

Homogenization of daily ECA&D temperature series

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September 17, 2018

abstract The daily maximum and minimum temperature series of the European Climate Assessment & Dataset are homogenised using the quantile matching approach. As the dataset is large and the detail of metadata is generally missing, an automated method locates breaks in the series based on a comparison with surrounding series and applies adjustments which are estimated using homogeneous segments of surrounding series as reference. A total of 6500 series have been processed and after removing duplicates and short series, about 2100 series have been adjusted. Finally, the effect of the homogenization of daily maximum and minimum temperature on trend estimation is shown to produce a much more spatially homogeneous and then plausible picture.

1 Introduction

Modifications to meteorological stations, such as relocation, replacement of the instrument, recalibration, new buildings in the neighborhood or growth of vegetation in the proximity, alter temperature measurements and introduce biases in the observational records that do not relate to weather and climate [Aguilar et al., 2003, Hartmann et al., 2013]. The analysis of climatic variability and climatic change requires homogeneous temperature series [Peterson et al., 1998]: these series do not confuse the climatic signal with artificial biases which are present in non-homogenous series [Menne and Williams Jr, 2009, Brunetti et al., 2006, Thorne et al., 2005, Begert et al., 2005]. Prior to climate analyses, actions are required aimed at the removal of step-like or gradual changes related to these non-climatic effects in observational records [Caussinus and Mestre, 2004].

The registration of activities on meteorological stations in the metadata keeps track of these changes and sufficiently detailed metadata allow a precise temporal localization of the breaks. Unfortunately the availability of metadata is often low, especially further back in time, and doesn't cover the whole set of inhomogeneities that affect the measurements [Caussinus and Mestre, 2004]. This implies that break-detection based on metadata only is not possible for many datasets, even though this approach is regarded as most accurate and reliable. This argument, and the sheer size of a dataset, motivates the use of an automated homogenization procedure [Caussinus and Mestre, 2004].

The aim of this study is to develop a pan-European homogeneous dataset of daily maximum and minimum temperature using such an automated homogenization procedure. It will use a recent agreement-based system to detect breaks [Kuglitsch et al., 2012] and the quantile matching technique [Trewin, 2013] in combination with a pairwise-comparison [Menne and Williams Jr, 2005] approach to determine adjustments. The elements in this approach are introduced below.

Automated homogenization procedures consists of two steps: break detection and adjustment calculation (which follow - or are integrated with - a quality check procedure) [Alexandersson, 1986, Caussinus and Mestre, 2004]. These have been focusing mainly on the detection of breaks in the monthly, seasonal or annual values and use statistical tests accompanied with penalizing functions [Alexandersson, 1986, Caussinus and Mestre, 2004, Menne and Williams Jr, 2005, Wang et al., 2007] or inspections on autocorrelation of residuals [Vincent, 1998]. Recent comparisons [Venema et al., 2013, Domonkos, 2013, Lindau and Venema, 2013] have pointed out advantages and drawbacks of the most common systems. Procedures that look for an agreement among methods (e.g. [Kuglitsch et al., 2012]) go one step further and take benefits from the reduced uncertainty in break location by looking for consensus.

Homogenization of annual or monthly averages does not automatically imply a homogenization of higher-order moments [Trewin, 2013] since the processes that generates inhomogeneities on daily

This article has been accepted for publication and undergone full peer review but has not been through the copyediting, typesetting, pagination and proofreading process, which may lead to differences between this version and the Version of Record. Please cite this article as doi: 10.1002/joc.5874

48 data-sets are non-linear, i.e. introduced inhomogeneities to the temperature measurements depend,
49 not in a linear way, on the temperature itself [Della-Marta and Wanner, 2006] and external factors
50 as cloud cover, wind strength and direction can modify extreme daily values differently than the
51 averaged conditions [Brandsma and Van der Meulen, 2008].

52 Some methods homogenize daily records simply by interpolating monthly adjustment to a daily
53 resolution via a polynomial [Vincent et al., 2002] or trigonometric regressions [Brunetti et al.,
54 2006]. While this approach assures that daily adjusted values reflect the same temporal behaviour
55 as those observed in the monthly series [Vincent et al., 2002], adjustments for the higher-order
56 moments are not guaranteed [Mestre et al., 2011]. A more advanced set of methods considers
57 the temperature distribution which, split into quantile bins, is compared with expected values
58 obtained from surrounding stations. A non-linear regression [?] or cubic smoothing splines [Mestre
59 et al., 2011] are used for the calculation of the correcting factors. Finally a method not based on
60 model parameterization or regressions is the quantile matching, which compares quantiles of the
61 distributions of measurements before and after the break and calculates adjustments by requiring
62 similarity between these distributions.

63 Ideally, adjustments are made by comparing measurements from the original and the disturbed
64 situation for overlapping periods [World Meteorological Organization, 2011]. The difference be-
65 tween these records eliminates the background climatic signal and highlights the effects of e.g. the
66 change in location. When such parallel measurements are not available, the most reliable source
67 of information about the climate background is the net of neighbouring stations being exposed to
68 the same climatic conditions as the target series [Della-Marta and Wanner, 2006, Aguilar et al.,
69 2003, Menne and Williams Jr, 2005, Venema et al., 2013].

70 Approaches to construct the reference series are weighted (or simple) averages of surrounding
71 recorded anomalies [Alexandersson and Moberg, 1997, Vincent et al., 2002, Begert et al., 2005,
72 Štěpánek et al., 2009], using a high-correlated homogeneous series [Della-Marta and Wanner, 2006]
73 or performing the *pairwise comparison* [Menne and Williams Jr, 2005, Trewin, 2013]. The calcu-
74 lation of an averaged series may incorporate the inhomogeneities of neighbouring series into the
75 reference series [Menne and Williams Jr, 2005, Della-Marta and Wanner, 2006], might get mislead-
76 ing features from uncorrelated series and can be affected by the change in number of contributing
77 series, introducing strong changes of mean and variance in the reference [Brunetti et al., 2006],
78 thus compromising the ability to represent statistical features of the climate background [Caus-
79 sinus and Mestre, 2004]. On the other hand, while the use of single series allows the analysis to
80 be independent from the changes in data availability, this approach is risky since it relies totally
81 on a series whose quality might not be certain and whose climatic features might not be consis-
82 tent with the target series. Isolation of the artificial signal with pairwise comparison, where each
83 reference series provides an estimate of the adjustment of the target series, is shown to be more
84 robust at detection undocumented changes [Menne and Williams Jr, 2009] and more reliable for
85 the calculation of estimates of the adjusting factors [Trewin, 2013].

86 The study is organised as follows: section two introduces the data set and the methods, section
87 three shows the results, a few case studies and the effects of the homogenisation on trends in
88 temperature. The study is discussed and concluded in section four.

89 2 Data & Methods

90 2.1 ECA&D Dataset

91 The European Climate Assessment & Dataset (ECA& D, [Klein Tank et al., 2002] [Klok and
92 Klein Tank, 2008] is a collection of daily station observations of currently 12 elements and contains
93 at the time of writing (July 2018) data from nearly 11100 European stations (more than 7500 with
94 temperature measurements) and is gradually expanding. ECA&D contains more than 200 tem-
95 perature series starting before 1900, but a strong increase in the number of series is found in the
96 1950s. Data from the station network at ECA&D is updated on a monthly basis using data kindly
97 provided directly by the National Meteorological and Hydrological Services (NMHSs), individual
98 researchers affiliated with a university, global data centres like the National Centers for Environ-
99 mental Information (NOAA, Asheville, US) or the synoptic messages from the NMHSs delivered
100 through the Global Telecommunication System [World Meteorological Organization, 2007].

101 Data coverage varies depending on the countries and on the time coverage (Figure 1. Early series
102 (start before 1890) are well distributed in western Europe with the exception of southern Italy
103 and northern Scandinavia. Further data-rescue work is in progress for the improvement of data
104 coverage in low density areas and periods.

105 The quality check procedure in ECA&D is documented elsewhere [ECA&D Project Team, 2012]
106 and at the time of writing simply insists on consistency between maximum and minimum temper-
107 ature, does not allow more than 5 repeating values and flags data when the difference from the
108 climatological value exceeds five times the standard deviation. A more sophisticated quality check
109 procedures for ECA&D is currently developed, as an evolution of [Štěpánek et al., 2009], based on
110 spatial consistency of measurements in cooperation with Global Change Research Institute, Brno
111 (CZ).

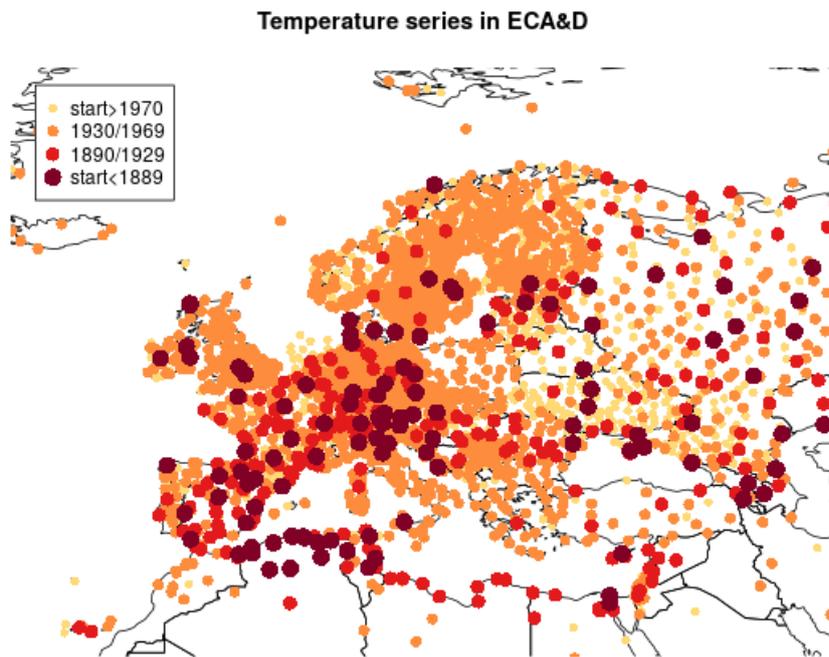


Figure 1: *Minimum and maximum temperature series in ECA&D. Colour code indicates the start of the series.*

112 2.2 Break Detection

113 The detection of breaks is done using a completely automated procedure which is blind to metadata.
114 The break detection method is inspired by the approach of [Kuglitsch et al., 2012], which seeks
115 agreement (i.e. common detected timing of breaks) of two out of three common break detection
116 methods (Prodige: [Causinus and Mestre, 2004], RHtest: [Wang et al., 2007] and GAHMDI:
117 [Toreti et al., 2012]. The main difference with the Kuglitsch et al. (2012) approach is that
118 both the selection of reference series and the combination of the three detection methods have
119 been automated. Both are performed separately on annual and winter/summer half means of
120 standardized differences between candidate and a maximum of 8 reference series, selected based on
121 completeness, correlation of annual averages (minimum of 0.6), and distance (maximum of 1000
122 km). At least three reference series must confirm a breakpoint in any of the temporal aggregations
123 in a pairwise approach.

124 The breakpoints are detected at annual resolution. Breakpoints detected in adjacent years by
125 different methods, reference series, or temporal aggregations, are considered the same breakpoint.

126 Simultaneous changes made to the measurement networks at a national scale are difficult to detect
 127 simply because surrounding reference series will suffer from the same resulting break. For these
 128 breaks documented metadata is required.

129 2.3 Calculation of Adjustments: Quantile Matching

130 The adjustment of daily temperature is inspired by the work of [Trewin, 2013] and is based on a
 131 quantile matching algorithm which compares the probability density distribution of temperature
 132 before and after the considered break, not taking into account metadata. By making use of a set of
 133 homogeneous references series, the climatic signal is accounted for and the assumption is made that
 134 the difference series between the candidate and the reference (in their homogeneous sub-periods)
 135 is random noise.

136 The adjustment process targets each series individually. The break detection produces a sequence
 137 $(t_1, t_2, t_3, \dots, t_n)$ of the timing of the detected breaks in the candidate series (from the most recent
 138 to the earliest). Following these breaks, the candidate is divided into $n + 1$ sub-series

$$\mathbf{S}_0(t|t_1 < t), \mathbf{S}_1(t|t_2 < t < t_1), \text{ etc.}, \quad (1)$$

139 which are homogeneous by definition [Caussinus and Mestre, 2004]. These segments will be con-
 140 sidered independently during the following steps of the process. Segments shorter than 5 years are
 141 not adjusted because of insufficient length required for a robust calculation of quantiles.

142 2.3.1 Reference selection and use

143 The references are selected from a box of 6° centered on the candidate station and with an elevation
 144 difference smaller than 500 meters. For high-elevation stations ($\geq 1000\text{m}$), this threshold is changed
 145 to find neighbouring stations within half the elevation of the candidate station (which increases the
 146 number of reference series for mountain stations). Among the series that fulfill these requirements,
 147 in case of densely covered areas, the set union of the 40 longest ones and the 20 starting earliest
 148 are chosen.

149 With the same splitting procedure used for the candidate series, the results of break detection are
 150 used to divide the reference series into homogeneous sub-series. Only the sub-series with at least 5
 151 years of overlap with both segments of the candidate (i.e. at either side of the break) are selected.
 152 This constraint helps to avoid series that have breaks in the same period to be references to each
 153 other (e.g. in case of simultaneous breaks in a national network). Since the presence of a trend on
 154 temperatures is likely to happen, the maximum length of the sub-series used for calculating the
 155 adjustments is set to 20 years, so that the changes in time of the moments don't alter the shape
 156 of the distribution, e.g. making it broader.

157 For each reference sub-series, the daily raw correlation with the segment of the candidate after the
 158 break is calculated. In order to limit the computational time but simultaneously preserve statistical
 159 significance, out of the set of reference series, the 18 series with the highest correlation are chosen,
 160 provided they have correlations higher than 0.75. Note that this threshold is higher than earlier
 161 suggested [Domonkos, 2013]. Figure 2 illustrates the selection of homogeneous references for a
 162 detected break in the candidate series.

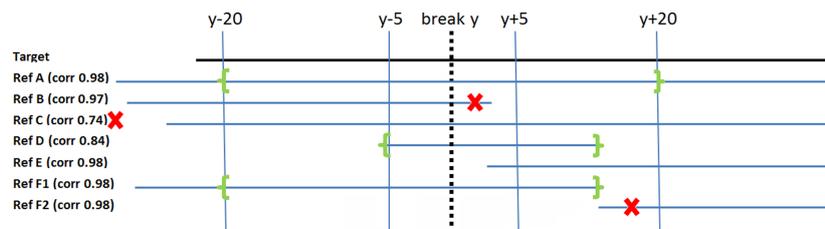


Figure 2: Fictional example of reference selection where the sections between green curly brackets are used and sections marked with a red cross are not. Reference series with correlation below 0.75 are discarded, as well as series with overlap shorter than 5 years or data recorded more than 20 years before or after the break.

163 In areas with a sparse network, which are common for the early periods, the number of available
 164 references may be low. In these cases non-split (and thus inhomogeneous) series (up to 5 in total,
 165 meeting the correlation, geographical and temporal overlapping requirements) are added to the
 166 reference set, avoiding stations having a sub-segment already selected. In any case a minimum
 167 of 3 references series is required for the procedure to be performed, otherwise the candidate is
 168 temporarily discarded. The larger size and density of the station network of ECA&D in comparison
 169 to the Australian dataset of [Trewin, 2013] makes that the availability of reference series is higher
 170 for the European situation.
 171 Note that the selection of reference series for the adjustment calculation is different than the one
 172 used for the break detection.

173 2.3.2 Calculation of quantile-based adjustments

174 Adjustment calculation has been developed taking inspiration from [Trewin, 2013]. It is performed
 175 backwards considering the breaks from t_1 to t_n successively. The segment of the candidate se-
 176 ries after the considered break (t_i) is termed the *basis* series (\mathbf{B}), while the *segment* (\mathbf{S}_i , where
 177 $i=1,2,\dots,n$) immediately before it is to be adjusted. For each reference \mathbf{R}_j , we define $\mathbf{R}_j^{\text{aft}}$ and $\mathbf{R}_j^{\text{bef}}$
 178 as the portion of the reference after and before the break t_i .

179 The quantile-based adjustments for daily data are calculated on a monthly base and applied on
 180 a daily resolution, depending on which quantile of the monthly distribution the daily data belong
 181 to. The distribution of temperatures is then considered for each month separately, introducing the
 182 seasonal cycle in the adjustments. Absolute temperatures from the target month m are considered,
 183 and values from the preceding month and the following month are gathered to reduce the noise
 184 and to make sure that a sufficient amount of data are available to determine the quantiles. This
 185 approach makes that weather types from the spring and autumn transition seasons, like March
 186 and May, contribute to determining the adjustment for homogenization of April temperatures,
 187 which is likely to be influenced both by typical March and May conditions. These temperature
 188 measurements are sorted in ascending order and e.g. the value associated to 10th quantile is
 189 calculated as the median value of all data points between the 7.5th and 12.5th quantiles. This
 190 process generates quantile sequences for data before and after the break in the target ($\mathbf{s}_{q,m}, \mathbf{b}_{q,m}$)
 191 and in the reference series ($\mathbf{r}_{j,q,m}^{\text{bef}}, \mathbf{r}_{j,q,m}^{\text{aft}}$) (thus obtaining 4 quantile sequences).

192 The adjustment for each of the quantiles is then calculated in a three-step approach. First the dif-
 193 ference between quantile sequences of the candidate series before and after the breaks is calculated,
 194 this difference is affected by both the artificial and the climatic signal. Secondly, to identify the
 195 climatic signal, the difference between quantile sequences of the reference series before and after
 196 the break is calculated. As third step, in order to isolate the artificial signal, the above differences
 197 are subtracted to each other. These steps are summarized by the following equation:

$$\mathbf{a}_{i,j,q,m} = (\mathbf{b}_{q,m} - \mathbf{s}_{i,q,m}) - (\mathbf{r}_{j,q,m}^{\text{aft}} - \mathbf{r}_{j,q,m}^{\text{bef}}) \quad (2)$$

198 In order to reduce noise, the adjustments are smoothed using a simple mean of adjustments from
 199 the neighbouring months and neighbouring quantiles.

$$\bar{\mathbf{a}}_{j,q,m} = \frac{\mathbf{a}_{j,q,m} + \mathbf{a}_{j,q+5,m} + \mathbf{a}_{j,q-5,m} + \mathbf{a}_{j,q,m+1} + \mathbf{a}_{j,q,m-1}}{5} \quad (3)$$

200 Further check has been performed to avoid situations in which negative slopes of the smoothed
 201 sequences cause, after their application, changes in the rank of the data (i.e. data changing quantile
 202 after homogenization). This check has interested a very small portion of the series, more details
 203 on this can be found in Appendix A

204 As mentioned above, the application of adjustments is performed considering each daily value of
 205 the series individually, depending on the location in the monthly temperature distribution.

206 A set of estimations of the correction is produced, each one corresponding to the different over-
 207 lapping periods each reference series \mathbf{R}_j has with the segments of the candidate series. The value
 208 to be corrected may belong to a different quantile in each of these overlapping periods. After
 209 determining these quantiles (\tilde{q}_j) the estimation (\tilde{v}_j) of the adjusted value related to \mathbf{R}_j is:

$$\tilde{v}_j = v + \mathbf{a}_{j,\tilde{q},m} \quad (4)$$

210 where v is the original value. The final adjusted value is then calculated taking the median of the
211 estimations:

$$\bar{v} = \text{median}(\tilde{v}_j) \quad (5)$$

212 where $j=(1,\dots,r)$.

213 The method described above has been applied to the whole ECA&D dataset. The high number of
214 breaks detected caused a high number of short homogeneous segments, which often were not long
215 enough to be homogenized or had too short overlapping period with the homogeneous segments
216 of the surrounding reference series, making it impossible to perform the quantile matching. These
217 portions have been integrated in the temporary homogenized version of the dataset, which has
218 undergone a second round of homogenization. In this second run the break detection and quantile
219 matching have been launched again, so that additional adjustments were calculated.

220 3 Results

221 3.1 Statistics of the adjustments

222 Figure 3 shows the number and timing of the detected breaks for the first iteration and for the
223 second iteration. The relatively high number of breaks detected in the original series during
224 the first iteration resulted in a high number of short homogeneous segments. These segments,
225 serving as references, are often not long enough to be homogenized or had a too short overlapping
226 period with the target series. In such a situation no adjustments are possible. Nevertheless, the
227 first homogenization iteration improved the number and the length of the homogeneous segments
228 which made it possible to adjust additional breaks in the second iteration. The number of series
229 for daily maximum and minimum series in the ECA&D dataset and the number of adjusted series
230 after the first and second iterations are shown in Table 1.

231 Since the second iteration takes the results of the previous iteration as input, the possibility exists
232 that this complex system diverges too strongly from the initial situation (e.g. positive feedbacks
233 in the iterative processes, tendency for the removal of local signals or introduction of forced trends
234 in the temperature series). This issue might be alleviated by applying the homogenization on the
235 original series, using the homogenized series as references in each iteration and setting a convergence
236 threshold to stop the system. In the approach documented here, we simply limit the number of
237 iterations to two.

238 The adjustments that have been applied to the breaks vary strongly with the month and the
239 quantiles. Figure 4 shows that the adjustment of the median is symmetric around -0.1°C and smaller
240 (and less frequent) adjustments are found for the second iteration. Furthermore the distribution
241 of adjustments of the median is wider for minimum temperatures and more peaked around for
242 maximum temperature. Averages of adjustments for the 5th, 50th and 90th quantiles (see table 2)
243 are more negative for TN. For both variables adjustments for lower quantiles are more negative,
244 indicating a general tendency to broaden the probability density distribution. This is consistent
245 with earlier findings [Lawrimore et al., 2011, Trewin, 2013, Thorne et al., 2016].

246 The peak in the distribution close to zero relates to:

- 247 1) the independence between break detection and adjustment calculation, i.e. a break may be
248 found but the comparison to the surrounding reference series does not give a reliable correction,
- 249 and 2) to the possibility that the median needs no adjustment but percentiles in the tails do.

250 This latter situation is illustrated in the scatterplots of adjustments for the 5th versus the 95th
251 quantiles, for the series where adjustments of the median are null (figure 5). These figures show
252 a centred and symmetric distribution and indicate that adjustments are not skewed towards more
253 positive or negative slopes. These figures also show that no thresholds on the absolute value of the
254 adjustments are used, contrasting with the approach of [Trewin, 2013] who applied adjustments
255 only if the resultant shift in annual mean exceeded the $\pm 0.3^\circ\text{C}$ threshold.

	TN	TX
Original series (complete set of considered series)	6438	6404
Homogeneous original series (original series labeled as homogeneous by the break detection)	560	670
Adjusted series it.1 (series that have been corrected during it.1)	2111	2007
Homogenized data-set it.1 (union of series that were already homogeneous and adjusted series)	2671	2677
Homogeneous series after it.1 (union of original hom. series and series successfully homogenized after it.1)	1165	1131
Adjusted series (it.2) (series that have been corrected during it.2)	1526	1571
Homogenized data-set (final) (union of series that were already homogeneous and adjusted series)	2691	2702
Final non homogeneous series (according to a third run of break detection)	1357	1260
Final homogeneous series (according to a third run of break detection)	1334	1441

Table 1: *Number of series involved in the stages of the homogenization process. Homogenized series, for each iteration, consist of the sum of the adjusted series and the already homogeneous series. Low percentage of homogenized series is due to the exclusion of short and duplicate series.*

	TN			TX		
	05 th perc	50 th perc	95 th perc	05 th perc	50 th perc	95 th perc
1 st it.	-0.12 °C	-0.11 °C	-0.10 °C	-0.06 °C	-0.05 °C	-0.06 °C
2 nd it.	-0.03 °C	-0.01 °C	-0.01 °C	-0.04 °C	-0.03 °C	-0.03 °C

Table 2: *Averages of adjustments for 5th, 50th and 95th percentile for TN and TX in first and second iteration*

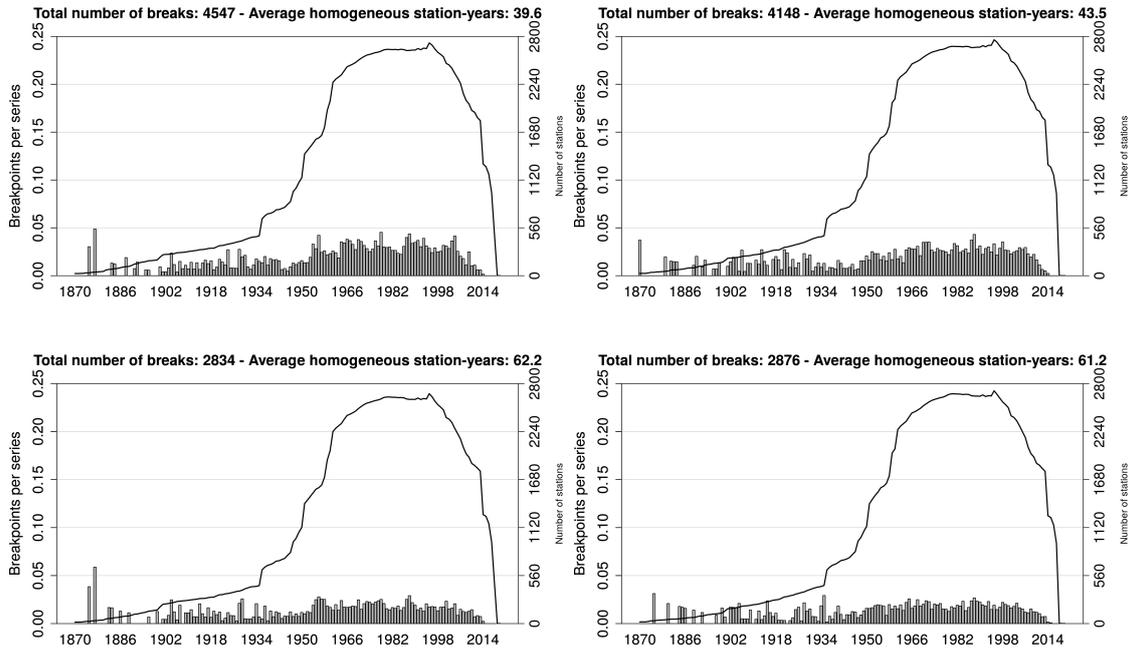


Figure 3: Statistics regarding application of break detection on ECA&D temperature data-set. Histogram describes number of breakpoints per series, line describes number of stations. Left (Right) panels are about minimum (maximum) temperatures. Top and bottom panels are respectively for first and second break detection runs.

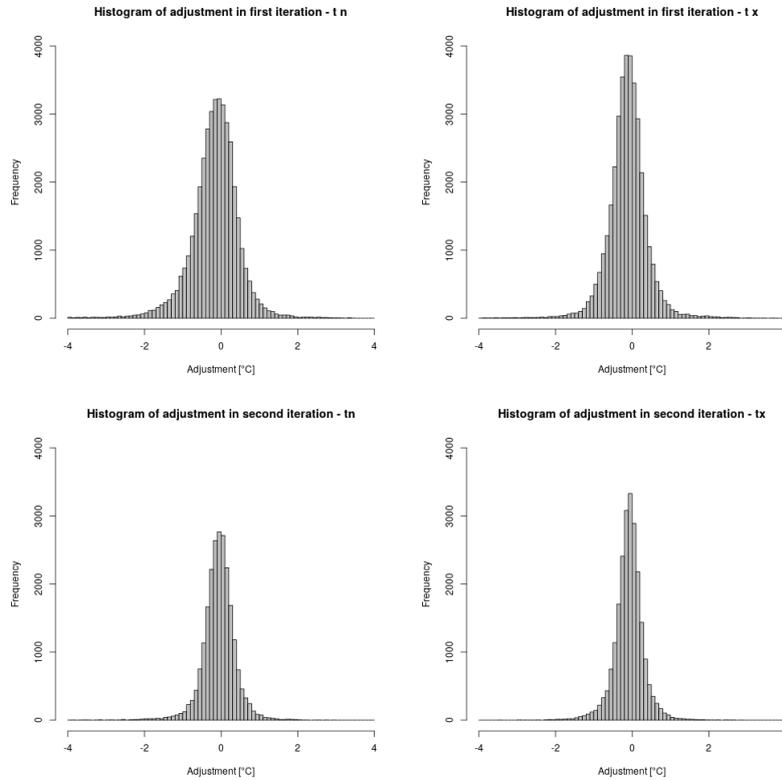


Figure 4: Histograms of adjustments for values in the median quantile (q_{50}), for t_n (left column) and t_x (right columns) and for first iteration (top row) and second iteration (bottom row). Difference in width between first and second iteration proves the different role of the two phases.

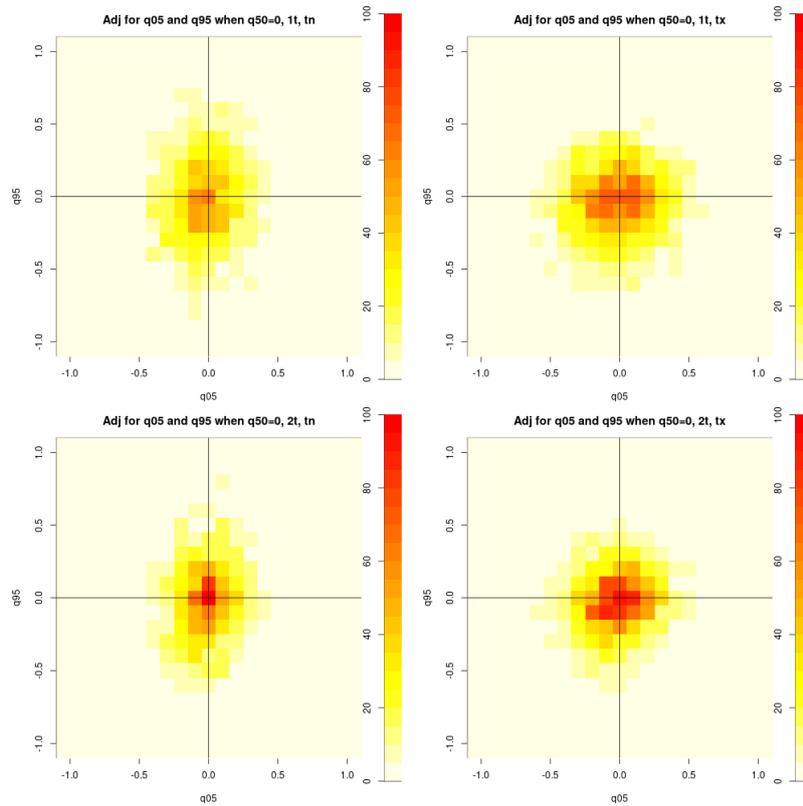


Figure 5: *Density scatterplot of adjustments for quantile 95 versus quantile 05 when adjustments for the median are null. Minimum temperature (left column), maximum temperature (right column), first iteration (top row), second iteration (bottom row).*

3.2 Case studies

In order to demonstrate the method in more detail, two case studies are presented.

3.2.1 Bamberg

An illustrative example is the adjustment of data from the station of Bamberg (Germany). Metadata reports a set of breaks (Table 3) which are only partially retrieved by the automatic break detection (first iteration: 1891,1952; second iteration: 1920).

The two documented breaks are not reproduced exactly, but the 1948/49 break is located within a few years. The further breaks that are detected are probably related to unrecorded changes in the features of the station.

1880-01 to 1948-12	Bamberg (Sternwarte - City Centre)	z=283m
1949-01 to 1995-03	Bamberg (South - Country Side)	z=243m
1995-03 to present	Bamberg (South - Country Side)	z=239m

Table 3: *Available metadata regarding station in Bamberg, Germany.*

The high density of stations in Germany and Austria provided by their respective meteorological services allows to have more than 18 reference series available for the break in 1952 on which we focus in this case study. The 18 highest correlated ones have been selected (figure 6).

Shape and location of the probability distributions of the non-homogenized temperatures before (light red) and after (light blue) the break show a clear distinction (figure 7). Shifts in quantile sequences varies from very low values for the tails of the distribution to 1.1°C for the median, showing the different effect of the break on mean and extreme values. Probability plots after the

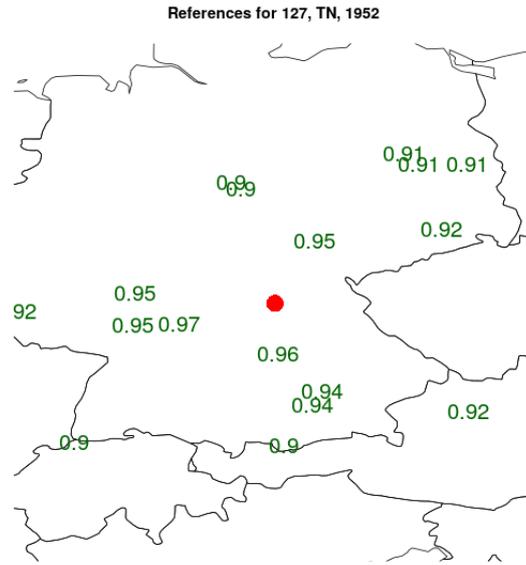


Figure 6: Reference series that have been used for the homogenization of the break in 1952 in the station of Bamberg (red dot). Numbers represent correlation calculated between references and basis series (considering the 20 years following the break).

272 two iterations of homogenization (red before and blue after the break) get closer to each other in
 273 different way depending on the quantiles, indicating that the difference between the two original
 274 distributions was not entirely due to the artificial intervention. The two sub-series (before and
 275 after the break) do not completely overlap due to the climatic variability that has been captured
 276 by the surrounding reference series and taken into account in determining the adjustments.
 277 Estimates of adjustments related to each reference (for this case study) are shown in figure 8. These
 278 are the results of the process described by equation 2, followed by smoothing and check of negative
 279 slopes. Figure 8 shows that the lower quantiles has stronger (more negative) adjustments than the
 280 upper quantiles which increases the width of the distribution. As an example of the adjustment
 281 process, the estimates related to the value measured on 22th May 1951 (6.2°C) are highlighted
 282 in figure 8. This measurement belongs to different quantiles (35th and 40th), depending on the
 283 reference series that is considered, since for each of these there is a different overlapping period with
 284 the target series. The final correction is taken as the median of the estimates, in this particular
 285 case the adjustment will be -1.3°C, with a final homogenized value of 4.9°C.

286 Effects on the series are evident when indices like the annual mean are plotted (figure 9, top panel).
 287 In this particular case, the first iteration is able to correct the series almost entirely, since the break
 288 that has been detected during the second iteration (1920) had very low adjustments. Comparison
 289 of corrections for the mean and for the two tails of the distribution show the expected differences:
 290 larger (smaller) corrections for the 5th (95th) quantile.

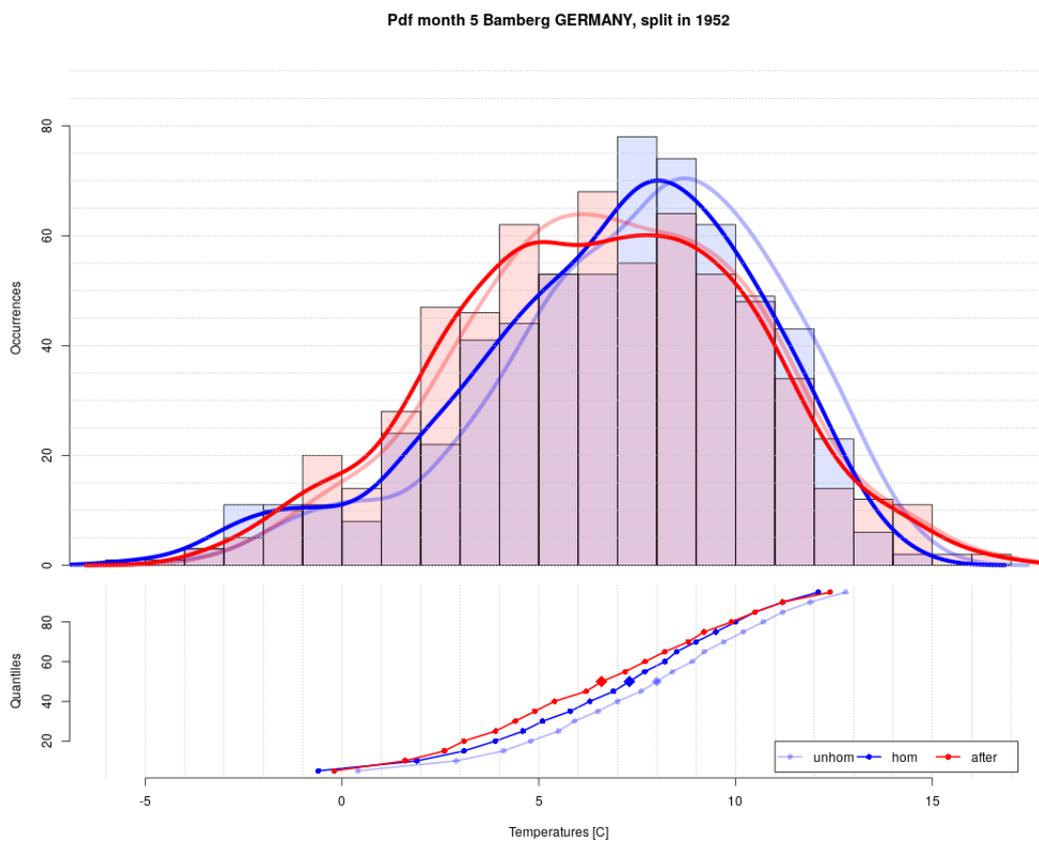


Figure 7: *Top: Histograms and pdf of adjusted minimum temperature in month May for the 20 years before 1952 (blue) and 20 years after that (red). Light blue and light red curves represent original pdfs. Bottom: quantile functions related to the above distributions, same colour code.*

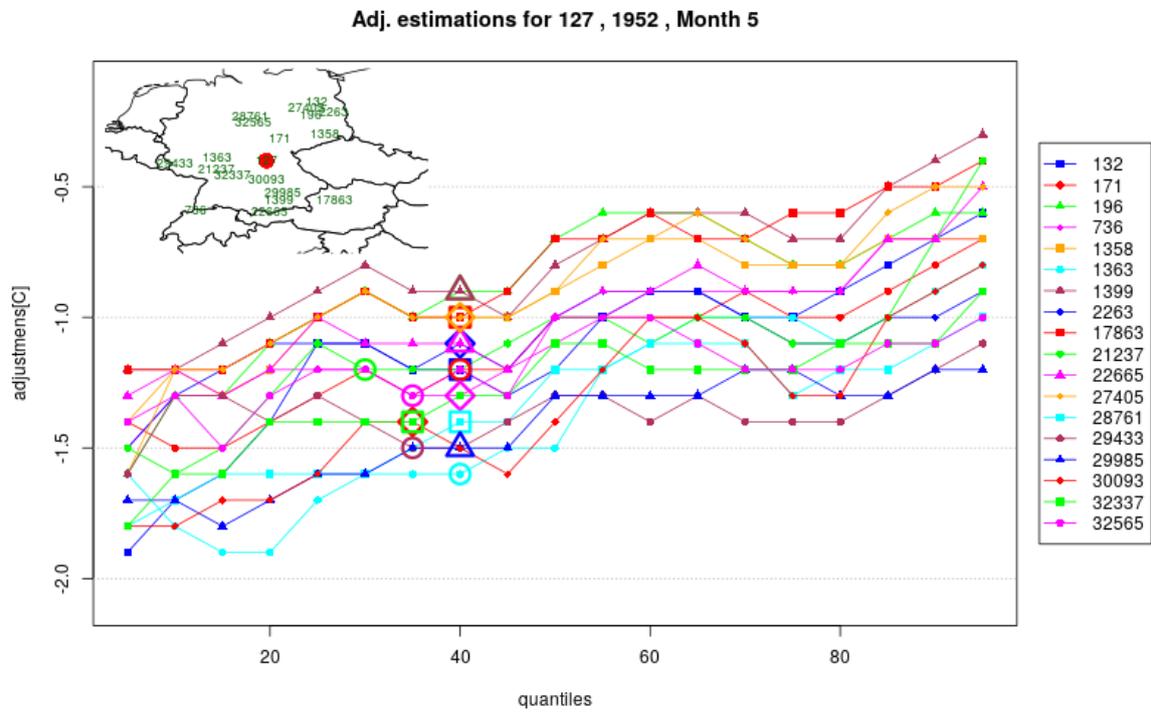


Figure 8: Estimation of adjustments for month May , station of Bamberg and break in 1952 after the smoothing process and the negative slop check.

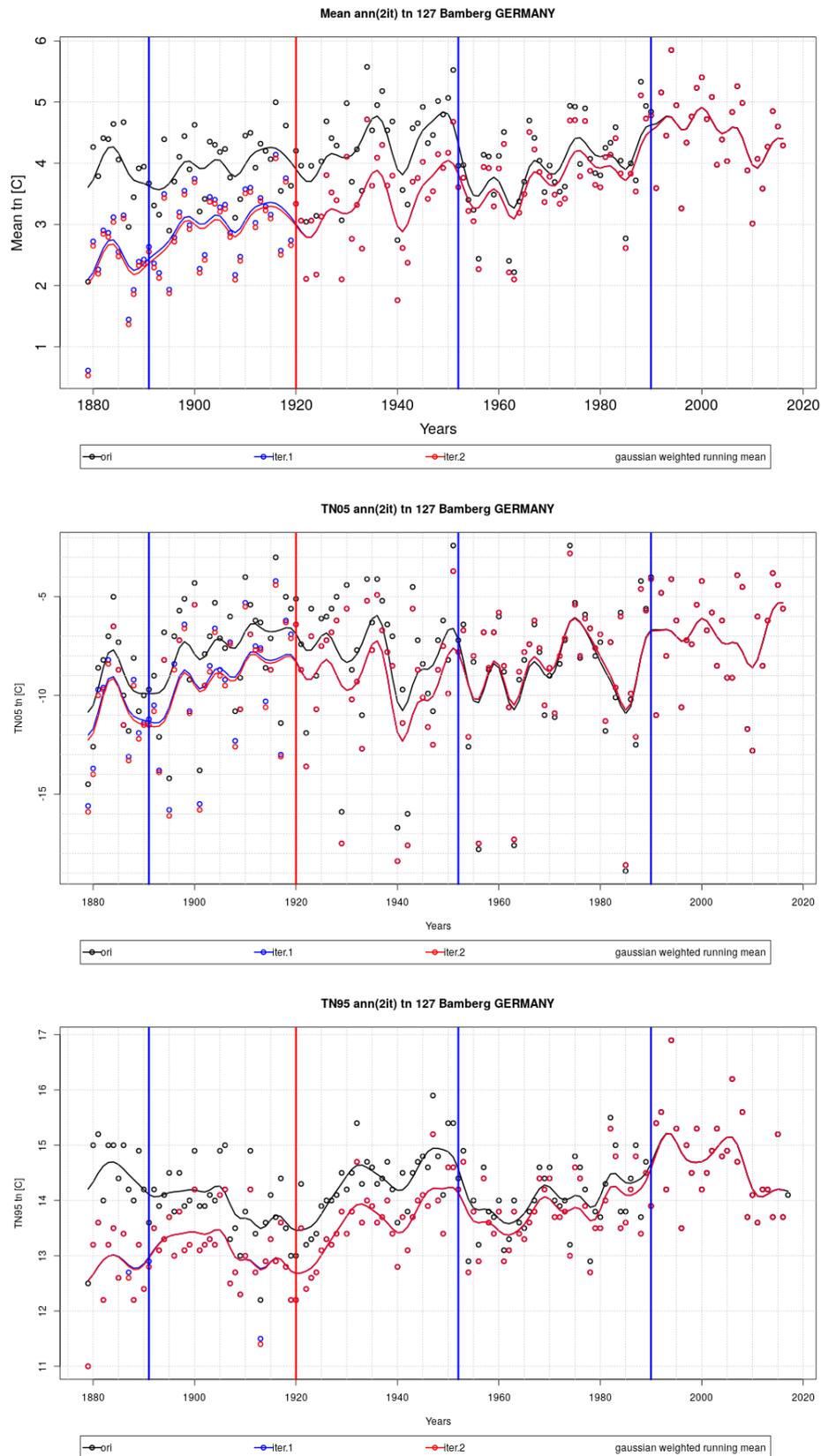


Figure 9: Annual mean (top), 5th quantile (centre), 95th quantile (bottom) time series for minimum temperatures in Bamberg. Black line: original series, blue: first iteration result, red: second iteration result. Vertical lines: output of break detection in the first (blue) and in the second (red) iteration.

291 **3.2.2 Salzburg**

292 A particularly representative case is the station of Salzburg (Austria), where the metadata reports
 293 a set of breaks (table 4). Break detection detects almost all these breaks, except the most recent
 294 one which is probably small in amplitude (only 3 meters of change in height), and detecting some
 295 further breaks which probably derive from unreported changes in the stations features. For this
 296 case study, focus will be on the break located near 1938 which is associated with the relocation of
 297 the station from the city to the nearby airstrip.

1863-01 to 1883-12	High School (Gymnasium Altes Borromäum)	z=424m
1884-01 to 1903-07	High School (Oberrealschule)	z=419m
1903-08 to 1941-02-28	Studiengebäude-Lehrerbildungsanstalt	z=423m
1939-03-01 to 1996-06-15	Airport station 1	z=434m
Since 1996-06	Airport station 2	z=437m

Table 4: Available metadata for station in Salzburg, Austria.

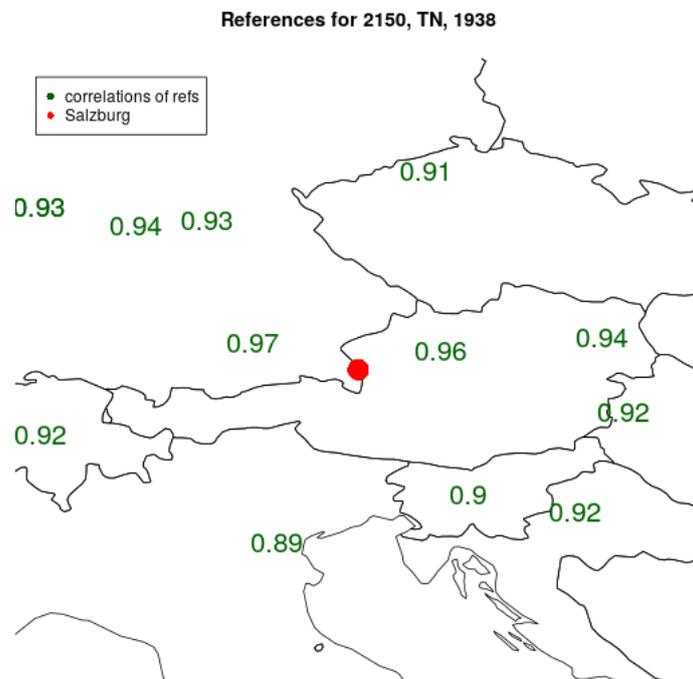


Figure 10: Reference series that have been used for the homogenization of the break in 1938 in the station of Salzburg (red dot). Numbers represent correlation calculated between references and basis series. Correlations are calculated using the 20 years following the break.

298 For the 1938 break, 12 reference series meet the requirements, see figure 10 . Shape and location
 299 of the probability distribution of temperatures before and after break in 1938 (figure 11) show a
 300 clear shift of the distribution of records before the break. Shift in quantiles varies from 0.6°C for
 301 the 5th quantile to 1.1°C for the median, showing the different effect of the break on mean and
 302 extreme values. Probability distributions after the 2 iterations of homogenization almost overlap
 303 each other, indicating that a great part of the difference was due to the relocation, while the
 304 remaining difference represents actual climatic variability that relates to the surrounding reference
 305 series.

306 Adjustments applied may be seen in figure 12, analogous to figure 8 with the difference that here
 307 the slope of the curve in the quantile-adjustment plot is not as steep as in the Bamberg case. The

308 highlighted marks are related to the measurement on 14th May 1938 (4.9°C), whose adjustment
 309 will be -0.7°C, with a final value of 4.2°C.

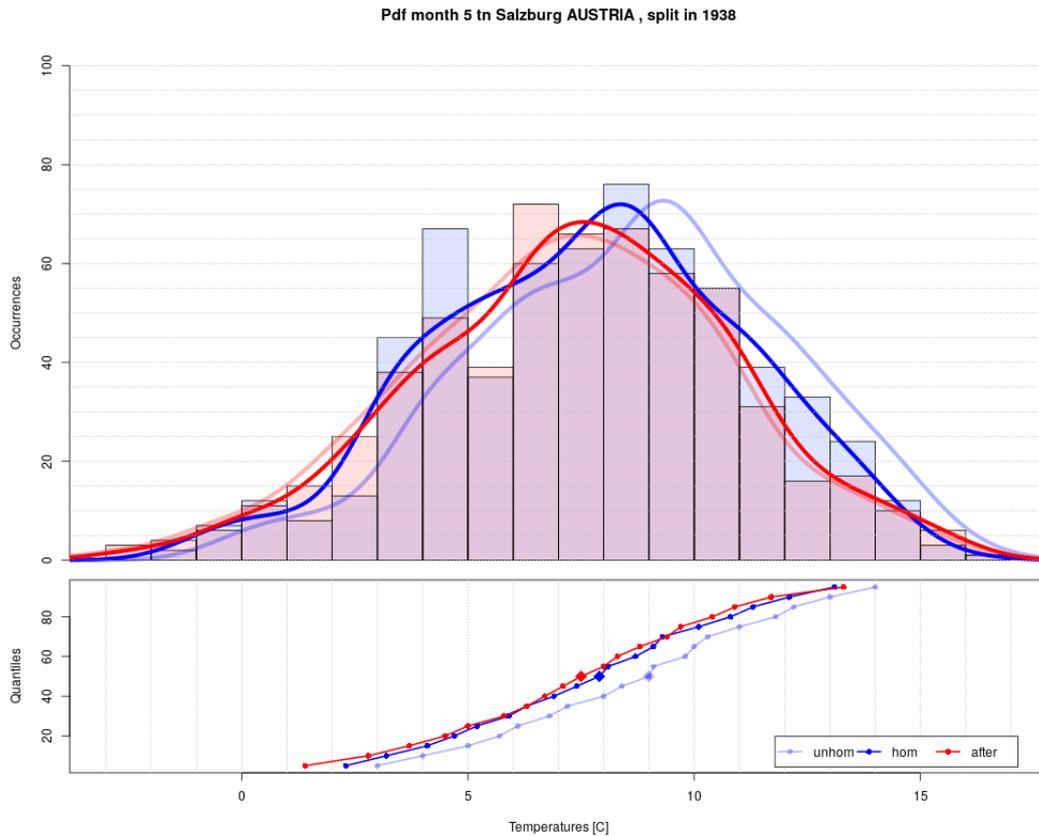


Figure 11: *Top: Histograms and probability distribution of adjusted minimum temperature in month May for the 20 years before 1938 (blue) and 20 years after that (red). Light blue and light red curves represent original distributions. Bottom: quantile functions related to the above distributions, same colour code.*

310 Effects on the series are evident when indices like the annual mean are plotted (figure 13, top
 311 panel). Interesting about this case is that first iteration (blue lines, when not covered by red)
 312 corrects the big breaks, such as the break in 1938. On the other hand, the second iteration is
 313 able to adjust two early breaks (red vertical lines) that were not detected during the first round
 314 because of the lack of long reference series in the early periods. The two new breaks are confirmed
 315 by the metadata (table 4). The amplitude of adjustments in this case is clearly lower, showing
 316 that second iteration works as an enhancement of adjustments from the first round. Appendix B
 317 shows that the second iteration homogenizes the older part of the series that was not corrected
 318 during the first round.

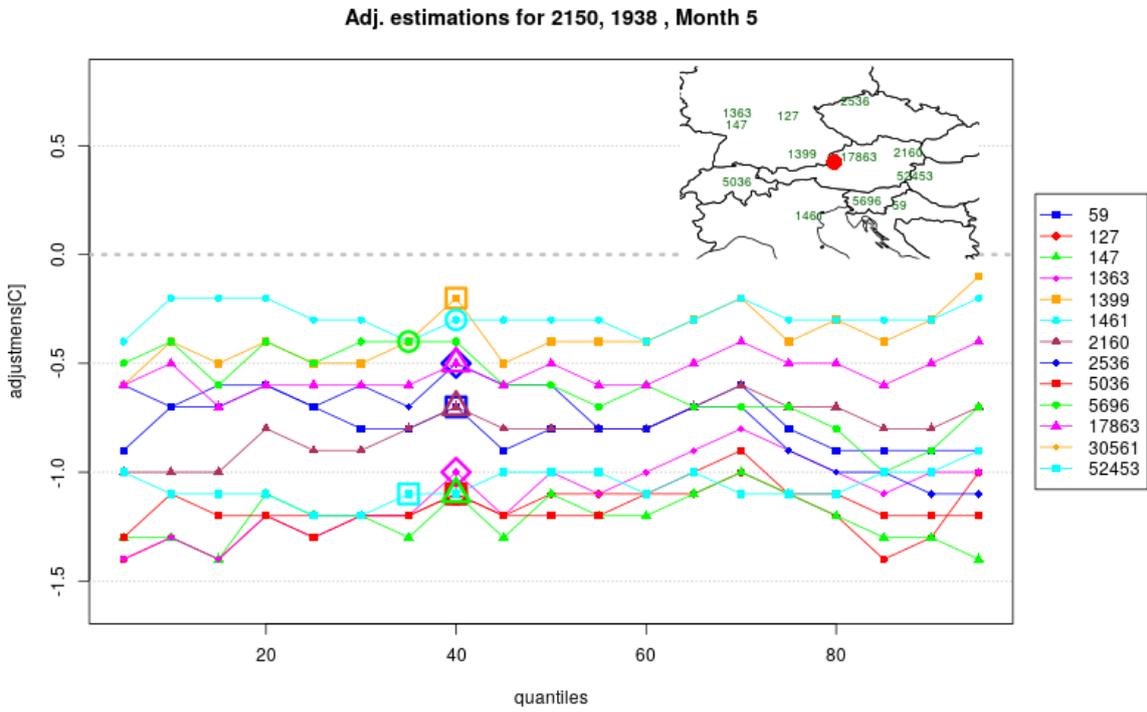


Figure 12: Estimation of adjustments for month May, station of Salzburg and break in 1938 after the smoothing process and the negative slope check. The inset shows the locations of the series used to calculate the adjustment.

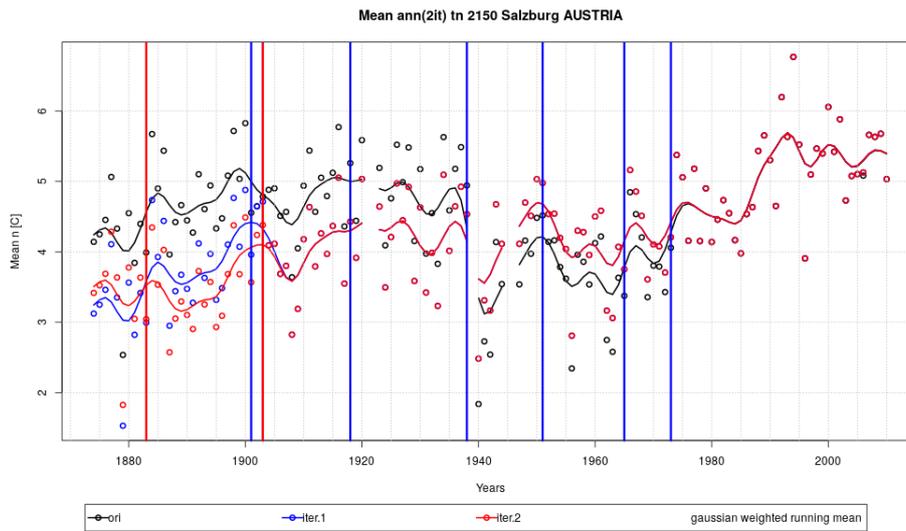


Figure 13: Annual mean time series for minimum temperatures in Salzburg. Black line: original series, blue: first iteration result, red: second iteration result. Vertical lines: output of break detection in the first (blue) and in the second (red) iteration.

3.3 Application to the complete data-set

Figure 14 shows trends in annual mean daily minimum and maximum temperature over the 1961-2010 period, before the homogenization (top panels) and after the homogenization (middle panels). The comparison of these figures shows the removal of several outliers and unrealistic low or high trends values. The trends based on the homogenized series are much more spatially homogeneous. The bottom panels of figure 14 show the difference in trend values between the non-homogenized and homogenised series which demonstrate that the adjustments go both ways - trends are increased and decreased by this procedure. A coherent spatial pattern of adjustments is not evident from this figure.

The almost complete disappearance of stations with negative and very large trends demonstrates the effectiveness of the method in recognizing and keeping the climate signals that dominates the series and removing outliers trends which are related to artificial signals. Even though the result is the convergence of the trends of all stations to positive values, it is important to notice that the aim of this process is not the removal of the negative trends. This phenomenon is an indirect effect of the homogenization procedure. Indeed all stations with excessively high trends (i.e. dark red circles) have been adjusted with negative factors, as shown in figure 14, bottom row. A further check is to search for stations that still showed a negative trend or a exceptionally high trend exceeding $0.6^{\circ}\text{C}/\text{dec}$. Figure 15 shows the locations of the stations related to these extreme trends. These are not isolated stations but it is shown that these values are consistent with trends of neighbouring stations. The low value trends are mainly located in Bulgaria and southern Romania, while the very large trends are mainly in the Northern Baltic area. The second case is likely to be the result of a widespread climatic effect, while the first might be the result of the influence of the series of Bucarest on the neighbours.

In Appendix C box plots show the distribution of the trends in the annual mean of TN and TX for the two successive iterations. These indicate a narrowing of the distribution of the trends together with a significant reduction of outliers.

Figure 16 describes the distribution of the trends on extreme indices (5th percentile of TN and 95th percentile of TX) in the original data-set and after the two stages of homogenization.

Beside the changes in the width of the distributions, changes in the first moment have been observed. The medians show a slight shift to higher values (table 5) for annual means. TN05 and especially TX95 show a more irregular behaviour, where overadjusted trends in the first iteration are refined in the second iteration.

	TN annmean	TX annmean	TN05	TX95
original	+0.31 $^{\circ}\text{C}/\text{dec}$	+0.35 $^{\circ}\text{C}/\text{dec}$	+0.41 $^{\circ}\text{C}/\text{dec}$	+0.35 $^{\circ}\text{C}/\text{dec}$
1 st iteration	+0.31 $^{\circ}\text{C}/\text{dec}$	+0.37 $^{\circ}\text{C}/\text{dec}$	+0.42 $^{\circ}\text{C}/\text{dec}$	+0.40 $^{\circ}\text{C}/\text{dec}$
2 nd iteration	+0.32 $^{\circ}\text{C}/\text{dec}$	+0.37 $^{\circ}\text{C}/\text{dec}$	+0.42 $^{\circ}\text{C}/\text{dec}$	+0.37 $^{\circ}\text{C}/\text{dec}$

Table 5: Median trend values of annual averaged daily minimum and maximum temperature, of TN05 and of TX95, for the original dataset, after the first iteration and after the second iteration.

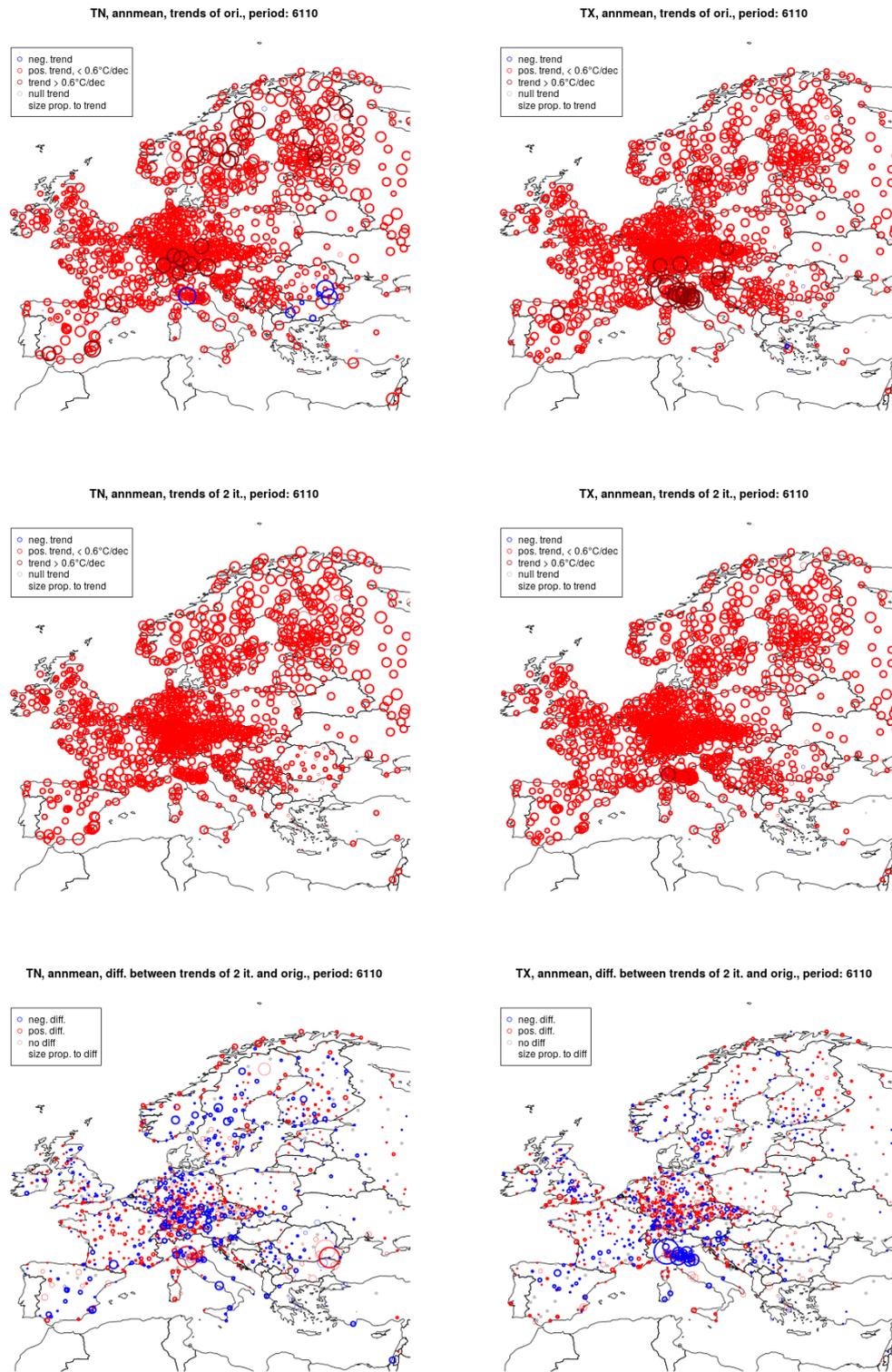


Figure 14: Maps of trends of annual mean in the period from 1961 to 2010 of original series (top), homogenized series (middle) and difference between the two (bottom) about minimum (left column) and maximum temperatures (right column). Blue circles indicate negative trends, red circles represent positive trends below $0.6^\circ\text{C}/\text{dec}$, dark red circles represent trends above $0.6^\circ\text{C}/\text{dec}$. Size of the circle is proportional to the amplitude of the trend. Thickness of the circle indicates significance of the trend itself (above 0.95). Code colour is chosen basing on the box plots of figure 18 (blue and brown values lie on the tails of the original distribution).

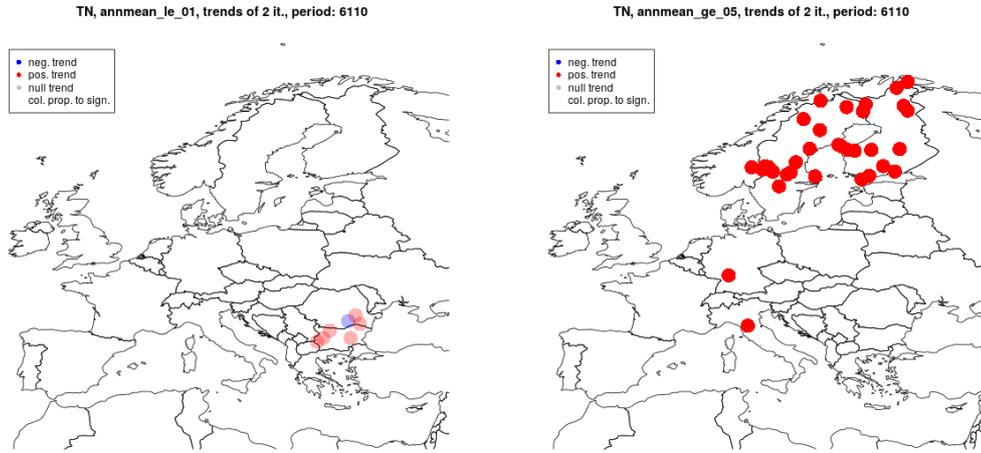


Figure 15: *Left: Map of series having trend of annual mean of minimum temperatures lower than $0.1^{\circ}\text{C}/\text{dec}$, transparency of the dots indicate non significant trend. Right: Map of series having trend of annual mean of minimum temperatures larger than $0.5^{\circ}\text{C}/\text{dec}$, transparency of the dots indicate non significant trend.*

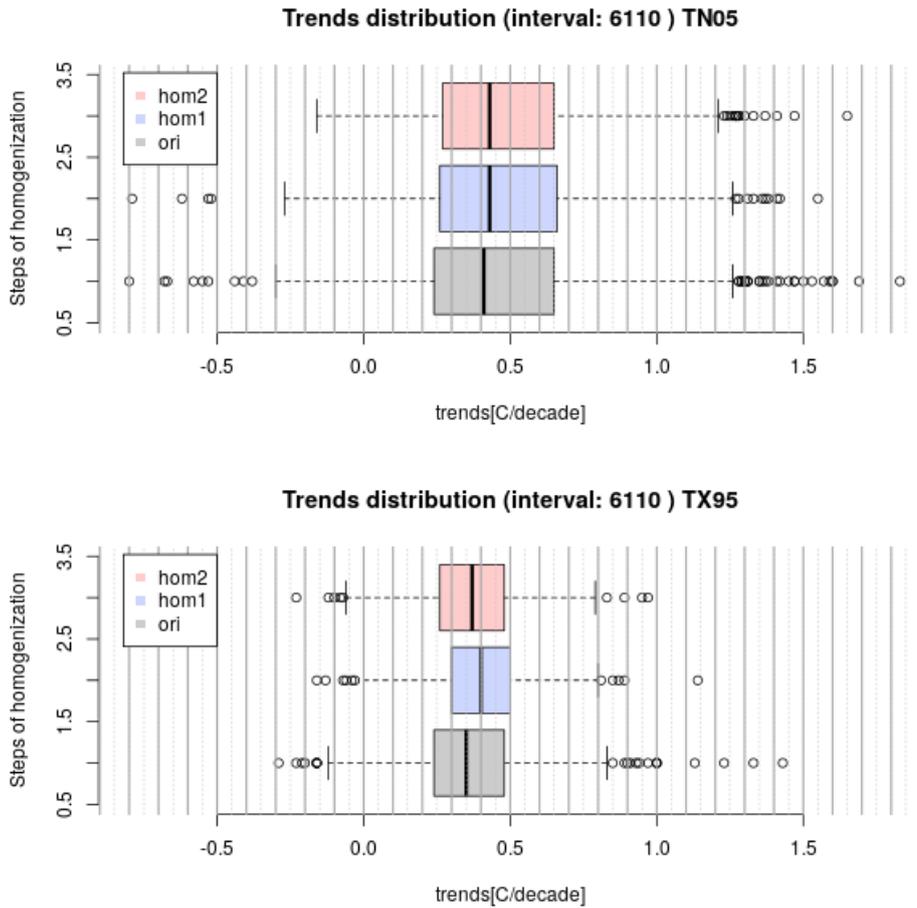


Figure 16: *Distribution of trends for TN05 (top) and TX95 (bottom). Each plot shows boxplot for original (grey), first iterations' result (blue) and second iterations' result (red)*

4 Discussion and Conclusions

A fully automatic homogenization method for daily temperature series has been presented and applied to the pan-European dataset ECA&D. The size of the dataset and the absence of detailed metadata for all series and stations requires a procedure which is *blind* for metadata and is able to handle a large variety of data quality conditions. These challenges are met by using a combination of the break detection method developed earlier by [Kuglitsch et al., 2012] and the quantile matching method which has been pioneered earlier by [Trewin, 2013] for use in large datasets. In order to distinguish between climatic signal and artificial signal in the breaks, a network of reference series in the vicinity of the record that needs adjustment is employed. In this study, the reference series are chosen from coordinate boxes of $6^\circ \times 6^\circ$ around the target series, with thresholds on altitude difference. A further selection, using daily raw correlation and length of overlapping period, makes this approach sufficiently flexible to cope with both high and low station density areas. These criteria take inspiration from the the nearest neighbour stations approach of [Menne and Williams Jr, 2009] and [Trewin, 2013].

The whole procedure is iterated twice. While the first iteration locates and adjusts the largest discontinuities, the second iteration is able to adjust more subtle changes such as earlier breaks, breaks with smaller amplitude or in areas with scarce station density. The more abundant presence of homogeneous sub series after the first iteration makes this possible. The homogenization has been able to adjust about 2700 ECA&D temperature series, while the remaining ones are duplicates or consist in short records. A final round of break detection has shown that only 1400 series can be considered completely homogenized, while on the rest minor breaks persist. No further iterations are made to remain as close as possible to the original dataset while adjusting for the largest inhomogeneities.

The trends in annual averaged values show a much stronger spatial consistency than before the adjustments. This illustrates the effectiveness in the removal of the artificial signals, thus making the climatic signal dominant. A comparison between trends prior and after the homogenization shows that changes in trends are both ways. The averages of the distributions of European trends of annual means are shifted slightly to warmer values (TN: +3.2%, TX: +5.7%). At the same time the interquantile range of the distributions of these trends are consistently reduced (TN: from 0.16 to 0.10°C/dec, TX: from 0.12 to 0.08°C/dec), indicating a higher uniformity of the values. Similar conclusions are reached when considering indices for extreme values, such as the 5th and 95th quantiles.

The strength of the quantile matching method is that each part of the distribution (i.e. quantile) is considered independently from the others. Previous studies on temperature probability distributions have focused on fitting the probability density functions with sophisticated functions or calculation of variations in the distribution parameters. On the other hand the quantile matching approach has a more heuristic approach, aiming at being more versatile and able to adapt to the wide spectrum of signals that artificial activities may lead to the records. In distinction with the [Trewin, 2013] approach, no linear interpolation between quantiles (used to obtain adjustments for percentiles between the multiples of 5) is included, reducing the parameterization of the process. Furthermore the calculation of adjustment is more conservative (use of averaged values and check of negative slopes) and the selection of reference series employs different criteria to give more importance to data availability and correlation.

Nonetheless further studies have been performed and are planned to understand minor controversial aspects of the described method. The dispersion of the reference series has been shown to affect the calculation of the adjustments (Appendix D). Therefore it is planned to inspect how to lend the reference selection an "angular even distribution", i.e. approximately same number of references on the North, on the South, on the East and on the West of the candidate series. The selection of reference series must also take into account the contribution of series with anomalous behaviour, as seen in Figure 15, where the negative trend in a station (Bucarest) might be one of the reasons of the lower trends observed in the surrounding area. The validation and the comparison of the results with other adjustment calculation methods are currently subject of further studies.

In conclusion, the method to adjust inhomogeneities in daily temperature discussed here is a purely statistical method. While the use of quantile matching favours the differentiation of adjustments for low and high daily extremes and have a seasonal cycle, these adjustments don't consider existing meteorological circumstances. Future work using a physical approach to calculating the adjust-

407 ments, in which actual weather contributes to the size of the adjustment, will give an alternative
408 estimate of the homogeneity adjustment.

409 5 Acknowledgements

410 We acknowledge the data providers in the ECA&D project (www.ecad.eu). Funding has been
411 received from the EU FP7 Collaborative Project UERRA (Uncertainties in Ensembles of Regional
412 ReAnanalysis), Grant agreement 607193, and the EU H2020 EUSTACE Project, Grant agreement
413 640171.

414 A Check of negative slopes in the adjustment sequences

415 During the application of the quantile matching method it might happen that the rank of mea-
416 surements is not preserved. This occurs if the adjustment of a high quantile is smaller than that
417 of a lower quantile.

418 This possible setback, that involves approximately 0.5% of the adjustment calculation, requires
419 a constraint in order to keep the rank of data when the sequence has a negative slope in the
420 adjustment - quantile plane. By definition a quantile sequence, including the result of the adjusting
421 process ($\tilde{s}_{j,q,m}$) must have a non-negative slope. This implies that for any q :

$$\bar{s}_{j,q+5,m} - \bar{s}_{j,q,m} \geq 0 \quad (6)$$

422 For each q , elements of the adjusted quantile sequence are calculated like:

$$\bar{s}_{j,q,m} = \mathbf{s}_{j,q,m} + \mathbf{a}_{j,q,m} \quad (7)$$

423 Thus:

$$\begin{aligned} & \tilde{s}_{j,m,q+5} - \tilde{s}_{j,m,q} = \\ & = (\mathbf{s}_{j,q+5,m} + \bar{\mathbf{a}}_{j,q+5,m}) - (\mathbf{s}_{j,q,m} + \bar{\mathbf{a}}_{j,q,m}) = \\ & = (\mathbf{s}_{j,q+5,m} - \mathbf{s}_{j,q,m}) + (\bar{\mathbf{a}}_{j,q+5,m} - \bar{\mathbf{a}}_{j,q,m}) \geq 0 \end{aligned}$$

424 And finally:

$$(\bar{\mathbf{a}}_{j,q+5,m} - \bar{\mathbf{a}}_{j,q,m}) \geq -(\mathbf{s}_{j,q+5,m} - \mathbf{s}_{j,q,m}) \quad (8)$$

425 This constraint is implemented fixing the adjustment related to the median and checking the two
426 tails quantile by quantile. In case of a too negative slope, the value is corrected moving it to the
427 closest acceptable value.

428 For instance, for the right tail of the distribution, if:

$$\mathbf{a}(\tilde{j}, 55, \tilde{m}) - \mathbf{a}(\tilde{j}, 50, \tilde{m}) < -\mathbf{s}(\tilde{j}, 55, \tilde{m}) + \mathbf{s}(\tilde{j}, 50, \tilde{m}) \quad (9)$$

429 then the corrected value is set to:

$$\bar{\mathbf{a}}(\tilde{j}, 55, \tilde{m}) = \mathbf{a}(\tilde{j}, 50, \tilde{m}) - \mathbf{s}(\tilde{j}, 55, \tilde{m}) + \mathbf{s}(\tilde{j}, 50, \tilde{m}) \quad (10)$$

430 B Munich

431 The need and the utility of the second iteration with the break detection and homogenization can
432 be appreciated when looking data from the station of Munich (Germany). Metadata reports a set
433 of breaks (Table 6) including one in the 1920s.

434 Here the big difference between red and blue lines in figure 17 indicates that the first iteration was
435 not able to correct the breaks in 1912 and 1926. Nevertheless the higher availability of data and
436 long homogeneous segments, together with a better signal to noise ratio has allowed to adjust the
437 earliest part of the series on the second run of the software.

1879-01 to 1954-07	Munich (Botanic Garden Nynphenburg)	z=515
1954-08 to 1999-03	Munich (Nynphenburg residential area)	z=515m
1999-04 to present	No location metadata (probably not changed) (changes in measuring times in 2001-04)	z=515m

Table 6: Available metadata regarding station in Munich, Germany.

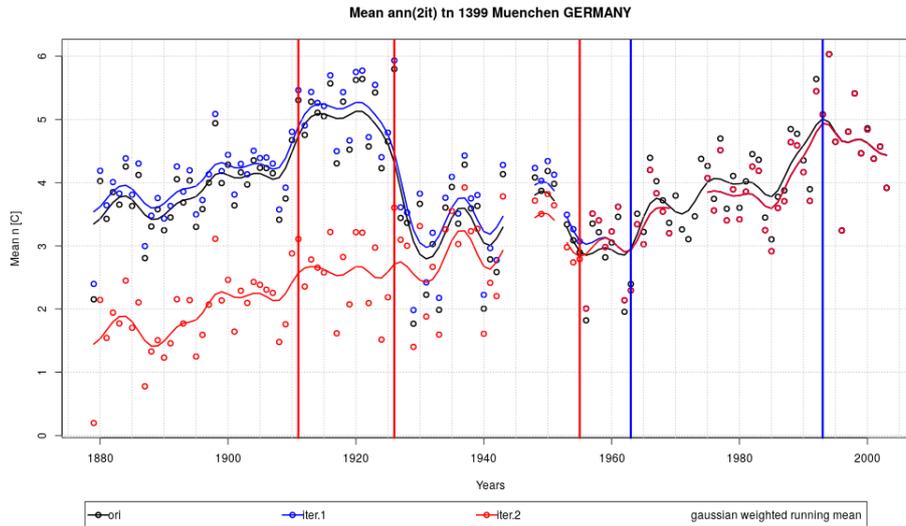


Figure 17: Annual mean time series for minimum temperatures in Munich. Same colour code as previous figures.

438 C Trend assessments on annual means

439 Assessment of trends before and after the homogenization has been computed on annual means,
440 showing a relevant narrowing of the distribution, especially between original and first iteration,
441 while the second iteration acts more like a refining of the result, see figure 18.

442 D Geographically-induced patterns on adjustments

443 Adjustment calculations may depend on the geographical distribution of reference series around
444 the target series. The homogenization of the break in 1938 of series of Salzburg (Austria) has been
445 checked considering separately the reference series that lay on the north (south, west and east), see
446 figure 19, top panels. Both pairs north-south and east-west present some differences, more evident
447 in the case east-west. Reason of this difference is highlighted in the adjustments sequences for
448 month of May of western and eastern series, where it is clear that four series in the western dataset
449 introduce larger (negative) adjustments. This series correspond to the 4 north-western German
450 stations, which differ for topographical characteristics with respect to the rest of the stations.

451 Therefore the extremely high variety of European topography and climate features require to
452 perform an accurate choice of the references for each target series, reasoning on its locations,
453 surroundings etc. Furthermore in some cases the sparseness of stations density doesn't allow to
454 be able to select the best stations. In further versions of the software an even angular distribution
455 of the references around the target will be implemented in case of regions with high density of
456 stations.

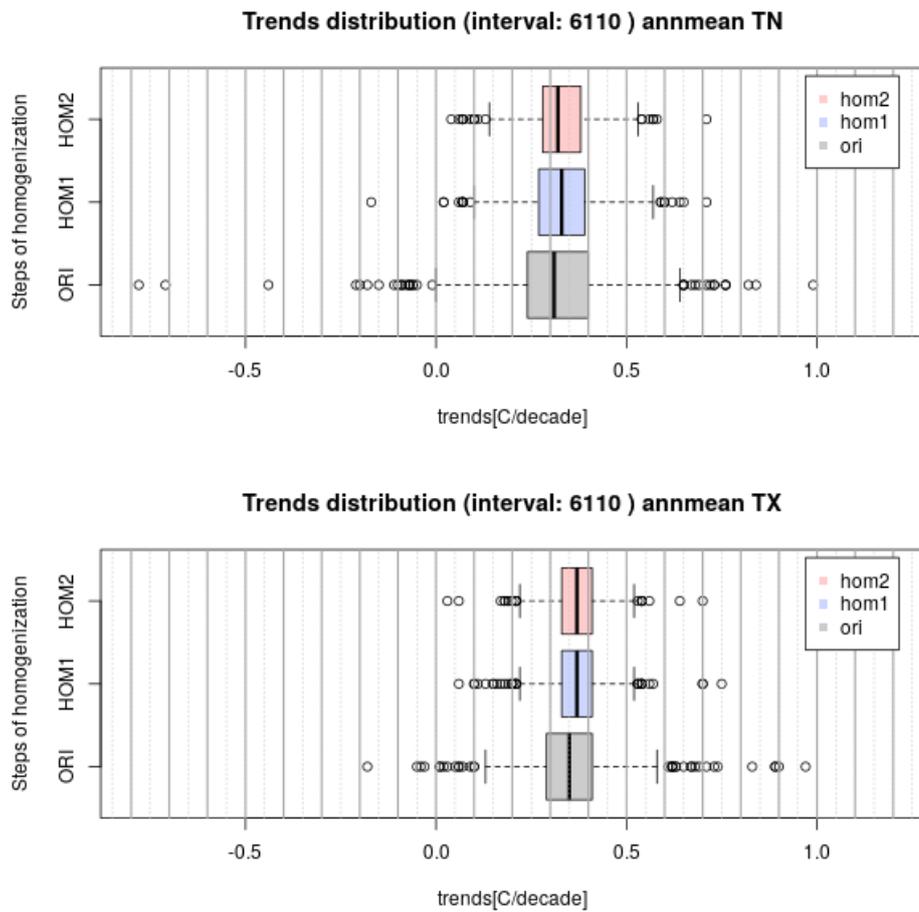


Figure 18: *Distribution of trends for annual mean of minimum (top) and maximum temperature. Each plot shows boxplot for original (grey), first iterations' result (blue) and second iterations' result (red)*

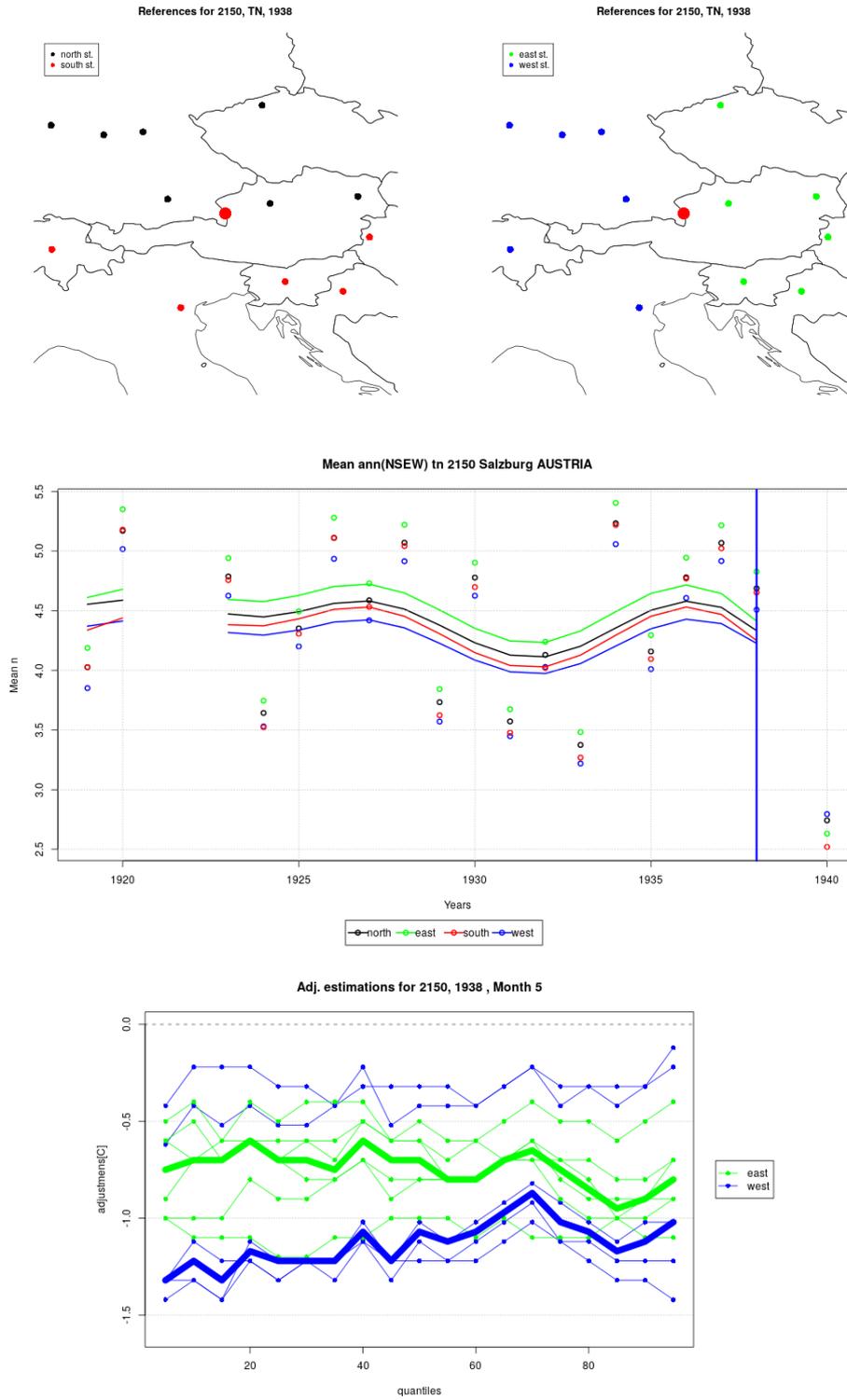


Figure 19: Annual mean of minimum temperatures in Salzburg after homogenization using 4 different sets of reference series: northern (black), eastern (blue), southern (red), western set (green).

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