evaluated the effects of several rounds of opportunistic chlamydia screening on multiple outcomes of chlamydia prevalence, 413 414 PID and epididymitis. Second, ACCEPt shows the value of randomisation. In a before-and-415 after design, without a control arm, the reduction in estimated chlamydia prevalence would 416 have been assumed to have resulted from increased testing. Third, the pragmatic 417 implementation and evaluation reflect the situation that would occur if a chlamydia 418 screening programme was rolled out, with the addition of theory-based implementation into routine practice.¹⁴ Fourth, cluster randomisation meant that all clinics in a town were 419 420 offered the same intervention package. Measurement of an infectious disease outcome at 421 the cluster level also captured the effects on transmission within social and sexual networks. 422 To our knowledge, this trial provides the least biased estimate to date of the effectiveness 423 of an intervention to increase opportunistic chlamydia testing. Measures to minimise bias 424 included enrolment of all clinics in a cluster to reduce contamination, allocation 425 concealment, use of routine data to minimise measurement bias and blinding of the 426 statistician in the analysis.

427

428 The trial also had limitations. First, we estimated chlamydia prevalence in the population by 429 conducting a survey in attenders at clinics, and only about 30% of those who took part were 430 men, reflecting clinic attendance patterns. To increase external validity, we enrolled 431 consecutive patients, irrespective of the reason for consultation; response rates were high, 432 about 70% in each survey and measured characteristics of participants in both surveys were similar. Second, our sample size assumed an absolute reduction in estimated chlamydia 433 434 prevalence of 2% in the intervention and no change in the control. We did not anticipate a 435 decrease in both groups. Nevertheless, we can rule out baseline differences as a source of 436 uncertainty because we estimated chlamydia prevalence at the start and end of the trial and 437 there were no important baseline imbalances between groups. Third, the trial was 438 conducted in rural towns so the results might not be generalisable to more ethnically and 439 socio-economically diverse urban areas. However, we accounted for area level socio-440 economic factors in our analysis. Fourth, we could not eliminate ascertainment bias from 441 the diagnosis of PID within clinics because clinical judgment can be influenced by knowledge 442 of a woman's chlamydia test status. Fifth, we were unable to collect data on partner 443 notification nor confirm treatment of all diagnosed cases, but, data from primary care in

Australia suggest nearly all cases are treated.²² Finally, as our intervention was pragmatic
and modified as clinics' needs changed, we could not determine which intervention
components were the most effective at increasing test uptake.

447

ACCEPt is one of the few controlled trials to have evaluated the effects of chlamydia testing 448 449 on prevalence. The Dutch Chlamydia Screening Implementation (CSI) trial evaluated yearly 450 register-based invitations to 16 to 29 year old women and men using self-collected 451 specimens.⁹ Uptake was too low to estimate chlamydia prevalence in the population, so the 452 investigators reported the proportion of chlamydia tests with a positive result (chlamydia 453 positivity) and found no difference between intervention and control areas, with poor uptake as a possible explanation. Baseline chlamydia positivity in ACCEPt (4.8%, 180/3758) 454 455 and the first CSI screening round (4.3%, 1851/43358) were similar. Unlike CSI, chlamydia 456 positivity in ACCEPt declined over time. This was not unexpected because, as testing rates 457 increase, the inclusion of lower risk individuals will decrease chlamydia positivity. The 458 reduction in the incidence of PID hospitalisation (RR 0.6, 95% CI 0.4 to 1.0) in ACCEPt was 459 compatible with the relative risk in a meta-analysis of previous trials.⁷ However, the 460 incidence of PID was low and the absolute difference was small (13.7 per 10,000 after a 461 mean of 3.1 years). One trial in Denmark that used routine data to estimate PID incidence 462 from hospital and community records, found comparably low rates of PID (around 50 per 463 10,000 women) but no difference between control and intervention (hazard ratio 1.12, 95% CI 0.7 to 1.8).²³ Neither ACCEPt nor the Danish trial found an effect of chlamydia testing on 464 465 epididymitis in men.

466

467 The ACCEPt intervention did not result in a clinically relevant difference in estimated 468 chlamydia prevalence between intervention and control clusters after three years. An 469 important reason might be that testing uptake needed to be sustained at higher levels for 470 longer, with higher levels of repeat testing. Our mathematical modelling suggested that 471 uptake of 20% by women and men under 30 years could reduce population prevalence by 65% in 10 years.²¹ Our intervention increased uptake to an average of 20%, which may have 472 473 increased further over time. Larger financial incentives might have increased testing rates 474 more, but at the time of the trial, our incentive payments were based on what was offered 475 for immunisation and other similar preventive activities in general practice. A one-year pilot 476 for the National Chlamydia Screening Programme in England achieved uptake of 45% with 477 GP incentives of up to £25 per test.²⁴ These payments were not sustained during the roll-out 478 with testing rates in general practice falling to below 10%.²⁵ The effects of higher levels of 479 test uptake on population prevalence of chlamydia remain unclear. In Great Britain, 480 population chlamydia prevalence in women aged 18 to 24 years was 3.1% in 1999 to 2000 and 3.2% in 2010 to 2011,⁵ during which testing coverage increased from 8% in 2008 to 481 about 30% in 2011.²⁵ In the United States of America, population chlamydia prevalence 482 483 among women aged 15 to 24 years was 4.1% (95% CI 2.4 to 6.8) in 1999 to 2000 and 3.8% (95% CI 2·4 to 6.0) in 2007 to 2008,²⁶ when testing coverage among women aged 15 to 24 484 485 years was reported to be more than 35% per year.²⁷

486

487 Reasons for the reduction in estimated chlamydia prevalence in both intervention and 488 control clusters, despite marked differences in test uptake, are unclear. We do not think 489 that awareness about the intervention or information given to control clusters contributed 490 to the reduction in prevalence for two reasons. First, only GPs in intervention clusters 491 received education about chlamydia testing and management and the PID and epididymitis 492 package included only information about criteria for diagnosis. Second, testing uptake did 493 not increase after survey 1. It is possible that treatment of chlamydia detected during the 494 first survey removed prevalent infections from the population, with insufficient time 495 between the surveys for rebound. The geographical separation of rural towns in Australia 496 might have reduced the opportunity for new infections to be introduced and, in survey 2, 497 78% of participants reported that their most recent sexual partner came from within the 498 same postcode. However, the surveys only included about 10% of all patients registered at 499 the clinics, so a marked reduction in *C* trachomatis transmission at the population level is 500 unlikely. Background antibiotic use may have affected both intervention and control, 501 particularly as the second survey took place during a severe influenza season.²⁸ We think that antibiotic use is an unlikely explanation because prescribing increased by only 5.6%502 across Australia between surveys,²⁹ chlamydia notifications in the population as a whole did 503 504 not decline,⁴ and statistical adjustment for reported antibiotic use in both surveys did not 505 affect our results. Regression to the mean, affecting both intervention and control, is the 506 most likely explanation. It is also possible that participants in survey 1 were at higher risk of

507 chlamydia than those in survey 2, even though participation rates and measured508 characteristics were similar.

509

510 The ACCEPt intervention was an intensive but highly pragmatic package that can be adapted 511 to different primary care clinics. Although it increased levels of opportunistic chlamydia 512 testing, it is likely that substantial investment would be required to increase and maintain 513 test coverage at higher levels. An ongoing economic evaluation will determine its cost-514 effectiveness. In high income countries with a low general population chlamydia prevalence, 515 public health specialists and policy makers should decide on an acceptable level for 516 chlamydia control and focus on the reduction of social and ethnic inequalities in chlamydia and its associated complications.³⁰ The ACCEPt results, in conjunction with evidence about 517 518 the feasibility of sustained uptake of opportunistic testing in primary care clinics, indicate 519 that substantial reductions in chlamydia prevalence or chlamydia-associated complications 520 might not be achievable.

521

522 ACCEPt Consortium:

523 Jane S Hocking (principal investigator), University of Melbourne; Christopher K Fairley, 524 Monash University; John Kaldor, Kirby Institute; Basil Donovan, Kirby Institute; Matthew 525 Law, Kirby Institute; Nicola Low, University of Bern; Meredith Temple-Smith, University of 526 Melbourne; Lena Sanci, University of Melbourne; Jane Gunn, University of Melbourne; 527 David Regan, Kirby Institute; David Wilson, Kirby Institute; Margaret Hellard, Burnet 528 Institute; Marian Pitts, Latrobe University; Anne Mitchell, Latrobe University; John Imrie, 529 University of Witwatersrand; Sepehr Tabrizi, Royal Women's Hospital; James Ward, South 530 Australian Health and Medical Research Institute; Robert Carter, Deakin University; Marion 531 Saville, Victorian Cytology Service; Marcus Chen, Melbourne Sexual Health Centre; Marie 532 Pirotta, University of Melbourne.

533

534 Author contributions:

JSH, MTS, RG, BD, ML, JG, DR, JK, CF, NL secured the funding for the trial. JSH was involved
in the conception and design of the study, data analysis and data interpretation and for

537 monitoring any reported adverse events or breaches of protocol. MTS, RG, BD, JG, JK, CKF

- and NL were involved in the conception and design of the study and data interpretation. ML
- 539 was involved in the conception and design of the study, data analysis and data
- 540 interpretation. SB was involved in the data analysis and interpretation. AV was involved in
- 541 data collection and enrolment of clinics, GPs and patients. LB and DR were involved in study
- 542 design and data collection. All authors contributed to the writing of this manuscript and
- 543 approved the final version for submission.
- 544

545 **Declaration of interests:**

- 546 All authors declare no competing interests.
- 547

548 Acknowledgements:

549 We thank: Jane Tomnay, from the Centre for Excellence in Rural Sexual Health, University of

550 Melbourne, for providing resources, support and advice on conducting this study in rural

areas; Dougie Boyle and his team for GRHANITE technical support; Carolyn Murray and Chris

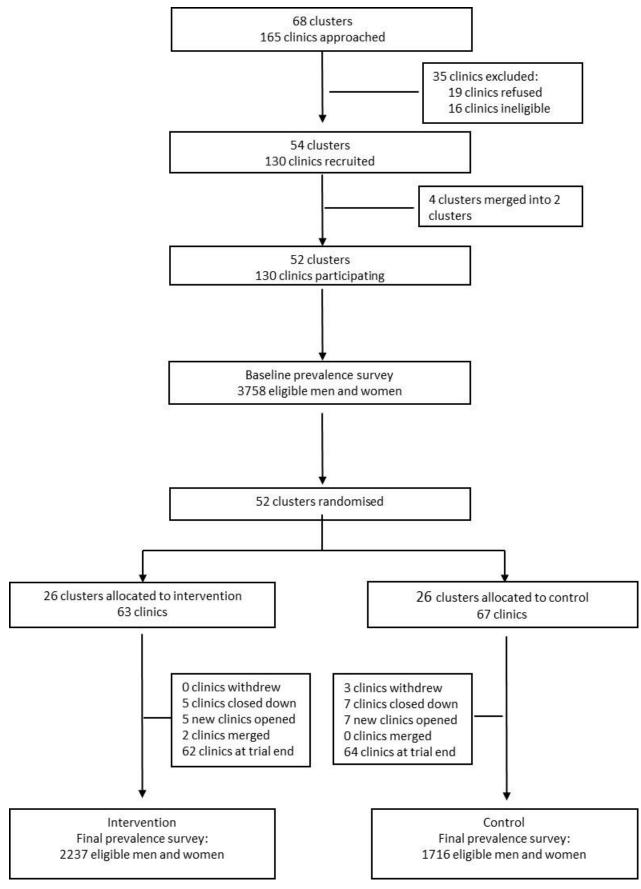
- 552 Bourne, from the Sexually Transmissible Infections Program Unit, NSW Health, for providing
- advice and resources; Simone Spark for her contribution with the set-up and ongoing
- conduct of the trial. ACCEPt was supported by a team of research officers, over 80 research
- assistants, and pathology providers who are detailed on www.accept.org.au.

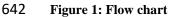
557 References

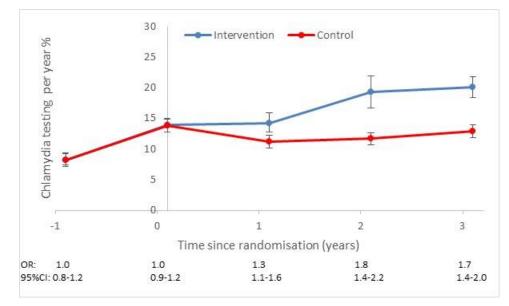
558

559 1. Centers for Disease Control and Prevention. Sexually Transmitted Disease Surveillance 2015. Atlanta: U.S. Department of Health and Human Services; 2016. 560 561 Royal Australian College of General Practitioners. Guidelines for preventive activities 2. 562 in general practice ("The Red Book"). 9th ed. Melbourne. East Melbourne: Royal Australian 563 College of General Practitioners; 2016. 564 Public Health England. Sexually transmitted infections and chlamydia screening in 3. 565 England, 2017. Health Protection Report 2018; 12(20). 566 Kirby Institute. HIV/AIDS, viral hepatitis and sexually transmissible infections in 4. 567 Australia Annual Surveillance Report 2016. Sydney: Kirby Institute, UNSW, 2016. 568 5. Sonnenberg P, Clifton S, Beddows S, et al. Prevalence, risk factors, and uptake of 569 interventions for sexually transmitted infections in Britain: findings from the National 570 Surveys of Sexual Attitudes and Lifestyles (Natsal). Lancet 2013; 382(9907): 1795-806. 571 6. Torrone E, Papp J, Weinstock H, Centers for Disease C, Prevention. Prevalence of Chlamydia trachomatis genital infection among persons aged 14-39 years--United States, 572 573 2007-2012. MMWR 2014; 63(38): 834-8. Low N, Redmond S, Uusküla A, et al. Screening for genital chlamydia infection. 574 7. 575 Cochrane Database of Systematic Reviews 2016; (8): Art. No: CD010866. 576 8. de Wit GA, Over EA, Schmid BV, et al. Chlamydia screening is not cost-effective at 577 low participation rates: evidence from a repeated register-based implementation study in 578 The Netherlands. Sexually Transmitted Infections 2015; 91(6): 423-9. 579 van den Broek IVF, van Bergen JEAM, Brouwers EEHG, et al. Effectiveness of yearly, 9. 580 register based screening for chlamydia in the Netherlands: controlled trial with randomised 581 stepped wedge implementation. BMJ 2012; 345: e4316. 582 10. Hocking JS, Low N, Guy R, et al. 12 PRT 09010: Australian Chlamydia Control 583 Effectiveness Pilot (ACCEPt): a cluster randomised controlled trial of chlamydia testing in general practice (ACTRN1260000297022). Lancet 2013; Accepted protocol 584 585 summary(http://www.thelancet.com/protocol-reviews-list). Walker J, Fairley C, Bradshaw C, et al. The difference in determinants of Chlamydia 586 11. 587 trachomatis and Mycoplasma genitalium in a sample of young Australian women. BMC 588 Public Health 2011; 11: 35. 589 Reynolds R, Oakman T. Genital chlamydia in southern New South Wales: an 12. 590 ecological analysis of testing and notification patterns 2004-2008. Australian Journal Rural 591 Health 2010; 18(4): 159-65. 592 13. Guy RJ, Ali H, Liu B, et al. Efficacy of interventions to increase the uptake of 593 chlamydia screening in primary care: a systematic review. BMC Infectious Diseases 2011; 11: 594 211. 595 14. Murray E, Treweek S, Pope C, et al. Normalisation process theory: a framework for 596 developing, evaluating and implementing complex interventions. BMC Medicine 2010; 8(63). 597 Australasian Sexual Health Alliance. Australian STI management guidelines for use in 15. 598 primary care. September 2014 2014. http://www.sti.guidelines.org.au/sexually-599 transmissible-infections/chlamydia#management (accessed 20 October 2014. 600 16. Yeung AH, Temple-Smith M, Fairley CK, et al. Chlamydia prevalence in young 601 attenders of rural and regional primary care services in Australia: a cross-sectional survey. 602 Med J Aust 2014; 200(3): 170-5.

- 603 17. Herzog SA, Althaus CL, Heijne JC, et al. Timing of progression from Chlamydia 604 trachomatis infection to pelvic inflammatory disease: a mathematical modelling study. BMC 605 Infectious Diseases 2012; 12: 187. 606 Workowski KA, Bolan GA. Sexually Transmitted Diseases Treatment Guidelines, 2015. 18. 607 MMWR 2015; 64(3): 1-135. 608 Bender N, Herrmann B, Andersen B, et al. Chlamydia infection, pelvic inflammatory 19. disease, ectopic pregnancy and infertility: cross-national study. Sexually Transmitted 609 610 Infections 2011; 87(7): 601-8. 611 Trojian TH, Lishnak TS, Heiman D. Epididymitis and orchitis: an overview. American 20. 612 Family Physician 2009; **79**(7): 583-7. Regan D, Wilson D, Hocking J. Coverage is the key for effective screening of 613 21. 614 Chlamydia trachomatis in Australia. Journal of Infectious Diseases 2008; 198(3): 349-58. 615 22. Chen M, Ryder N, Donovan B. Completeness and timeliness of treatment for 616 Chlamydia within a sexual health service. International Journal of STD & AIDS 2004; 15: 762-617 4. 618 23. Andersen B, van Valkengoed I, Sokolowski I, Moller JK, Ostergaard L, Olesen F. 619 Impact of intensified testing for urogenital Chlamydia trachomatis infections: a randomised 620 study with 9-year follow-up. Sexually Transmitted Infections 2011; 87(2): 156-61. 621 24. Pimenta JM, Catchpole M, Rogers PA, et al. Opportunistic screening for genital 622 chlamydial infection. 11: Prevalence among healthcare attenders, outcome and evaluation 623 of positive cases. Sexually Transmitted Infections 2003; 79: 22-7. 624 25. Chandra NL, Soldan K, Dangerfield C, et al. Filling in the gaps: estimating numbers of 625 chlamydia tests and diagnoses by age group and sex before and during the implementation 626 of the English National Screening Programme, 2000 to 2012. Euro surveillance 2017; 22(5). 627 26. Datta SD, Torrone E, Kruszon-Moran D, et al. Chlamydia trachomatis trends in the 628 United States among persons 14 to 39 years of age, 1999-2008. Sexually Transmitted 629 Diseases 2012; 39(2): 92-6. 630 27. Hoover KW, Leichliter JS, Torrone EA, et al. Chlamydia screening among females aged 15-21 years--multiple data sources, United States, 1999-2010. MMWR supplements 631 632 2014; 63(2): 80-8. 633 28. Department of Health. Australian Influenza Surveillance Report Reporting Period: 26 634 September to 9 October 2015. Australian Government; 2015. 635 Department of Health. Expenditure and Prescriptions twelve months to 30 June 29. 636 2015. 2016. http://www.pbs.gov.au/info/statistics/pbs-expenditure-prescriptions-30-june-637 2015 (accessed 31 December 2017 2017). 638 30. Unemo M, Bradshaw CS, Hocking JS, et al. Sexually transmitted infections: challenges 639 ahead. Lancet Infect Dis 2017.
- 640



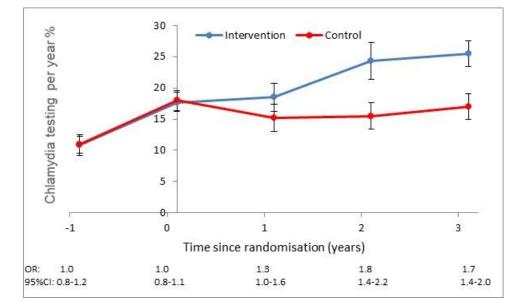






A)

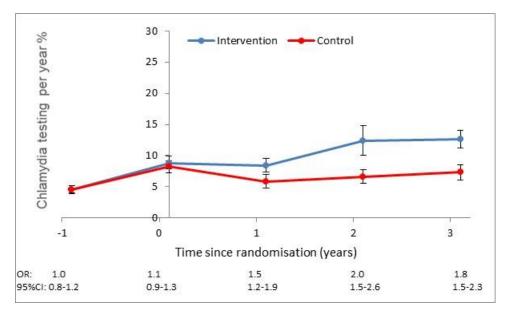






647

B)



649

650

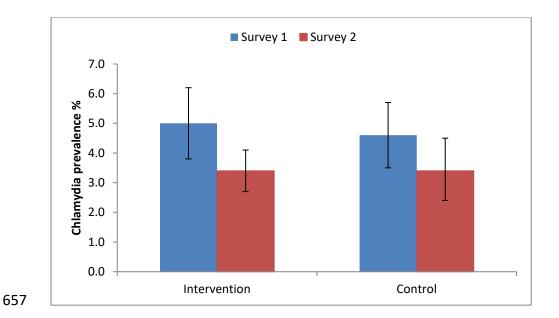
C)

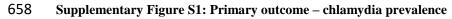
651 Figure 2: Chlamydia testing per 100 per year by time since randomisation – A) Overall, B) Females, C)

652 Males. (Time since randomisation refers to time since the start of the intervention period in each cluster and is

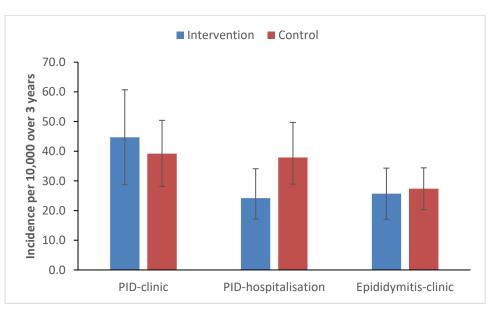
defined as: -1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before

randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2
as (13-24) months after randomisation; 3 as (25-36) months after randomisation.





659



660

661 Supplementary Figure S2: Secondary morbidity outcomes (PID-clinic – incidence of PID diagnosed at participating clinics; PID-hospitalisation – incidence of PID associated hospitalisations; Epididymitis-clinic –

663 incidence of epididymitis diagnosed at participating clinics)

Table 1: Baseline characteristics of clusters and clinics and characteristics of those participating in each prevalence survey

revalence survey Variable	Intervention	Control	Australian average§
Number of clusters	26	26	-
Number of clinics at randomisation	63	67	-
Number of GPs*	305	281	-
GPs who are female n (%)*	120 (39)	112 (40)	41%
GP age group (n, %)*			
≤ 44 yrs.	149 (49)	145 (52)	33%
45-59 yrs.	131 (43)	110 (39)	45%
≥ 60 yrs.	25 (8)	26 (9)	22%
Socioeconomic status of clusters n (%)†	20 (0)	===(>)	/0
Relatively most disadvantaged Q1	9 (35)	7 (27)	
Q2	14 (54)	14 (54)	-
Q3	2(8)	5 (19)	
Q3 04	1 (4)	0(0)	
Relatively least disadvantaged Q5	0 (0)	0(0)	
Remoteness area of clusters n (%):	0(0)	0(0)	
Inner Regional	14 (54)	12 (46)	_
Outer Regional	11 (42)	12 (46)	
Remote	1 (9)	2 (8)	
Total population 16 to 29 year olds§	1 (2)	2 (0)	
· · · · ·	10288 (34)	11924 (36)	
16-19 yrs n(%)	9912 (32)	10407 (32)	
20-24 yrs n(%)	10327 (34)	10480 (32)	
25-29 yrs n(%)	14798 (49)	15961 (49)	
Females n(%)			
Chlamydia testing in the 12 months prior to	2802/34143	3107/37775	-
recruitment n/N, % (95%CI)	8.2 (7.0 to 9.4)	8.2 (7.2 to 9.2)	
Chlamydia prevalence prior to randomisation	92/1833	88/1925	-
n/N, % (95%CI) ¶	5.0 (3.8 to 6.2)	4.6 (3.5 to 5.7)	
	Prevalence survey 1	1025	
No of participants in the analysis	1833	1925	-
Response rate % (95%CI)	$68 \cdot 8 (61 \cdot 7 \text{ to } 75 \cdot 1)$	71.8 (66.9 to 76.3)	
Females n (%)	1276 (70)	1394 (72)	-
Age group n (%)	122 (24)	521 (27)	
16-19 yrs.	432 (24)	521 (27)	-
20-24 yrs.	714 (39)	697 (36)	
25-29 yrs.	687 (37)	707 (37)	
No. of partners last 12 months n (%)	100 (11)	100 (0)	
Missing	199 (11)	166 (9)	-
0/1	1109 (68)	1201 (68)	
2	253 (15)	267 (15)	
3+	272 (17)	291 (17)	
	Prevalence survey 2	1717	
No of participants in the analysis	2237	1716	-
Response rate % (95%CI)	72 (67.4 to 76.5)	67 (61.0 to 72.0)	
Females n (%)	1616 (72)	1229 (72)	-
Age group n (%)	570 (25)	101 (25)	
16-19 yrs.	570 (25)	424 (25)	-
20-24 yrs.	834 (37)	696 (41)	
25-29 yrs.	833 (37)	596 (35)	
No. of partners last 12 months n (%)			
Missing	182 (8)	119 (7)	-
0/1	1458 (71)	1104 (69)	
2	255 (12)	204 (13)	
3+	342 (17)	289 (18)	

667 CI = confidence interval. NB: Not all percentages add up to 100% because of rounding.

668 * Based on the number of GPs recruited at baseline. General practice is dynamic with GPs departing and/or 669 joining clinics throughout the trial. Over 1200 GPs were recruited during the trial. † Socioeconomic status based 670 on quintiles (Q) of each cluster's index of relative socio-economic disadvantage Australian Bureau of Statistics. 671 Socio-Economic Indexes for Areas (SEIFA) (2011). ‡ Remoteness area is based on the Australian Statistical 672 Geography Standard Remoteness Structure of each cluster. Australian Bureau of Statistics. Australian Statistical 673 Geography Standard (ASGS): Remoteness Structure, (2011). § Population of 16 to 29 year old men and women 674 in the cluster. Australian Bureau of Statistics. Population by Age and Sex, Australian States and Territories. 675 (2006). It includes people who have never been sexually active. || n= number of patients aged 16 to 29 years 676 who had at least one chlamydia test in the 12 months prior to the clinic's recruitment. ¶ Based on prevalence 677 survey 1 results: n=number of people who test NAAT positive for chlamydia; N=number of people tested. § Australian National GP Workforce Statistics 2010-2011. Department of Health. GP Workforce Statistics -678

2001-02 to 2016-17. 2017.

679 680 http://health.gov.au/internet/main/publishing.nsf/Content/General+Practice+Statistics-1

Table 2: Primary and morbidity secondary outcomes

In	tervention		Control	U		Adjusted Treatment effect	
n/N	Prevalence (%) (95%CI)	n/N	Prevalence (%) (95%CI)	OR* (95% CI)	P value	OR* (95% CI)	P value
92/1833	5.0(3.8 to 6.2)	88/1925	4.6 (3.5 to 5.7)				
76/2237	3.4 (2.7 to 4.1)	59/1716	3.4 (2.4 to 4.5)	0.9 (0.5 to 1.5)	0.6522	0.9 (0.5 to 1.6)†	0.6727
	-1.6 (-2.9 to -0.3)		-1.1 (-2.7 to 0.5)				
	-0.5 (95%CI: -2.6	to 1.5 p=0.6097)					
n/N	Incidence (95%CI)	ice n/N Incidence RR P value RR		RR (95% CI)	P value		
239/65519	44.7 (28.7 to 60.7)	237/60384	39·2 (28·1 to 50·4)	1.1 (0.7 to 1.8)	0.5622	1.2 (0.8 to 1.9)**	0.4553
	5.5 per 10,000 (95%CI: −1	3.4 to 24.3 p=0.5	693)§§				
57/23527	24.2 (17.2 to 34.1)	88/23219	37.9 (28.9 to 49.7)	0.6 (0.4 to 1.0)	0.0444	$0.6 (0.4 \text{ to } 1.0)^{++}$	0.0407
-	-13.7 per 10,000 (95%CI: -	-26.9 to -0.5 p=0.	0423)§§				
106/41168	25.7 (17.2 to 34.3)	106/38717	27.4 (20.3 to 34.4)	0.9 (0.6 to 1.4)	0.7676	0.9 (0.6 to 1.4)**	0.6790
	-1.6 per 10,000 (95%CI: -	-12.4 to 9.1 p= 0.7	7660)§§				
	n/N 92/1833 76/2237 n/N 239/65519 57/23527 106/41168	(95%CI) 92/1833 5.0 (3.8 to 6.2) 76/2237 3.4 (2.7 to 4.1) -1.6 (-2.9 to -0.3) -0.5 (95%CI: -2.6) n/N Incidence (95%CI) 239/65519 44.7 (28.7 to 60.7) 5.5 per 10,000 (95%CI: -1) 57/23527 24.2 (17.2 to 34.1) -13.7 per 10,000 (95%CI: -1) 106/41168 25.7 (17.2 to 34.3)	n/N Prevalence (%) n/N 92/1833 5.0 (3.8 to 6.2) $88/1925$ 76/2237 3.4 (2.7 to 4.1) $59/1716$ -1.6 (-2.9 to -0.3) -0.5 (95% CI: -2.6 to 1.5 p= 0.6097); n/N Incidence n/N 239/65519 44.7 (28.7 to 60.7) $237/60384$ 5.5 per 10,000 (95% CI: -13.4 to 24.3 p= 0.5 $57/23527$ 24.2 (17.2 to 34.1) $88/23219$ -13.7 per 10,000 (95% CI: -26.9 to -0.5 p= $0.106/41168$	n/N Prevalence (%) (95%CI) n/N Prevalence (%) (95%CI) 92/1833 $5 \cdot 0$ ($3 \cdot 8$ to $6 \cdot 2$) $88/1925$ $4 \cdot 6$ ($3 \cdot 5$ to $5 \cdot 7$) 76/2237 $3 \cdot 4$ ($2 \cdot 7$ to $4 \cdot 1$) $59/1716$ $3 \cdot 4$ ($2 \cdot 4$ to $4 \cdot 5$) $-1 \cdot 6$ ($-2 \cdot 9$ to $-0 \cdot 3$) $-1 \cdot 1$ ($-2 \cdot 7$ to $0 \cdot 5$) $-1 \cdot 1$ ($-2 \cdot 7$ to $0 \cdot 5$) $-0 \cdot 5$ (95%CI: $-2 \cdot 6$ to $1 \cdot 5$ p= $0 \cdot 6097$)‡ Incidence n/N Incidence n/N Incidence (95%CI) (95%CI) (95%CI) 239/65519 $44 \cdot 7$ ($28 \cdot 7$ to $60 \cdot 7$) $237/60384$ $39 \cdot 2$ ($28 \cdot 1$ to $50 \cdot 4$) 5 \cdot 5 per 10,000 (95%CI: $-13 \cdot 4$ to $24 \cdot 3$ p= $0 \cdot 5693$)§§ $57/23527$ $24 \cdot 2$ ($17 \cdot 2$ to $34 \cdot 1$) $88/23219$ $37 \cdot 9$ ($28 \cdot 9$ to $49 \cdot 7$) $-13 \cdot 7$ per 10,000 (95%CI: $-26 \cdot 9$ to $-0 \cdot 5 = 0 \cdot 0423$)§§ $57/23527$ $24 \cdot 2$ ($17 \cdot 2 = 34 \cdot 1$) $88/23219$ $37 \cdot 9$ ($28 \cdot 9 = 10 \cdot 49 \cdot 7$)	n/N Prevalence (%) (95%CI) n/N Prevalence (%) (95%CI) OR* (95%CI) 92/1833 $5 \cdot 0$ ($3 \cdot 8$ to $6 \cdot 2$) $88/1925$ $4 \cdot 6$ ($3 \cdot 5$ to $5 \cdot 7$) $76/2237$ $3 \cdot 4$ ($2 \cdot 7$ to $4 \cdot 1$) $59/1716$ $3 \cdot 4$ ($2 \cdot 4$ to $4 \cdot 5$) $0 \cdot 9$ ($0 \cdot 5$ to $1 \cdot 5$) $-1 \cdot 6$ ($-2 \cdot 9$ to $-0 \cdot 3$) $-1 \cdot 1$ ($-2 \cdot 7$ to $0 \cdot 5$) $-1 \cdot 1 (-2 \cdot 7 \text{ to } 0 \cdot 5)$ $-0 \cdot 5$ (95% CI: $-2 \cdot 6$ to $1 \cdot 5 \text{ p} = 0 \cdot 6097$); n/N Incidence n/N Incidence RR (95%CI) (95%CI) (95% CI) (95% CI) 239/65519 $44 \cdot 7$ ($28 \cdot 7$ to $60 \cdot 7$) $237/60384$ $39 \cdot 2$ ($28 \cdot 1$ to $50 \cdot 4$) $1 \cdot 1$ ($0 \cdot 7$ to $1 \cdot 8$) $5 \cdot 5$ per 10,000 (95% CI: $-13 \cdot 4$ to $24 \cdot 3$ p= $0 \cdot 5693$)§§ $57/23527$ $24 \cdot 2$ ($17 \cdot 2$ to $34 \cdot 1$) $88/23219$ $37 \cdot 9$ ($28 \cdot 9$ to $49 \cdot 7$) $0 \cdot 6$ ($0 \cdot 4$ to $1 \cdot 0$) $-13 \cdot 7$ per 10,000 (95% CI: $-26 \cdot 9$ to $-0 \cdot 5$ p= $0 \cdot 0423$)§§ $106/41168$ $25 \cdot 7$ ($17 \cdot 2$ to $34 \cdot 3$) $106/38717$ $27 \cdot 4$ ($20 \cdot 3$ to $34 \cdot 4$) $0 \cdot 9$ ($0 \cdot 6$ to $1 \cdot 4$)	$\begin{tabular}{ c c c c c } \hline \mathbf{N} & $\mathbf{Prevalence}(\%)$ & \mathbf{N} & $\mathbf{Prevalence}(\%)$ & \mathbf{OR}^* & \mathbf{Pvalue} \\ \hline \mathbf{N} & $\mathbf{Pvalence}(\%)$ & \mathbf{OR}^* & \mathbf{Pvalue} \\ \hline $\mathbf{95\%CI}$ & $-1.1\ (-2.7\ to\ 0.5)$ & $-0.5\ (95\%CI$ & $-2.6\ to\ 1.5\ p=0.6097)^{\ddagger}$ & $10\ (95\%CI$ & $\mathbf{95\%CI}$ & $$	Image: NN Prevalence (%) (95%CI) n/N Prevalence (%) (95%CI) OR* P value (95%CI) OR* 92/1833 $5 \cdot 0$ ($3 \cdot 8$ to $6 \cdot 2$) $88/1925$ $4 \cdot 6$ ($3 \cdot 5$ to $5 \cdot 7$) $3 \cdot 4$ ($2 \cdot 7$ to $4 \cdot 1$) $59/1716$ $3 \cdot 4$ ($2 \cdot 4$ to $4 \cdot 5$) $-1 \cdot 6$ ($-2 \cdot 9$ to $-0 \cdot 3$) $0 \cdot 9$ ($0 \cdot 5$ to $1 \cdot 5$) $0 \cdot 6522$ $0 \cdot 9$ ($0 \cdot 5$ to $1 \cdot 6$) $-1 \cdot 1 (-2 \cdot 7$ to $0 \cdot 5$) m/N Incidence (95%CI) n/N Incidence (95%CI) RR P value (95% CI) RR 239/65519 $4 \cdot 7$ ($28 \cdot 7$ to $60 \cdot 7$) $237/60384$ $39 \cdot 2$ ($28 \cdot 1$ to $50 \cdot 4$) $1 \cdot 1$ ($0 \cdot 7$ to $1 \cdot 8$) $0 \cdot 5622$ $1 \cdot 2$ ($0 \cdot 8$ to $1 \cdot 9$)** $5 \cdot 5$ per 10,000 (95%CI: $-13 \cdot 4$ to $24 \cdot 3$ p= $0 \cdot 5693$)§§ $37 \cdot 9$ ($28 \cdot 9$ to $49 \cdot 7$) $0 \cdot 6$ ($0 \cdot 4$ to $1 \cdot 0$) $0 \cdot 0444$ $0 \cdot 6$ ($0 \cdot 4$ to $1 \cdot 0$)† $-13 \cdot 7$ per 10,000 (95%CI: $-26 \cdot 9$ to $-0 \cdot 5$ p= $0 \cdot 0423$)§§ $37 \cdot 9$ ($28 \cdot 9$ to $34 \cdot 4$) $0 \cdot 9$ ($0 \cdot 6$ to $1 \cdot 4$) $0 \cdot 9$ ($0 \cdot 6$ to $1 \cdot 4$)** $106/41168$ $25 \cdot 7$ ($17 \cdot 2$ to $34 \cdot 3$) $106/38717$ $27 \cdot 4$ ($20 \cdot 3$ to $34 \cdot 4$) $0 \cdot 9$ ($0 \cdot 6$ to $1 \cdot 4$) $0 \cdot 9$ ($0 \cdot 6$ to $1 \cdot 4$)**

OR = Odds Ratio; RR = Risk Ratio; CI = Confidence Interval

* OR is for the interaction between randomisation group and survey 1 versus 2. † The adjusted model includes gender, age group, number of opposite lifetime partners, number of opposite sex partners in the last 12 months, condom use last encounter, sexual health related concerns, any antibiotic use last 3 months, and socioeconomic status of clusters. ‡ Unadjusted treatment effect calculated as the difference between survey2 – survey1 for intervention clusters and survey 2 – survey 1 for control clusters.. **§** Based on diagnoses recorded in the medical records software at each clinic. Incidence is a measure of the incidence proportion among women attending the clinic during a 3 year intervention period. Age group is limited to those aged 16 to 33 years. ¶ Any hospital admission for a primary diagnosis based on the following ICD-10 codes: ICD10 codes N70.0, N70.1, N70.9, N71.0, N71.1, N71.9, N73.0, N73.1, N73.2, N73.3, N73.4, N73.5, N73.8, N73.9, A56.1, N74.4. Age group is limited to those aged 15 to 34 years. Incidence is a measure of the incidence proportion among those attending the hospital during a 3 year intervention period. || RR is for intervention versus control. ** The adjusted model contains age group and socioeconomic status of cluster. †† The adjusted model contains age group only. ‡‡ Based on diagnoses recorded in the medical records software at each clinic. Incidence is a measure of the incidence proportion among men attending the clinic during a 3 year intervention period. Age group is limited to those aged 16 to 29 years. **§§**Unadjusted absolute difference between intervention and control incidence.

supprementary ruble 51. Secondary dateons		ntervention		ontrol	Unadjusted		Adjuste	
			(b. 7	D (1) (0()	Treatment eff		Treatment e	
	n/N	Proportion (%) (95%CI)	n/N	Proportion (%) (95%CI)	OR* (95% CI)	P value	OR* (95% CI)	P value
Chlamydia testing by time since randomisation †		()3/001)		()5/001)	()) /0 (1)		()5/0 (1)	
-1	2802/34143	8·2 (7·0 to 9·4)	3107/37775	8.2 (7.2 to 9.2)	1.0 (0.8 to 1.2)	0.9810	1.0 (0.9 to 1.3)	0.6733
$^{-1}_{0}$	5893/42418	13.9 (12.8 to 15.0)	6115/44262	13.8 (12.5 to 15.2)	1.0 (0.3 to 1.2) 1.0 (0.9 to 1.2)	0.9810	1.0 (0.9 to 1.3) 1.0 (0.9 to 1.2)¶	0.0733
0	6264/43968	13.9(12.8 to 15.9) 14.2(12.6 to 15.9)	5005/44666	13.8(12.5 to 15.2) 11.2(9.5 to 12.9)	1.0 (0.9 to 1.2) 1.3 (1.1 to 1.6)	0.0093	1.4 (1.1 to 1.7)	0.0045
1		· · · · · · · · · · · · · · · · · · ·					· · · · · · · · · · · · · · · · · · ·	
2	8494/44005	19.3 (16.7 to 21.9)	5090/43438	11.7 (10.1 to 13.3)	1.8 (1.4 to 2.2)	<.0001	$2 \cdot 0 (1 \cdot 6 \text{ to } 2 \cdot 5)$	<.0001
	8779/43676	20.1 (18.4 to 21.8)	5168/40156	12.9 (11.2 to 14.5)	1.7 (1.4 to 2.0)	<-0001	1.8 (1.5 to 2.1)	<-0001
Absolute difference‡		11.9 (10.3 to 13.4)		4.6 (3.3 to 6.0)				
Relative change in testing: year 3 vs year -1		2.8 (2.4 to 3.2)§		1.6 (1.4 to 1.9) §	1.7 (1.4 to 2.1)	<.0001	1.7 (1.4 to 2.1)	<.0001
Chlamydia positivity by time since randomisation †								
-1	250/2364**	10.6 (9.0 to 12.2)	310/2704**	11.5 (8.6 to 14.3)	0.9 (0.7 to 1.3)	0.5720	0.9 (0.6 to 1.2)¶	0.4812
0	535/5071**	10.6 (9.6 to 11.5)	485/5309**	9.1 (8.0 to 10.3)	1.2 (1.0 to 1.4)	0.0605	1.2 (1.0 to 1.4)	0.0756
1	573/5230**	11.0 (9.7 to 12.2)	457/4152**	11.0 (9.1 to 12.9)	1.0 (0.8 to 1.2)	0.9638	0.9 (0.8 to 1.2)	0.5993
2	634/7084**	8.9 (7.7 to 10.2)	438/4341**	10.1 (8.7 to 11.5)	0.9 (0.7 to 1.1)	0.2310	0.9 (0.7 to 1.1)	0.2077
3	498/7316**	6.8(5.6 to 8.0)	394/4478**	8.8 (7.7 to 9.9)	0.8 (0.6 to 1.0)	0.0173	0.7 (0.6 to 0.9)	0.0151
Absolute difference [†]		-3.8 (-6.0 to -1.5)		-2.7 (-5.1 to -0.2)			· · · · · ·	
Relative change in positivity: year 3 vs year -1		0.6 (0.5 to 0.8) §		0.7 (0.5 to 1.0) §	0.8 (0.6 to 1.2)	0.3271	0.8 (0.5 to 1.2)	0.3304
Chlamydia retesting (10-15 months) after a negative test		0 0 (0 0 10 0 0) 3		0 / (0 2 10 1 0/3		0 02/1	0 0 (0 0 10 1 2)	0 0000.
result by time since randomisation [†]								
-1	167/2193	7.6 (6.3 to 8.9)	189/2489	7.6 (6.3 to 8.8)	1.0 (0.8 to 1.3)	0.9804	1.0 (0.8 to 1.3)	0.8220
$^{-1}_{0}$	397/4716							
0		8.4 (7.1 to 9.7)	272/5000	$5 \cdot 4 \ (4 \cdot 6 \ \text{to} \ 6 \cdot 3)$	1.6 (1.3 to 2.0)	<.0001	1.7 (1.3 to 2.1)	<.0001
1	506/4880	10.4 (8.7 to 12.1)	269/3862	7.0 (6.1 to 7.9)	1.5 (1.2 to 1.9)	<.0001	1.6 (1.3 to 2.0)	<.0001
2	610/6690	9.1 (8.3 to 9.9)	310/4075	7.6 (6.3 to 8.9)	1.2 (1.0 to 1.5)	0.0603	1.3 (1.1 to 1.6)¶	0.0130
3	260/7010	3.7 (2.2 to 5.2)	81/4236	1.9 (1.2 to 2.7)	$2 \cdot 0 \ (1 \cdot 1 \text{ to } 3 \cdot 5)$	0.0182	$2 \cdot 1 \ (1 \cdot 2 \text{ to } 3 \cdot 8)$	0.0147
Absolute difference‡		-3.9 (-5.6 to -2.3)		-5.7 (-7.0 to -4.4)				
Relative change in retesting: year 3 vs year -1		0.5 (0.3 to 0.7)		0.2 (0.2 to 0.4) §	2.0 (1.1 to 3.5)	0.0196	$2 \cdot 0 (1 \cdot 1 \text{ to } 3 \cdot 6) \ $	0.0149
Chlamydia retesting (6 weeks - 6 months) after a positive								
test result by time since randomisation [†]								
-1	46/250	18.4 (13.6 to 23.2)	49/310	15.8 (10.7 to 20.9)	1.2 (0.7 to 1.9)	0.4543	$1.2 (0.7 \text{ to } 1.9)^{++}$	0.4673
0	118/535	22.1 (18.2 to 25.9)	93/485	19.2 (15.2 to 23.2)	1.2 (0.9 to 1.7)	0.2976	1.2 (0.9 to 1.7)	0.2572
1	159/573	27.7 (23.9 to 31.6)	90/457	19.7 (16.2 to 23.2)	1.6 (1.2 to 2.1)	0.0017	1.6 (1.2 to 2.2)	0.0015
2	158/634	24.9 (22.2 to 27.7)	84/438	19.2 (14.0 to 24.3)	1.4 (1.0 to 2.0)	0.0608	$1.4 (1.0 \text{ to } 2.0)^{++}$	0.0611
3	113/498	22.7 (18.9 to 26.5)	68/394	17.3 (13.1 to 21.4)	1.4 (1.0 to 2.0)	0.0537	1.6 (1.1 to 2.2)	0.0103
Absolute difference:	115/190	4.3 (-1.5 to 10.0)	00/371	1.5 (-5.8 to 8.7)	1 1 (1 0 to 2 0)	0 0007	10(11022)	0 0105
Relative change in retesting: year 3 vs year –1		1.3 (0.9 to 1.9) §		$1 \cdot 1 (0 \cdot 7 \text{ to } 1 \cdot 9) $	$1 \cdot 2 (0 \cdot 6 \text{ to } 2 \cdot 2)$	0.6284	$1 \cdot 2 (0 \cdot 6 \text{ to } 2 \cdot 3)$	0.5277
Repeat chlamydia infection (6 weeks – 6 months) after a		1.3 (0.9 to 1.9) §		1.1 (0.7 to 1.9) §	1.2(0.010(2.2))	0.0204	1.2(0.0102.3)	0.3211
positive test result by by time since randomisation [†]								
	5/46	10.0(1.2 to 20.5)	10/40	20.4 (6.2 to 34.7)	$0.5(0.1 \pm 0.17)$	0.2470	0.3 (0.1 to 1.0)	0.0515
-1		10.9 (1.3 to 20.5) 12.7 (7.0 to 18.4)	10/49		0.5 (0.1 to 1.7)			
0	15/118	12.7 (7.0 to 18.4)	12/93	12.9 (7.0 to 18.8)	1.0 (0.5 to 2.0)	0.9621	1.0 (0.5 to 2.0)	0.9680
l	22/159	13.8 (9.5 to 18.2)	10/90	11.1 (6.5 to 15.7)	$1 \cdot 3 \ (0 \cdot 7 \text{ to } 2 \cdot 3)$	0.3871	$1.3 (0.7 \text{ to } 2.3)^{\dagger\dagger}$	0.4378
2	20/158	12.7 (6.3 to 19.0)	13/84	15.5 (7.1 to 23.8)	0.8 (0.3 to 1.8)	0.5770	0.8 (0.3 to 1.7)	0.4837
3	12/113	10.6 (4.4 to 16.9)	5/68	7.4 (0.0 to 14.8)	1.5 (0.4 to 5.1)	0.5187	1.4 (0.4 to 5.1)	0.5650
Absolute difference‡		-0.3 (-12.3 to 11.8)		$-13 \cdot 1(-23 \cdot 3 to -2 \cdot 8)$				
Relative change in repeat infection: year 3 vs year -1		1.0 (0.3 to 3.4) §		0.3 (0.1 to 0.7) §	3.1 (0.7 to 13.8)	0.1287	$3 \cdot 6(0 \cdot 9 \text{ to } 14 \cdot 2) \ \P$	0.0724

Supplementary Table S1: Secondary outcomes – chlamydia testing, positivity, re-testing and repeat infection

OR = Odds Ratio; CI = Confidence Interval

* OR is for intervention versus control. \dagger Time since randomisation refers to time since the start of the intervention period in each cluster and is defined as : -1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2 as (13-24) months after randomisation; 3 as (25-36) months after randomisation. \ddagger Absolute difference between relative year 3 and year -1. \$ Odds ratio for relative year 3 vs year -1. \parallel Odds ratio for intervention vs control for year 3 vs year -1. \ast The denominator for chlamydia positivity is not equivalent to the number of chlamydia tests undertaken because chlamydia test results are not always stored electronically in the medical records at each clinic. Positivity is calculated based on the number of tests for which test results were stored electronically. \P Adjusted for gender, age group and socioeconomic status of clusters. \dagger Adjusted for age and gender. \ddagger Adjusted for gender only.

Subgroup	abgroup		Intervention		Control	Unadjus Treatment		Interaction test
		n/N	Prevalence (%) (95%CI)	n/N	Prevalence (%) (95%CI)	OR* (95% CI)	P value	P value†
Male	Survey 1	33/557	5.9 (4.0 to 7.8)	27/531	5.1(3.5 to 6.7)	0.7		
	Survey 2	19/621	3.1(1.8 to 4.4)	18/487	3.7 (2.0 to 5.4)	(0.3 to 1.6)	0.3814	0.4547
Female	Survey 1	59/1276	4.6(3.3 to 5.9)	61/1394	4.4 (3.1 to 5.6)	1.0		
	Survey 2	57/1616	3.5(2.5 to 4.6)	41/1229	3.3 (2.1 to 4.5)	(0.5 to 1.8)	0.9732	
Age [16-19 years]	Survey 1	35/432	8.1 (5.3 to 10.9)	35/521	6.7 (4.4 to 9.1)	0.9		
	Survey 2	34/570	6.0(3.9 to 8.1)	23/424	5.4 (3.0 to 7.9)	(0.4 to 2.1)	0.8028	
Age [20-24 years]	Survey 1	43/714	6.0 (4.0 to 8.0)	41/697	5.9 (3.9 to 7.9)	0.7		
	Survey 2	26/834	3.1(1.9 to 4.3)	28/696	4.0(2.6 to 5.4)	(0.4 to 1.4)	0.3341	0.7352
Age [25-29 years]	Survey 1	14/687	2.0(1.0 to 3.1)	12/707	1.7 (1.0 to 2.4)	1.2		
	Survey 2	16/833	1.9 (1.0 to 2.9)	8/596	1.3 (0.2 to 2.5)	(0.4 to 4.0)	0.7834	
Aboriginal/Torres Strait Islander	Survey 1	8/103	7.8 (2.5 to 13.0)	6/124	4.8 (2.1 to 7.6)	0.9		
C	Survey 2	8/164	4.9(1.4 to 8.3)	4/128	3.1 (0.9 to 5.3)	(0.2 to 4.0)	0.9373	
Not (Aboriginal/Torres Strait Islander)	Survey 1	77/1578	4.9(3.6 to 6.1)	75/1662	4.5 (3.3 to 5.7)	0.9		0.9295
	Survey 2	68/2015	3.4(2.7 to 4.0)	54/1562	3.5 (2.3 to 4.6)	(0.5 to 1.5)	0.6546	
Cluster chlamydia prevalence prior to	Survey 1	11/590	1.9(1.0 to 2.7)	7/487	1.4 (0.6 to 2.3)	0.7		
randomisation $< 3\%$ ‡	Survey 2	24/729	3.3(2.1 to 4.5)	20/531	3.8(1.8 to 5.8)	(0.3 to 1.8)	0.4159	
Cluster chlamydia prevalence prior to	Survey 1	81/1243	6.5(5.3 to 7.7)	81/1438	5.6 (4.6 to 6.6)	0.9		0.5813
randomisation $\geq 3\%$ [‡]	Survey 2	52/1508	3.4(2.6 to 4.3)	39/1185	3.3(1.9 to 4.6)	(0.5 to 1.6)	0.7559	
Cluster chlamydia testing in the 12 months	Survey 1	14/299	4.7 (1.9 to 7.4)	13/281	4.6 (2.4 to 6.8)	1.0		
prior to randomisation < 6%§	Survey 2	14/322	4.3 (2.4 to 6.3)	11/262	4.2 (0.2 to 8.2)	(0.3 to 3.7)	0.9604	
Cluster chlamydia testing in the 12 months	Survey 1	78/1534	$5 \cdot 1 (3 \cdot 7 \text{ to } 6 \cdot 5)$	75/1644	4.6 (3.3 to 5.8)	0.9		0.7964
prior to randomisation $\geq 6\%$ §	Survey 2	62/1915	$3 \cdot 2 (2 \cdot 5 \text{ to } 4 \cdot 0)$	48/1454	3.3 (2.2 to 4.4)	(0.5 to 1.5)	0.6052	

Supplementary Table S2: Subgroup analyses of primary outcome – chlamydia prevalence Subgroup

OR = Odds Ratio; CI = Confidence Interval

* OR is for the interaction between randomisation group and survey 1 versus 2. † Interaction test between study group and variable of interest. ‡ Based on survey 1 prevalence cut-point of 3% used in the randomisation algorithm. § Based on chlamydia testing cut-point of 6% used in the randomisation algorithm.

	Intervention			Control	Unadju Treatmen		Adjusted Treatment effect	
Primary outcome	n/N	Prevalence (%) (95%CI)	n/N	Prevalence (%) (95%CI)	OR† (95% CI)	P value	OR†‡ (95% CI)	P value
Chlamydia prevalence								
Survey 1	78/1623	4.8 (3.5 to 6.1)	33/793	4.2 (2.5 to 5.8)	0.7	0.3990	0.7	0.4342
Survey 2	67/2009	3.3 (2.6 to 4.1)	27/669	4.0 (1.6 to 6.5)	(0.3 to 1.6)		(0.3 to 1.7)	
Difference: survey 2-survey 1		-1.5 (-2.9 to -0.1)		-0.1 (-3.1 to 3.0)				
Treatment effect: difference (intervention-control)		-1·4 (95% CI	-4.7 to 1.9	9 p=0.4096)§				

Supplementary Table S3: Secondary analysis of primary outcome chlamydia prevalence, in clinics that adhered to their randomisation allocation*

OR = Odds Ratio; CI = Confidence Interval

* Analysis is limited to intervention clinics testing $\geq 10.7\%$ and control clinics testing < 10.7% in the 12 months prior to survey 2, irrespective of the cluster's testing performance before the start of the trial. A total of 23 intervention and 11 control clusters were included in the analysis. \dagger OR is for the interaction between randomisation group and survey 1 versus 2. \ddagger Adjusted for gender, age group, socioeconomic status of clusters. \$ Unadjusted treatment effect calculated as the difference between survey2– survey1 for intervention clusters and survey 2 – survey 1 for control clusters.

		Intervention		Co	ontrol	Unadjusted Treatment effect		Interaction test	
Subgroup	Chlamydia testing by time since randomisation*	n/N	Proportion (%) (95%CI)	n/N	Proportion (%) (95%CI)	OR (95% CI)	P value	P value†	
Males	-1	636/14061	4.5 (3.9 to 5.2)	711/15731	4.5 (4.0 to 5.1)	1.0 (0.8 to 1.2)	0.9936	0.9058	
	0	1560/17998	8.7 (7.8 to 9.5)	1532/18780	8·2 (7·2 to 9·1)	1·1 (0·9 to 1·3)	0.4201	0.2040	
	1	1560/18490	8.4 (7.4 to 9.5)	1115/19061	5.8 (4.7 to 7.0)	1.5 (1.2 to 1.9)	0.0015	0.0792	
	2	2304/18552	12.4 (10.1 to 14.8)	1216/18465	6.6 (5.5 to 7.7)	2.0 (1.5 to 2.6)	<.0001	0.1225	
	3	2296/18212	12.6 (11.2 to 14.0)	1230/16952	7.3 (6.1 to 8.5)	1.8 (1.5 to 2.3)	<.0001	0.1748	
Females	-1	2166/20082	10.8 (9.1 to 12.5)	2396/22044	10.9 (9.5 to 12.2)	1.0 (0.8 to 1.2)	0.9380		
	0	4333/24420	17.7 (16.2 to 19.3)	4583/25482	18.0 (16.4 to 19.6)	1.0 (0.8 to 1.1)	0.8253		
	1	4704/25478	18.5 (16.2 to 20.8)	3890/25605	15.2 (13.0 to 17.4)	1.3 (1.0 to 1.6)	0.0373		
	2	6190/25453	24·3 (21·4 to 27·3)	3874/24973	15.5 (13.4 to 17.6)	1.8 (1.4 to 2.2)	<.0001		
	3	6483/25464	25.5 (23.5 to 27.5)	3938/23204	17.0 (14.9 to 19.1)	1.7 (1.4 to 2.0)	<.0001		
Age	-1	1016/11834	8.6 (7.4 to 9.7)	1121/13267	8·4 (7·4 to 9·5)	1.0 (0.8 to 1.2)	0.8607	0.8203	
[16-19 years]	0	1925/14282	13.5 (12.0 to 14.9)	1994/15522	12.8 (11.5 to 14.2)	$1 \cdot 1 \ (0 \cdot 9 \text{ to } 1 \cdot 2)$	0.5143	0.4995	
	1	2114/14463	14.6 (12.6 to 16.6)	1721/15287	11.3 (9.5 to 13.0)	1.3 (1.1 to 1.7)	0.0097	0.3212	
	2	2798/14145	19.8 (17.2 to 22.3)	1741/14781	11.8 (10.0 to 13.6)	1.8 (1.5 to 2.3)	<.0001	0.8235	
	3	2738/13744	19·9 (18·1 to 21·7)	1713/13576	12.6 (10.5 to 14.7)	1.7 (1.4 to 2.1)	<.0001	0.9613	
Age	-1	1136/11675	9.7 (8.2 to 11.3)	1224/12499	9.8 (8.5 to 11.1)	1.0 (0.8 to 1.2)	0.9498		
[20-24 years]	0	2349/14786	15.9 (14.6 to 17.2)	2362/14608	16.2 (14.5 to 17.9)	1.0 (0.8 to 1.1)	0.7884		
	1	2572/15283	16.8 (14.9 to 18.7)	1957/15076	13.0 (10.9 to 15.0)	1.4 (1.1 to 1.7)	0.0063		
	2	3412/15416	22.1 (19.3 to 25.0)	1999/14697	13.6 (11.6 to 15.7)	1.8 (1.4 to 2.3)	<.0001		
	3	3542/15311	23.1 (21.2 to 25.1)	2073/13754	15.1 (13.2 to 16.9)	1.7 (1.4 to 2.0)	<.0001		
Age	-1	650/10634	6.1 (5.1 to 7.2)	762/12009	6·3 (5·3 to 7·4)	1.0 (0.8 to 1.2)	0.7493		
[25-29 years]	0	1619/13350	12.1 (11.0 to 13.2)	1759/14132	12.4 (11.1 to 13.8)	1.0 (0.8 to 1.1)	0.7072		
	1	1578/14222	11.1 (9.6 to 12.6)	1327/14303	9.3 (7.9 to 10.7)	1.2 (1.0 to 1.5)	0.0700		
	2	2284/14444	15.8 (13.1 to 18.5)	1350/13960	9.7 (8.3 to 11.0)	1.8 (1.4 to 2.2)	<.0001		
	3	2499/14621	17.1 (15.5 to 18.7)	1382/12826	10.8 (9.4 to 12.1)	1.7 (1.4 to 2.0)	<.0001		

Supplementary Table S4: Subgroup analyses of secondary outcomes – chlamydia testing

OR = Odds Ratio; CI = Confidence Interval

* Time since randomisation refers to time since the start of the intervention period in each cluster and is defined as : 1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2 as (13-24) months after randomisation; 3 as (25-36) months after randomisation. † Interaction test between study group and variable of interest (gender or age).

		Intervention		C	Control	Unadjusted Treatment effect		Interaction test	
Subgroup	Chlamydia positivity by time since randomisation*†	n/N	Proportion (%) (95%CI)	n/N	Proportion (%) (95%CI)	OR (95% CI)	P value	P value‡	
Male	-1	98/519	18.9 (15.6 to 22.1)	107/602	17.8 (14.1 to 21.5)	$1 \cdot 1 \ (0 \cdot 8 \text{ to } 1 \cdot 5)$	0.6482	0.2450	
	0	174/1288	13.5 (11.4 to 15.6)	167/1322	12.6 (10.9 to 14.4)	1.1 (0.9 to 1.4)	0.5135	0.4224	
	1	190/1228	15.5 (13.6 to 17.3)	148/882	16.8 (14.1 to 19.4)	0.9 (0.7 to 1.1)	0.4059	0.4472	
	2	225/1821	12.4 (10.1 to 14.6)	151/994	15.2 (12.9 to 17.5)	0.8 (0.6 to 1.0)	0.0776	0.2773	
	3	179/1827	9.8 (7.3 to 12.3)	133/1044	12.7 (10.4 to 15.1)	0.7 (0.5 to 1.0)	0.0865	0.9583	
Female	-1	152/1845	8.2 (6.9 to 9.6)	203/2102	9.7 (6.5 to 12.8)	0.8 (0.6 to 1.2)	0.3796		
	0	361/3783	9.5 (8.6 to 10.5)	318/3987	8.0 (6.6 to 9.3)	1.2 (1.0 to 1.5)	0.0659		
	1	383/4002	9.6 (8.2 to 10.9)	309/3270	9.4 (7.5 to 11.4)	1.0 (0.8 to 1.3)	0.9185		
	2	409/5263	7.8 (6.6 to 9.0)	287/3347	8.6 (7.2 to 9.9)	0.9(0.7 to 1.1)	0.3643		
	3	319/5489	5.8 (4.9 to 6.8)	261/3434	7.6 (6.4 to 8.8)	0.8 (0.6 to 0.9)	0.0160		
Age	-1	118/842	14.0 (12.0 to 16.0)	125/967	12.9 (9.6 to 16.2)	$1 \cdot 1 \ (0 \cdot 8 \text{ to } 1 \cdot 5)$	0.5746	0.1410	
[16-19 years]	0	256/1642	15.6 (13.3 to 17.9)	190/1704	11.2 (8.8 to 13.5)	1.5 (1.1 to 1.9)	0.0071	0.0082	
	1	253/1764	14.3 (12.1 to 16.5)	197/1402	14.1 (11.2 to 16.9)	1.0 (0.8 to 1.4)	0.8699	0.1486	
	2	277/2291	12.1 (10.0 to 14.2)	179/1477	12.1 (9.6 to 14.6)	1.0 (0.7 to 1.3)	0.9860	0.2613	
	3	203/2251	9.0(6.7 to 11.4)	161/1460	11.0 (9.4 to 12.7)	0.8 (0.6 to 1.1)	0.1717	0.2522	
Age	-1	107/974	11.0 (8.4 to 13.5)	147/1087	13.5 (10.5 to 16.6)	0.8 (0.6 to 1.1)	0.1927		
[20-24 years]	0	220/2023	10.9 (9.8 to 12.0)	208/2050	10.1 (8.8 to 11.5)	1.1 (0.9 to 1.3)	0.3978		
	1	228/2142	10.6 (8.9 to 12.4)	201/1637	12.3 (9.8 to 14.7)	0.9 (0.6 to 1.1)	0.2609		
	2	267/2867	9.3 (7.8 to 10.8)	187/1708	10.9 (9.3 to 12.6)	0.8 (0.7 to 1.1)	0.1380		
	3	201/2939	6.8 (5.7 to 7.9)	173/1802	9.6 (8.4 to 10.8)	0.7 (0.6 to 0.9)	0.0007		
Age	-1	25/548	4.6 (2.8 to 6.3)	38/650	5.8 (2.7 to 9.0)	0.8 (0.4 to 1.5)	0.4437		
[25-29 years]	0	59/1406	4.2(3.1 to 5.3)	87/1555	5.6(4.6 to 6.6)	0.7 (0.5 to 1.0)	0.0582		
	1	92/1324	6.9(5.1 to 8.8)	59/1113	5.3 (4.1 to 6.5)	1.3 (0.9 to 1.9)	0.1105		
	2	90/1926	4.7(3.6 to 5.8)	72/1156	6.2 (4.9 to 7.6)	0.7 (0.5 to 1.0)	0.0701		
	3	94/2126	4.4 (3.5 to 5.3)	60/1216	4.9(3.6 to 6.3)	0.9 (0.6 to 1.3)	0.5213		

Supplementary Table S5: Subgroup analyses of secondary outcomes – chlamydia positivity

OR = Odds Ratio; CI = Confidence Interval

* Time since randomisation refers to time since the start of the intervention period in each cluster and is defined as : -1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2 as (13-24) months after randomisation; 3 as (25-36) months after randomisation. †The denominator for chlamydia positivity is not equivalent to the number of chlamydia tests undertaken because chlamydia test results are not always stored electronically in the medical records at each clinic. Positivity is calculated based on the number of tests for which test results were stored electronically. ‡ Interaction test between study group and variable of interest (gender or age).

	ary rubic 501 Subgroup analysis of secondary		Intervention		Control	Unadjuste Treatment e	Interaction test	
Subgroup	Chlamydia retesting by time since randomisation*	n/N	Proportion (%) (95%CI)	n/N	Proportion (%) (95%CI)	OR (95% CI)	P value	P value†
Male	-1	22/443	5.0 (2.7 to 7.2)	22/513	4.3 (2.9 to 5.7)	1.2 (0.7 to 2.1)	0.5964	0.5895
	0	56/1158	4.8 (3.2 to 6.5)	29/1206	$2 \cdot 4 (1 \cdot 2 \text{ to } 3 \cdot 6)$	$2 \cdot 1(1 \cdot 1 \text{ to } 3 \cdot 8)$	0.0179	0.3697
	1	62/1087	5.7 (4.1 to 7.3)	31/770	4.0 (2.0 to 6.0)	1.4 (0.8 to 2.6)	0.2142	0.7525
	2	85/1653	$5 \cdot 1 (4 \cdot 1 \text{ to } 6 \cdot 2)$	40/879	4.6 (3.1 to 6.0)	$1 \cdot 1 \ (0 \cdot 8 \text{ to } 1 \cdot 7)$	0.5160	0.6263
	3	48/1693	2.8 (1.4 to 4.3)	9/950	0.9 (0.4 to 1.5)	3.1 (1.4 to 6.6)	0.0044	0.1735
Female	-1	145/1750	8.3 (6.7 to 9.9)	167/1976	8.5 (7.0 to 9.9)	1.0 (0.7 to 1.3)	0.8768	
	0	341/3558	9.6 (7.9 to 11.2)	243/3794	6.4 (5.4 to 7.4)	1.5 (1.2 to 2.0)	0.0004	
	1	444/3793	11.7 (9.7 to 13.7)	238/3092	7.7 (6.7 to 8.7)	1.6 (1.3 to 2.0)	<.0001	
	2	525/5037	10.4 (9.4 to 11.5)	270/3196	8.4 (6.9 to 10.0)	1.3 (1.0 to 1.6)	0.0374	
	3	212/5317	4.0 (2.4 to 5.6)	72/3286	$2 \cdot 2 (1 \cdot 3 \text{ to } 3 \cdot 1)$	1.9 (1.0 to 3.3)	0.0372	
Age	-1	78/760	10.3 (8.1 to 12.5)	88/883	10.0 (7.8 to 12.1)	1.0 (0.7 to 1.4)	0.8449	0.5510
[16-19 years]	0	164/1523	10.8 (8.6 to 12.9)	116/1644	7.1 (5.9 to 8.2)	1.6 (1.2 to 2.1)	0.0009	0.7790
	1	251/1822	13.8 (11.4 to 16.2)	123/1411	8.7 (7.2 to 10.2)	1.7 (1.3 to 2.2)	0.0001	0.3541
	2	302/2515	12.0 (10.3 to 13.7)	160/1610	9.9 (8.1 to 11.7)	1.2 (1.0 to 1.6)	0.0937	0.9964
	3	112/2723	4.1 (2.5 to 5.7)	46/1752	2.6 (1.5 to 3.8)	1.6 (0.9 to 2.8)	0.1173	0.1092
Age	-1	67/899	7.5 (5.9 to 9.1)	68/983	6.9 (5.2 to 8.7)	$1 \cdot 1 \ (0 \cdot 8 \text{ to } 1 \cdot 5)$	0.6490	
[20-24 years]	0	166/1856	8.9 (7.7 to 10.2)	105/1892	5.5 (4.4 to 6.7)	1.7 (1.3 to 2.2)	<.0001	
	1	182/1890	9.6 (7.9 to 11.4)	110/1438	7.6 (5.8 to 9.5)	1.3 (0.9 to 1.8)	0.1232	
	2	211/2529	8.3 (6.9 to 9.8)	103/1485	6.9 (5.3 to 8.6)	1.2 (0.9 to 1.7)	0.1970	
	3	105/2583	$4 \cdot 1 \ (2 \cdot 0 \text{ to } 6 \cdot 2)$	28/1530	1.8 (0.9 to 2.7)	2.3 (1.1 to 4.7)	0.0248	
Age	-1	22/534	4.1 (2.6 to 5.6)	33/623	5.3 (3.2 to 7.3)	0.8 (0.4 to 1.3)	0.3360	
[25 to 29	0	67/1337	5.0(3.7 to 6.3)	51/1464	3.5 (2.4 to 4.6)	1.5 (1.0 to 2.2)	0.0696	
years]	1	73/1168	6.3 (4.8 to 7.7)	36/1013	3.6(2.5 to 4.6)	1.8 (1.2 to 2.6)	0.0019	
	2	97/1646	5.9 (4.8 to 7.0)	47/980	4.8 (3.3 to 6.3)	1.2 (0.9 to 1.8)	0.2538	
	3	43/1704	2.5 (1.5 to 3.5)	7/954	0.7 (0.2 to 1.3)	3.5 (1.6 to 7.8)	0.0020	

Supplementary Table S6: Subgroup analysis of secondary outcomes – retesting within 10 to 15 months after a negative test

OR = Odds Ratio; CI = Confidence Interval

* Time since randomisation refers to time since the start of the intervention period in each cluster and is defined as : -1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2 as (13-24) months after randomisation; 3 as (25-36) months after randomisation. † Interaction test between study group and variable of interest (gender or age).

	Intervention Control		Control	Unadjusted Treatment effect		Adjusted Treatment effect		
Retesting after a positive test	n/N	Proportion (%) (95%CI)	n/N	Proportion (%) (95%CI)	OR* (95% CI)	P value	OR* (95% CI)	P value
Chlamydia retesting (< 3 weeks) after a positive test result by time since randomisation [†]							· · · ·	
	30/250	12.0 (6.3 to 17.7)	38/310	12.3 (6.9 to 17.6)	1.0 (0.5 to 2.0)	0.9469	$1.0 (0.5 \text{ to } 2.2)^{\parallel}$	0.9252
) 50/535	9.3 (6.3 to 12.4)	70/485	14.1 (11.2 to 17.6)	0.6 (0.4 to 0.9)	0.0232	$0.7 (0.4 \text{ to } 1.0)^{\text{I}}$	0.0749
	57/573	9.9 (5.9 to 14.0)	71/457	15.5 (11.9 to 19.2)	0.6 (0.4 to 1.0)	0.0502	$0.6 (0.4 \text{ to } 1.0)^{\parallel}$	0.0464
	2 61/634	9.6 (7.3 to 11.9)	66/438	15.1 (11.0 to 19.2)	0.6 (0.4 to 0.9)	0.0131	$0.6 (0.4 \text{ to } 0.9)^{\text{I}}$	0.0101
	3 48/498	9.6 (6.6 to 12.7)	44/394	11.2 (7.5 to 14.9)	0.8 (0.5 to 1.4)	0.5135	$0.9 (0.6 \text{ to } 1.5)^{\text{\$}}$	0.7375
Chlamydia retesting (6 weeks – 4 months) after a positive test result by time since randomisation [†]								
	36/250	14·4 (10·4 to 18·4)	34/310	11.0 (6.8 to 15.1)	1.4 (0.8 to 2.3)	0.2367	$1.3 (0.8 \text{ to } 2.3)^{\parallel}$	0.2737
) 93/535	17.4 (14.4 to 20.4)	73/485	15·1 (11·1 to 19·0)	1.2 (0.8 to 1.7)	0.3510	$1.2 (0.8 \text{ to } 1.7)^{\text{¶}}$	0.2914
	1 108/573	18.8 (15.7 to 22.0)	62/457	13.6 (9.4 to 17.7)	1.5 (1.0 to 2.2)	0.0516	$1.5 (1.0 \text{ to } 2.3)^{\text{¶}}$	0.0544
	2 117/634	18.5 (15.7 to 21.2)	61/438	13.9 (10.3 to 17.5)	1.4 (1.0 to 2.0)	0.0523	$1.4 (1.0 \text{ to } 2.0)^{\parallel}$	0.0531
	3 92/498	18.5 (14.6 to 22.3)	49/394	12.4 (9.1 to 15.8)	1.6 (1.1 to 2.3)	0.0177	$1.7 (1.1 \text{ to } 2.6)^{\text{II}}$	0.0086

Supplementary Table S7: Secondary outcomes – alternative definitions of retesting after a positive test

OR = Odds Ratio; CI = Confidence Interval

* OR is for intervention versus control. \dagger Time since randomisation refers to time since the start of the intervention period in each cluster and is defined as : to 1 is defined as the (24-13) months before randomisation; 0 is defined as (12-1) months before randomisation, prevalence survey 1 was conducted during this time; 1 as (1-12) months after randomisation; 2 as (13-24) months after randomisation; 3 as (25-36) months after randomisation|| Adjusted for age and gender. ¶ Adjusted for gender, age group and socioeconomic status of clusters.