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48	Abstract	<p>Few studies have investigated the perception of vestibular stimuli when they occur in sequences. Here, three experiments ($n_{\text{total}} = 33$) are presented that focus on intravestibular motion sequences and the underlying perceptual decision-making process. Natural vestibular stimulation (yaw rotation or translation) was used to investigate the discrimination process of the direction of a subsequent spatially congruent or incongruent translation or rotation. The few existing studies focusing on unimodal motion sequences have uncovered self-motion aftereffects, similar to the visual motion aftereffect, possibly due to altered processing of sensory stimuli. An alternative hypothesis predicts a shift of spatial attention due to the cue motion influencing perception of the subsequent motion stimulus. The results show that participants systematically misjudged the direction of motion stimuli well above the detection threshold if the direction of the preceding cue motion stimulus was congruent with the direction of the target (a motion aftereffect). Hierarchical drift diffusion models were used to analyze the data. The results suggest that altered perceptual decision-making and the resulting misperceptions are likely to originate in altered processing of sensory vestibular information.</p>	
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Canal–otolith interactions alter the perception of self-motion direction

Gianluca Macauda¹ · Andrew W. Ellis¹ · Luzia Grabherr¹ · Roman B. Di Francesco¹ · Fred W. Mast¹

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Abstract

Few studies have investigated the perception of vestibular stimuli when they occur in sequences. Here, three experiments ($n_{\text{total}} = 33$) are presented that focus on intravestibular motion sequences and the underlying perceptual decision-making process. Natural vestibular stimulation (yaw rotation or translation) was used to investigate the discrimination process of the direction of a subsequent spatially congruent or incongruent translation or rotation. The few existing studies focusing on unimodal motion sequences have uncovered self-motion aftereffects, similar to the visual motion aftereffect, possibly due to altered processing of sensory stimuli. An alternative hypothesis predicts a shift of spatial attention due to the cue motion influencing perception of the subsequent motion stimulus. The results show that participants systematically misjudged the direction of motion stimuli well above the detection threshold if the direction of the preceding cue motion stimulus was congruent with the direction of the target (a motion aftereffect). Hierarchical drift diffusion models were used to analyze the data. The results suggest that altered perceptual decision-making and the resulting misperceptions are likely to originate in altered processing of sensory vestibular information.

Keywords Motion aftereffect · Drift diffusion model · Vestibular cognition · Perceptual decision-making

Navigation through three-dimensional space requires keeping track of self-motion relative to an external reference frame (DeAngelis & Angelaki, 2012). Self-motion perception relies on information provided by the vestibular end organs, which are sensitive to both rotation (semicircular canals, SCCs) and translation (otolith organs). Everyday motion typically leads to combined rotatory and translatory input. Indeed, the otolith and SCC signals strongly interact in order to correctly estimate self-motion perception. The processing of vestibular information requires both inputs at the earliest level (Carriot, Jamali, Brooks, & Cullen, 2015; Cullen, 2012). For example, otolith information is ambiguous with respect to the physical motion stimulus (*tilt-translation ambiguity*), and combined otolith and SCC signals are necessary to disambiguate the sensory input (Angelaki, McHenry, Dickman, Newlands, & Hess, 1999; Merfeld, Zupan, & Gifford, 2001). Interestingly, scant attention has been paid to the perception of vestibular stimuli

and the underlying perceptual decision-making process when they occur in succession. Vestibular thresholds are commonly determined when a defined motion stimulus is presented in isolation (e.g. Grabherr, Nicoucar, Mast, & Merfeld, 2008). Two previous studies stressed the necessity to look at the influence of intravestibular interaction—that is, the interplay of otolith and SCC signals—on vestibular direction detection thresholds for better understanding real-world situations (Crane, 2016; MacNeilage, Turner, & Angelaki, 2010). So far, a small number of studies have used nulling and staircase paradigms to investigate motion sequences for vestibular or visuo-vestibular stimuli, and they have found self-motion aftereffects. These aftereffects resulted in the increased intensities of motion required to cancel out the adaptor stimulus (Crane, 2012a, 2012b; Cuturi & MacNeilage, 2014). However, in these studies, the adaptor and target stimuli always consisted of either linear translations or rotations separately.

In the visual domain, a motion aftereffect (MAE) results after prolonged exposure to coherent visual motion: A subsequently presented pattern of stationary dots appears to move in the opposite direction (the waterfall illusion; Adams, 1834). This perceptual phenomenon provides a window into the neural and computational mechanisms that underlie visual motion perception (Anstis, Verstraten, & Mather, 1998; Cuturi &

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66 MacNeilage, 2014; Konkle, Wang, Hayward, & Moore,
67 2009). The dominating explanation for this static MAE sug-
68 gests a selective adaptation mechanism of motion-sensitive
69 populations of neurons in the primary visual cortex (Huk,
70 Ress, & Heeger, 2001), which implies that MAEs occur as a
71 result of altered stimulus processing (sensitivity). According to
72 such models, prolonged exposure to a specific motion subse-
73 quently leads to a reduced firing rate and responsiveness of
74 those neurons (Huk et al., 2001). Assuming that distinct pop-
75 ulations of neurons code different directions that are constantly
76 compared, prolonged stimulation with coherent motion in one
77 direction would result in a decreased firing rate of the neurons
78 with that specific directional sensitivity. Thus, when subse-
79 quently viewing static patterns, the patterns appear to move
80 in the opposite direction (Anstis et al., 1998). More recent
81 studies started using dynamic instead of static probes. The
82 dynamic MAE differed from the static MAE, leading to the
83 idea that different neural populations are involved (see Mather,
84 Pavan, Campana, & Casco, 2008, for a review). A simplified
85 model associates the static MAE with early areas of visual
86 cortex, whereas higher-level aspects of the dynamic MAE are
87 associated with the middle temporal (MT) and medial superior
88 temporal (MST) areas. Although the dorsal part of area MST is
89 involved in processing self-motion, the role of area MT seems
90 to be limited to motion perception in general (Chowdhury,
91 Takahashi, DeAngelis, & Angelaki, 2009; Ilg, 2008).

92 However, effects of sequential motion may alternatively
93 manifest themselves in the form of a bias in perceptual deci-
94 sion-making, such that a leftward cue stimulus results in an
95 increased tendency to give a leftward response, and vice versa.
96 This is in line with the spatial-attention account (Posner,
97 Snyder, & Davidson, 1980), according to which an abstract
98 higher-level spatial (direction) cue exerts an influence on the
99 perception of the second motion stimulus. Spatial congruence
100 of the cue and target speeds up responses to the target for short
101 interstimulus intervals (ISIs). Long ISIs, however, result in an
102 impeding effect for spatially congruent trials. This phenome-
103 non is known as *inhibition of return* (IOR). Importantly, this
104 effect is not caused by low-level sensory properties of the cue,
105 but rather by abstract spatial information. In fact, the alloca-
106 tion of spatial attention can also be induced cross-modally (for
107 reviews, see Driver & Spence, 1998; McDonald, Green,
108 Störmer, & Hillyard, 2012; Spence, 2010).

109 To date, research on spatial cueing in the vestibular system
110 has been rather scarce (see Figliozzi, Guariglia, Silvetti,
111 Siegler, & Doricchi, 2005, for an exception). This is striking,
112 since vestibular stimuli are essential in the spatial processing
113 of self-motion perception. Moreover, it has been hypothesized
114 for a long time that vestibular processing is closely related to
115 spatial attention (Ferrè, Longo, Fiori, & Haggard, 2013;
116 Figliozzi et al., 2005; Shuren, Hartley, & Heilman, 1998;
117 Silberpfennig, 1941; Vallar, Sterzi, Bottini, Cappa, &
118 Rusconi, 1990).

119 Here we present three experiments with sequences of mo-
120 tions that focus on intravestibular interaction and their influ-
121 ence on the perceptual decision-making process. Yaw rotation
122 or translation preceded a spatially congruent or incongruent
123 translation or rotation, where all trials consisted of either trans-
124 lations followed by rotations, or vice versa. Participants per-
125 formed a self-motion direction discrimination task for the sec-
126 ond motion stimulus, and we also measured reaction times.

127 Taken together, the self-motion aftereffect and spatial-
128 attention accounts predict different response patterns in
129 terms of participants' choices and reaction times. On the
130 one hand, inspired by the MAE literature, a self-motion
131 aftereffect suggests a diminished ability to process a mo-
132 tion stimulus when its direction is congruent with that of
133 the preceding cue stimulus. This effect would be reflected
134 mainly in the participants' responses (correct or incorrect).
135 On the other hand, a spatial-attention account suggests ei-
136 ther facilitation in terms of faster responses (at short ISIs)
137 or IOR (at long ISIs) for congruent motion directions.
138 Facilitation could imply a bias for the motion direction that
139 is congruent with the direction of the preceding stimulus,
140 resulting in decreased reaction times (RTs), while IOR
141 would imply interference leading to longer RTs. Signal
142 detection theory distinguishes between a response bias (de-
143 cision criterion/intercept) and stimulus processing (d-
144 prime/slope). However, the classical signal detection ap-
145 proach does not take RTs into account, and it is not appro-
146 priate for dealing with uncertainty in the decision-making
147 process. (Clark, Yi, Galvan-Garza, Bermúdez Rey, &
148 Merfeld, 2018). When combined with choices, the time
149 taken to respond contains important information about
150 the cognitive process leading to a perceptual decision
151 (Shadlen & Kiani, 2013). Here we apply the drift diffusion
152 model (DDM; Ratcliff & McKoon, 2008) to better exploit
153 the information available in the data. The DDM, based on
154 the joint analysis of participants' choices and RTs, has be-
155 come a widely used cognitive model (Shadlen & Kiani,
156 2013). It allows participants' performance to be
157 decomposed into different subprocesses, represented by
158 the parameters of the DDM. According to the model, noisy
159 sensory evidence is accumulated over time until a decision
160 boundary is reached (Ratcliff & McKoon, 2008). The
161 speed of evidence accumulation is called the *drift rate*.
162 Since the DDM deals with two-alternative forced choice
163 tasks, evidence accumulated for one of the response op-
164 tions is counted as evidence against the alternative. The
165 point at which the process of evidence accumulation starts
166 is called the *starting point*. Two other relevant parameters
167 are the *boundary separation*, which defines the amount of
168 evidence necessary for a response, and the *nondecision*
169 *time*. Hence, DDMs are helpful in capturing the relevant
170 processes underlying decision-making in the perception of
171 subsequent motion stimuli.

172 **Experiment 1**

173 **Method**

174 Experiment 1 was based on a modification of the spatial-
 175 cueing paradigm by Posner, Snyder, and Davidson (1980),
 176 set up to examine potential influences of self-motion stimuli
 177 on the perception of subsequent motion stimuli. Similar to
 178 spatial cueing in the visual system, a cue stimulus (first mo-
 179 tion) was expected to direct spatial attention to one side of
 180 space, leading to faster accurate responses when the cue and
 181 target were congruent at short ISIs, and slower accurate re-
 182 sponses when the cue and target motion were congruent at
 183 long ISIs. The spatial-cueing effect is usually found for briefly
 184 presented stimuli with a short ISI (Ruz & Lupiáñez, 2002).
 185 Therefore, each of the two motion stimuli lasted 200 ms, and
 186 the ISIs ranged from 50 to 600 ms.

187 Participants were first translated horizontally (interaural y -
 188 axis, cue) and then rotated about the earth-vertical axis (yaw
 189 rotation, target). The first motion was stronger than the second
 190 motion, in order to ensure that participants would correctly
 191 perceive the cue and shift their attention accordingly. A trial
 192 was considered *congruent* when both motions had the same
 193 spatial direction (e.g., leftward interaural translation followed
 194 by leftward rotation). A trial was defined as *incongruent* when
 195 the second motion was directed in the opposite spatial direc-
 196 tion (e.g., leftward interaural translation followed by right-
 197 ward rotation). Performance in neutral trials served as a base-
 198 line; in these trials, yaw rotation was preceded by a horizontal
 199 forward or backward motion (x -axis). On the basis of the
 200 spatial-cueing paradigm, neutral trials were not expected to
 201 influence the subsequent motion. For both leftward/
 202 rightward and forward/backward translations, the first motion
 203 was 3.6 times above the threshold established in a preliminary
 204 study and similar to that in published data (Valko, Lewis,
 205 Priesol, & Merfeld, 2012), and the second motion was 2.6
 206 times above the threshold reported by Grabherr, Nicoucar,
 207 Mast, and Merfeld (2008).

208 **Participants** Twelve healthy participants participated in the
 209 first experiment (eight male, four female; mean age 27, range
 210 24–30 years). All participants were right-handed according to
 211 a German version of the handedness questionnaire by
 212 Chapman and Chapman (1987). None of the participants re-
 213 ported a history of neurological, vestibular, or attentional dis-
 214 orders. The study was approved by the ethics committee of the
 215 University of Bern, and all participants gave written informed
 216 consent prior to the experiment in accordance with the
 217 Declaration of Helsinki.

218 **Motion stimuli** A six-degree-of-freedom motion platform
 219 (6DOF2000E; Moog Inc., East Aurora, NY) and in-house
 220 software were used to generate the motion stimuli. The cue

stimuli consisted of translations with single-cycle sinusoidal
 acceleration and a frequency f of 5 Hz [$a(t) = A \sin(2\pi ft)$, $T =$
 $1/f$] along both the y -axis (left/right) and the x -axis (forward/
 backward), as had been used in previous studies (see, e.g.,
 Crane, 2012a; Grabherr et al., 2008). The acceleration ampli-
 tude (A) was set to 0.25m/s^2 , resulting in a peak velocity of
 0.016m/s ($v_{\text{max}} = AT/\pi$) and a displacement of 0.0016 m
 $(\Delta p = AT^2/2\pi)$. For the target stimuli, we used the same accel-
 eration profile as for the cue stimuli, but with yaw rotations
 about an earth-vertical axis (left/right). The acceleration am-
 plitude was set to 24 deg/s^2 ($v_{\text{max}} = 1.53\text{ deg/s}$, $\Delta p =$
 0.153 deg).

Experimental design Cue congruence (three levels: congruent,
 incongruent, and neutral) and ISI (four levels: 50, 100, 200,
 and 600 ms) were varied within each participant. In *congruent*
 trials, interaural y -axis translations were followed by yaw ro-
 tations to the corresponding side. In *incongruent* trials, y -axis
 translations preceded yaw rotations in the opposite direction.
 Neutral trials consisted of a translation along the naso-
 occipital x -axis and a subsequent yaw rotation. In-house soft-
 ware based on LabVIEW (National Instruments, Austin,
 Texas) was used to record participants' response and RTs.

Experimental procedure Participants were seated in a car seat
 with a five-point harness, which was mounted on the motion
 platform. Their head was fixated by means of a helmet. The
 experiments were conducted in darkness, and participants
 were blindfolded to prevent the perception of surrounding
 visual cues. White noise was delivered to in-ear headphones
 at approximately 60 dB to mask sounds from the engine of the
 motion platform. Participants were instructed to indicate the
 perceived direction of the second motion as fast as possible by
 pressing the corresponding button with either their left or right
 hand. The first motion served as a time-varying warning cue,
 which indicated the beginning of the second motion but did
 not predict its direction. Practice trials were administered until
 participants understood the task. Once a participant was com-
 fortable with the task, the motion sets were presented in three
 blocks. Participants were given sufficient time to rest between
 blocks. In total, 192 trials were presented randomly (48 con-
 gruent, 48 incongruent, and 96 neutral). A trial consisted of a
 cue stimulus (200 ms), a variable ISI (50–600 ms), a target
 stimulus (200 ms), time to respond (max. 2,500 ms), time for
 return to the origin (1,100 ms), and an intertrial interval of
 about 1,000 ms.

Data analysis All analyses were computed in R (R Core Team,
 2013) using the brms (Bürkner, 2017) and rstan (Guo et al.,
 2017) packages, which implement Bayesian inference proce-
 dures. These procedures provide posterior probability distri-
 butions for the estimated parameters. For all calculated statis-
 tical models, samples of each parameters' posterior

271 distribution were drawn with a Hamiltonian Monte Carlo sampling
 272 algorithm implemented in Stan. Samples were generated
 273 by four independent Markov chains, each with 1,000 warm-up
 274 samples, followed by another 1,000 samples drawn from the
 275 posterior distribution. Those 1,000 samples for each Markov
 276 chain were retained for further statistical inference. To confirm
 277 that the samples for each chain converged to the same poste-
 278 rior distribution, the R-hat statistic was used, along with visual
 279 inspection of the shape of the posterior distribution and the
 280 chains (Gelman et al., 2014). For all calculated models the R-
 281 hat statistics were below 1.05. Together with the visually
 282 inspected chains, this indicated that all Markov chains con-
 283 verged to the same posterior distribution of the estimated pa-
 284 rameters. These posterior distributions can be interpreted as
 285 Bayesian credible intervals. Calculations were performed on
 286 UBELIX (<http://www.id.unibe.ch/hpc>), the HPC cluster at the
 287 University of Bern. All analyses, models, and data are freely
 288 accessible on the Open Science Framework (OSF; <https://osf.io/46nqw/>).

290 To see whether participants were able to correctly perceive
 291 the direction of the yaw rotation, their performance was ana-
 292 lyzed with a multilevel Bayesian logistic regression model
 293 implemented in brms, including the factors cue congruence
 294 (congruent, incongruent, neutral) and ISI (50, 100, 200, 600
 295 ms). To analyze RTs in more detail, multilevel Bayesian mul-
 296 tiple regression models, with a lognormal likelihood function
 297 including the factors correctness (correct, wrong), cue congru-
 298 ence (congruent, incongruent, neutral), and ISI (50, 100, 200,
 299 600 ms), were calculated for Experiments 1 and 2.

300 To jointly analyze participants' responses and RTs, and to
 301 study the bias introduced by the first motion, we modeled the
 302 data using two different Bayesian multilevel DDMs
 303 (Vandekerckhove, Tuerlinckx, & Lee, 2011) estimated in
 304 brms (Bürkner, 2017). This procedure allows for parameter
 305 estimates for groups (fixed effects) while considering varia-
 306 tion between participants. For the first experiment, the neutral
 307 conditions were not included in the statistical models, since
 308 they were uninformative with respect to the underlying pro-
 309 cesses of congruent and incongruent trials. Moreover, the con-
 310 gruent and incongruent conditions were split into their specific
 311 directions in order to obtain a separate factor for the cue and
 312 target motion. This allowed us to fit a model in which the cue
 313 motion influenced the starting point for the direction of the
 314 target motion.

315 To investigate the decision-making process of self-motion
 316 perception in more detail, two different DDMs with either a
 317 flexible drift rate or a flexible starting point were calculated for
 318 Experiments 1 and 2. For the flexible drift rate, we estimated
 319 the fixed effects of the cue motion, the target motion, and the
 320 ISI, together with by-participant random effects. For the
 321 starting point, we estimated random intercepts for participants,
 322 but included no covariates. In the model with the flexible
 323 starting point for the drift rate, we estimated the fixed effects

of the cue motion and the ISI, together with by-participant
 random effects. Moreover, for the starting point, fixed effects
 were estimated of the cue motion and the ISI, along with by-
 participant random effects. Models for Experiments 1 and 2
 were compared using the leave-one-out cross-validation
 (LOO; Vehtari, Gelman, & Gabry, 2017). The LOO is a mea-
 sure of a model's predictive accuracy. In the following com-
 parisons, we report the model-specific LOO information cri-
 terion (LOOIC) and the LOOIC difference between the two
 models, as well as their standard errors of the LOOIC (see also
 Wallis et al., 2017, for a similar procedure). The LOOIC esti-
 mates the expected log pointwise predictive density.
 Multiplication by - 2 converts the measure to the deviance
 scale. Hence, lower LOOIC values indicate a better model fit.
 For all DDM models, we used weakly informative prior dis-
 tributions for the fixed-effect parameters (see the
[supplementary online material](#)). All other priors were set to
 the default implemented in brms and can be checked in the
 analysis file available on the OSF.

To assess whether the DDMs calculated for Experiments 1
 and 2 provided a good description, we generated 500 com-
 plete datasets (responses and RTs) from each model's poster-
 ior predictive distribution. In a first step, within each dataset,
 we calculated the mean response probabilities and median RTs
 for the upper and lower responses for each participant in each
 experimental condition. In a second step, the 500 datasets
 were summarized by different quantiles (.025, .1, .5, .9, and
 .975) for both the predicted response probabilities and the
 median RTs for lower and upper responses, for each partici-
 pant in each experimental condition. The predicted medians of
 the response probabilities and RTs were then compared to the
 observed responses and RTs. To quantify the distance between
 the models' predictions and the observed data, a distance mea-
 sure was calculated. This measure for the specific models was
 calculated as the sum of the squared differences between ob-
 servations and the median predictions for each participant in
 each experimental condition. The distance measures for the
 different models are presented in Table 1. Lower values cor-
 respond to a closer match between the generated and the ob-
 served datasets.

$$\text{Distance} = \sum_{i=1}^n (\text{observation} - \text{prediction})^2$$

Table 1 Distance measures for the different models in Experiments 1 and 2

Model	Experiment 1		Experiment 2		
	No Bias	Bias	No Bias	Bias	
Response probability "Right"	1.91	4.31	1.60	6.26	t1.4
Median RT "Right"	33.45	36.12	22.14	26.46	t1.5
Median RT "Left"	59.30	60.56	37.80	45.29	t1.6

366 The medians and quantiles of the response probabilities
367 and RTs were then averaged across all participants. A compar-
368 ison of the generated medians to the observed values over
369 all participants is presented in Supplementary Tables 1–4.
370 Moreover, they are visually compared to the observed data
371 in Figs. 3, 4, and 7 below. The described model checks are
372 based on standard procedures to assess the model fit
373 (Singmann, 2017, 2018).

374 Results

375 **Accuracy** The logistic regression showed that motion discrim-
376 ination performance was below chance level when the cue and
377 target motions were congruent. This suggests that participants
378 misperceived the direction of the target motion. Direction discrim-
379 ination performance was above chance level when a yaw
380 rotation was preceded by an incongruent or a neutral forward–
381 backward motion (Fig. 1, Exp. 1), suggesting that participants
382 correctly perceived the direction of the yaw rotation. The dis-
383 crimination performance did not differ between the incongru-
384 ent and neutral motions, since all 95% credible intervals (CIs)
385 for the posterior distributions of the difference included zero.
386 Overall, the response patterns indicate that the direction of the
387 translation leads to a misperception of the direction of the
388 rotation when the directions were congruent—that is, a self-
389 motion aftereffect—except at an ISI of 600 ms. When taking
390 into account the RTs presented in Fig. 2, interpretation of the
391 combined measures (accuracy and RTs) remained difficult and
392 inconclusive. Thus, DDMs were calculated in order to provide
393 more insights into the underlying process of the self-motion
394 aftereffect.

395 **Reaction times** The analysis of the RTs indicated no difference
396 between correct (estimate of the mean = 1.52 s, 95% CI =
397 [1.36, 1.71]) and wrong (estimate = 1.51 s, 95% CI = [1.40,
398 1.62]) responses when the cue and target motions were con-
399 gruent at an ISI of 50 ms. For the same ISI, participants were
400 slightly faster for correct (estimate = 1.50 s, 95% CI = [1.40,
401 1.61]) than for wrong (estimate = 1.73 s, 95% CI = [1.54,
402 1.95]) responses when the cue and target stimuli were incon-
403 gruent. This difference was even more pronounced in the neu-
404 tral condition (correct responses: estimate = 1.31 s, 95% CI =
405 [1.25, 1.38]; wrong responses: estimate = 1.66 s, 95% CI =
406 [1.54, 1.80]). RTs in wrong trials for an ISI of 50 ms did not
407 differ between the congruent, incongruent, and neutral condi-
408 tions. The slowest RTs for wrong responses were observed at
409 an ISI of 600 ms over all conditions (estimate = 2.41 s, 95% CI
410 = [2.16, 2.69]). Parameter estimates of the means and 95% CIs
411 for all conditions are depicted in Fig. 2.

412 **Drift diffusion models** The LOOICs for the model with the
413 perceived direction of the response and the drift-rate param-
414 eter depending of the direction of the first motion (no-bias

model), the model with the starting point also depending on
the direction of the first motion (bias model), and their LOOIC
difference are presented in Table 2 (Exp. 1 columns). The
LOOIC values indicate that the fit for the model with the
drift-rate parameter depending on the direction of the first
motion (no bias) was better, which is reflected in its lower
LOOIC.

A comparison of the posterior predictive distributions
and the observed data for both models in Experiment 1 is
presented in Fig. 3; in the left panels are comparison results
for the model that included cue and target motion as well as
ISI as predictors of the drift rate (no-bias model), and in the
right panels are results for the model that included target
motion and ISI as predictors for the drift rate, with the
starting point depending on the cue motion in combination
with ISI (bias model). The first row in Fig. 3 shows compar-
isons of the predicted (circles) and observed (crosses)
response probabilities of responding “right” for the no-bias
and bias models across all participants. This comparison
shows that the predicted and observed response probabili-
ties of responding “right” are closer in the no-bias than in
the bias model. Moreover, the 80% (fat gray bars) and 95%
(thin gray lines) credible intervals constructed from the
quantiles averaged across participants are narrower in the
no-bias model. These results indicate that the model with
no bias describes the observed response probabilities better
than the bias model. Moreover, the ranges of the predic-
tions for the response probabilities made by the no-bias
model are narrower. This reflects decreased uncertainty
about the model’s posterior parameter estimates and means
that the model is able to make more precise predictions.

The second and third rows in Fig. 3 show the predicted
(black circles and gray lines and bars) and observed (color
crosses) median RTs for “left” and “right” responses. The
intervals of the predicted median RTs are equally wide for
both models. The observed and predicted median RTs over
all conditions are also equally distant. Overall, the model for
which the drift rate depends on the cue and target motions as
well as the ISI, but with no predictor for starting point (no-bias
model), describes the observed data better, especially the re-
sponse probabilities. This conclusion is also supported by that
model’s lower LOOIC. Therefore, the model with no shifted
starting point was preferred.

To have more trials for each condition, and therefore more
precise parameter estimates for our models, we decided not to
distinguish between the explicit directions of the cue and tar-
get motions and to summarize the directions in the factor cue
congruence. The response was then coded in terms of accura-
cy. Thus, we estimated an additional model with fixed effects
of congruence and ISI, together with random effects of partic-
ipants for the drift rate (no-bias model, but with correctness as
the response variable). For the starting point, we estimated a
by-participant random intercept model.

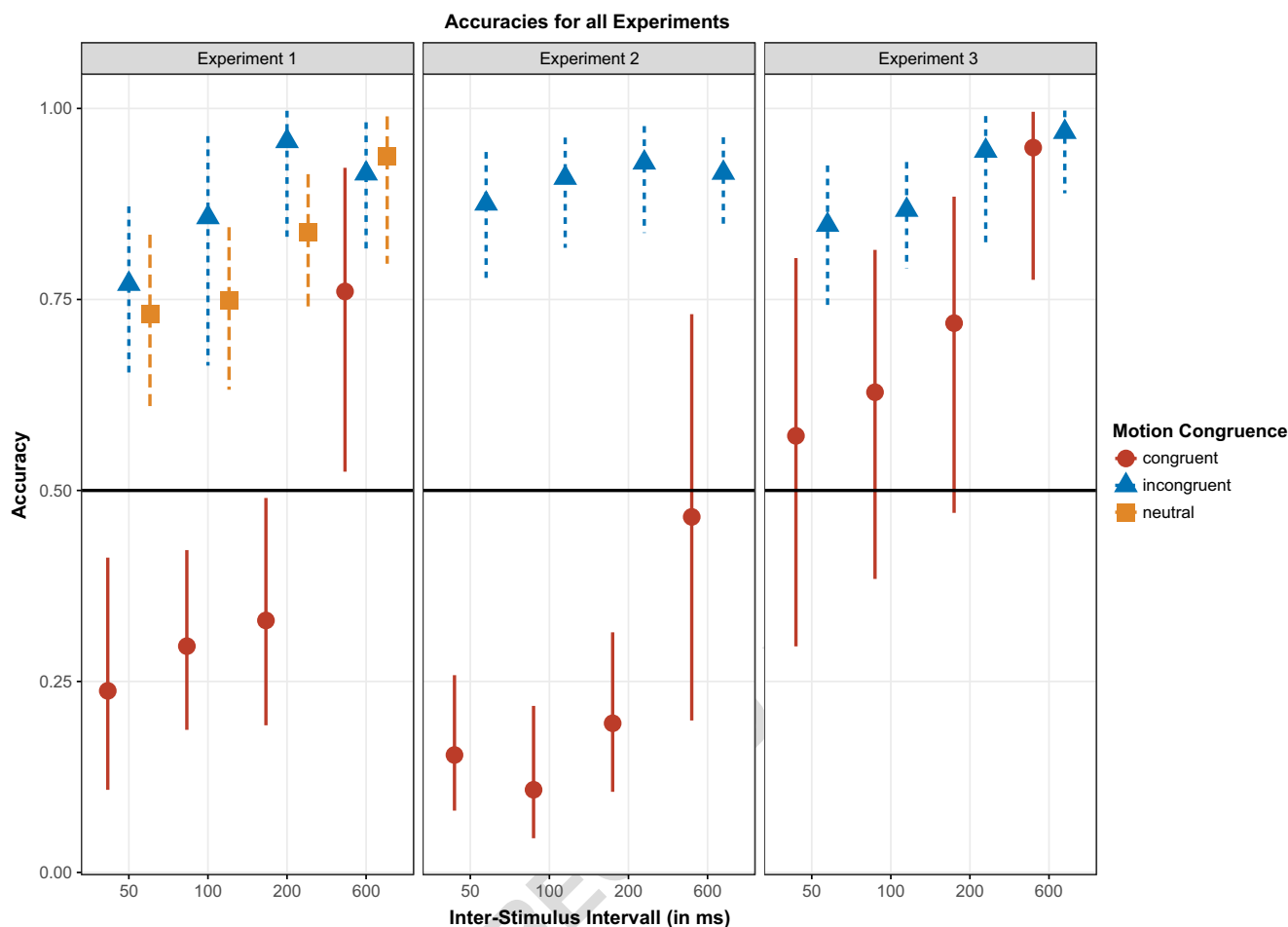


Fig. 1 Participants' performance in all three experiments. Dots, triangles, and squares represent the inverse logit-transformed parameter estimates of the different conditions in the logistic regressions. The lower and upper

ends of the lines surrounding them represent the 95% credible intervals. The black lines at .5 illustrate chance level

468 The goodness of fit for this model, based on posterior
 469 predictive checks, is presented in Fig. 4, left panels.
 470 The model predicts participants' responses well, since
 471 the observed median response probabilities (crosses)
 472 are close to the predicted response probabilities
 473 (circles) and are within both the 80% and 95% intervals
 474 of the predictive posterior distribution. Median RTs for
 475 the correct responses match the observed median RTs,
 476 as well. However, the model has difficulties describing
 477 the RTs for the wrong responses, especially in the in-
 478 congruent conditions, since the observed median RTs
 479 are at the upper end of the 80% and 95% intervals of
 480 the posterior predictive distributions.

481 The DDM with correctness as a response variable re-
 482 vealed that evidence for wrong responses was accumulat-
 483 ed when the cue and target motions were spatially con-
 484 gruent for ISIs of 50, 100, and 200 ms. The parameter
 485 estimates for both incongruent cues and congruent cues
 486 at an ISI of 600 ms indicate sensory evidence accumula-
 487 tion for the correct response (Fig. 5).

Discussion

488
 489 In Experiment 1, the results of logistic regression showed that
 490 there are compelling intravestibular self-motion aftereffects in
 491 terms of direction discrimination performance when a yaw
 492 rotation is preceded by a translation in the same spatial direc-
 493 tion. Thus, stimulation of the otoliths influences a sensory
 494 decision-making process that is based on information provid-
 495 ed by the semicircular canals, so that congruent motions are
 496 systematically misperceived. We concluded that the perceptual
 497 decision about the direction of the target motion uses an
 498 estimate of spatial direction obtained not only from the semi-
 499 circular canals, but also from the preceding stimulation of the
 500 otoliths. Importantly, self-motion aftereffects were found for a
 501 specific time window of ISIs of 50–200 ms, but not at 600 ms.
 502 This time window is much shorter than that reported in the
 503 vestibular MAE literature (Crane, 2012a). A graphical sum-
 504 mary of these misperceptions for congruent trials and the un-
 505 derlying decision-making processes is illustrated in Fig. 6.
 506 The RT data in Experiment 1 show that participants generally

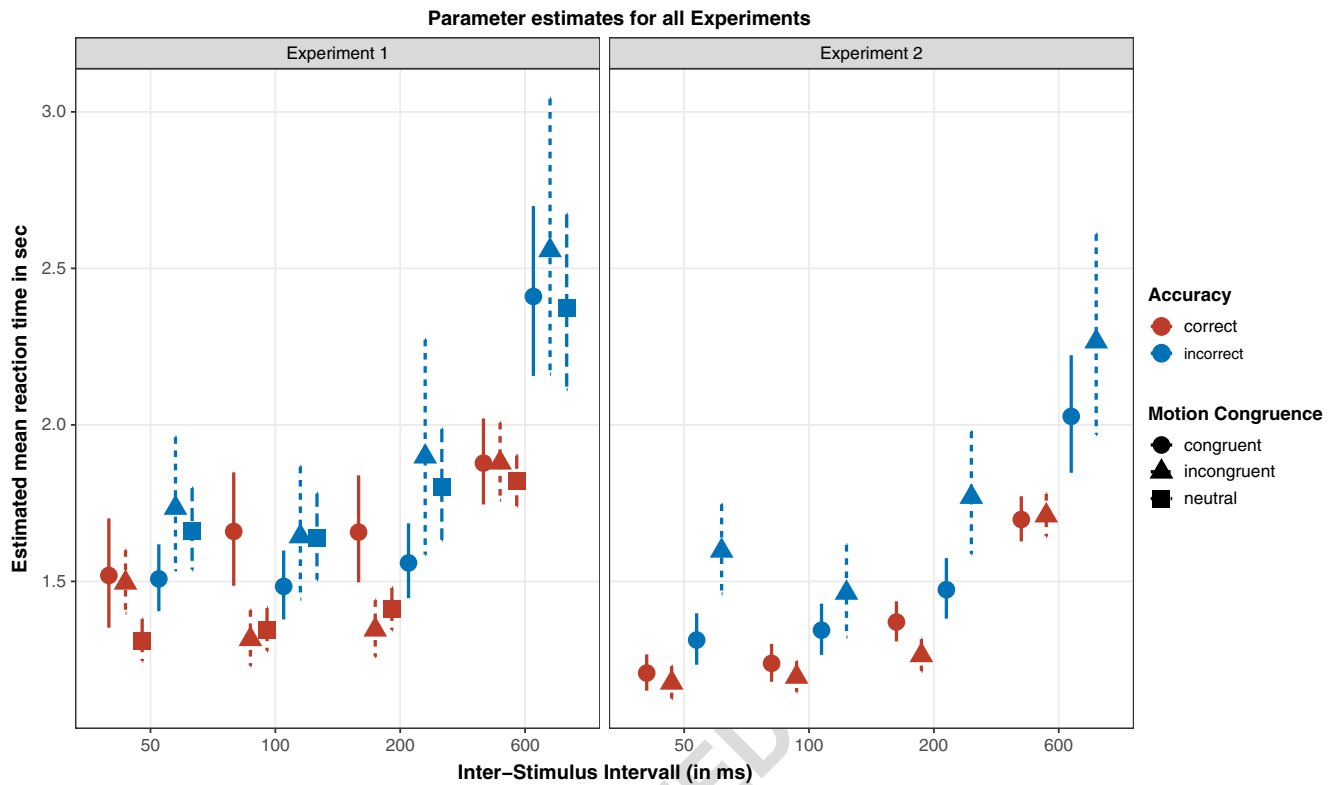


Fig. 2 Participants' RT data in Experiments 1 and 2. Dots, triangles, and squares represent the transformed parameter estimates of the means from of the different conditions in the lognormal regressions. The lower and upper ends of the lines surrounding them represent the 95% credible intervals

507 responded faster for correct responses. This was not the case in the congruent conditions for ISIs of 50, 100, and 200 ms. In these conditions, the RTs for wrong responses were even slightly faster. For these ISIs, self-motion aftereffects were observed. This might suggest that participants were convinced to respond correctly but did in fact accumulate evidence for the wrong direction.

514 **Experiment 2**

515 **Method**

516 The results of Experiment 1 raise the question of whether the consistent directional misperception of the second motion (self-motion aftereffect) would also occur if the rotation and

translation were reversed. Specifically, we investigated the influence of a rotation (cue) on a subsequent translation (target). The motion trajectories remained unchanged. Only congruent and incongruent cues were presented, since no rotation could serve as a “neutral” cue for lateral translations in the horizontal plane. Both pitch and roll rotations in an upright participant would inevitably lead to otolith stimulation, due to a resulting deviation from the direction of gravity.

Participants Twelve new participants were recruited for Experiment 2. Two of the participants had to abort the experiment, and we included the remaining ten participants in the analysis (six male, four female; mean age 29, range 23–56 years). All participants were right-handed according to a German version of the handedness questionnaire by Chapman and Chapman (1987). None of the participants

t2.1 **Table 2** LOOIC values for the models in Experiments 1 and 2

t2.2	Experiment 1			Experiment 2		
	No Bias	Bias	No Bias – Bias	No Bias	Bias	No Bias – Bias
t2.4	2,794.30	3,114.74	– 321.45	4,348.22	5,415.79	– 1,067.57
t2.5	80.42	75.25	46.99	132.25	71.64	87.48

LOOIC = leave-one-out information criterion, SE = standard error

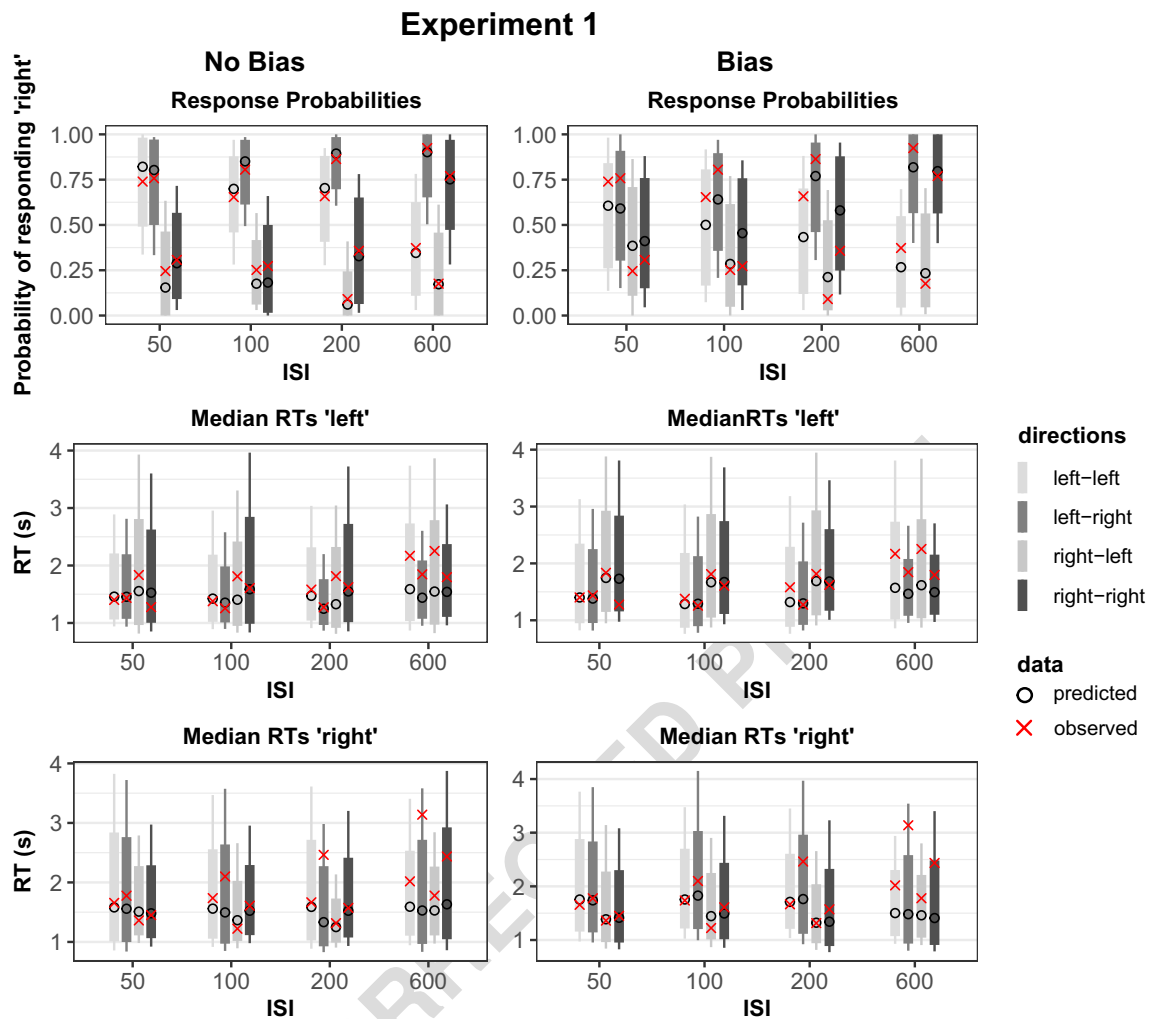


Fig. 3 Comparison of the posterior predictive distributions and the observed data from Experiment 1, for the model with the drift rate depending on cue and target motion but no shift in the starting point (on the left; no-bias model) and for the model with the drift rate depending on target motion and ISI and with the starting point depending on the cue motion in combination with ISI (on the right; bias model). The median predictions for response probabilities and median RTs averaged across all participants are plotted as black circles. The

bold gray bars represent 80% credible intervals, computed from the .1 and .9 quantiles. The thinner gray bars represent 95% intervals, computed from the .025 and .975 quantiles. Together, these indicate the uncertainty of the models' predictions, with larger intervals indicating more uncertainty. The means of the observed response probabilities and the medians of the observed RT data, averaged across all participants, are shown as crosses. The different ISIs are presented on the x-axis

Q4

534 reported a history of relevant neurological, vestibular, or at-
 535 tentional disorders. The study was approved by the ethics
 536 committee of the University of Bern, and all participants gave
 537 written informed consent prior to the experiment in accord-
 538 ance with the Declaration of Helsinki.

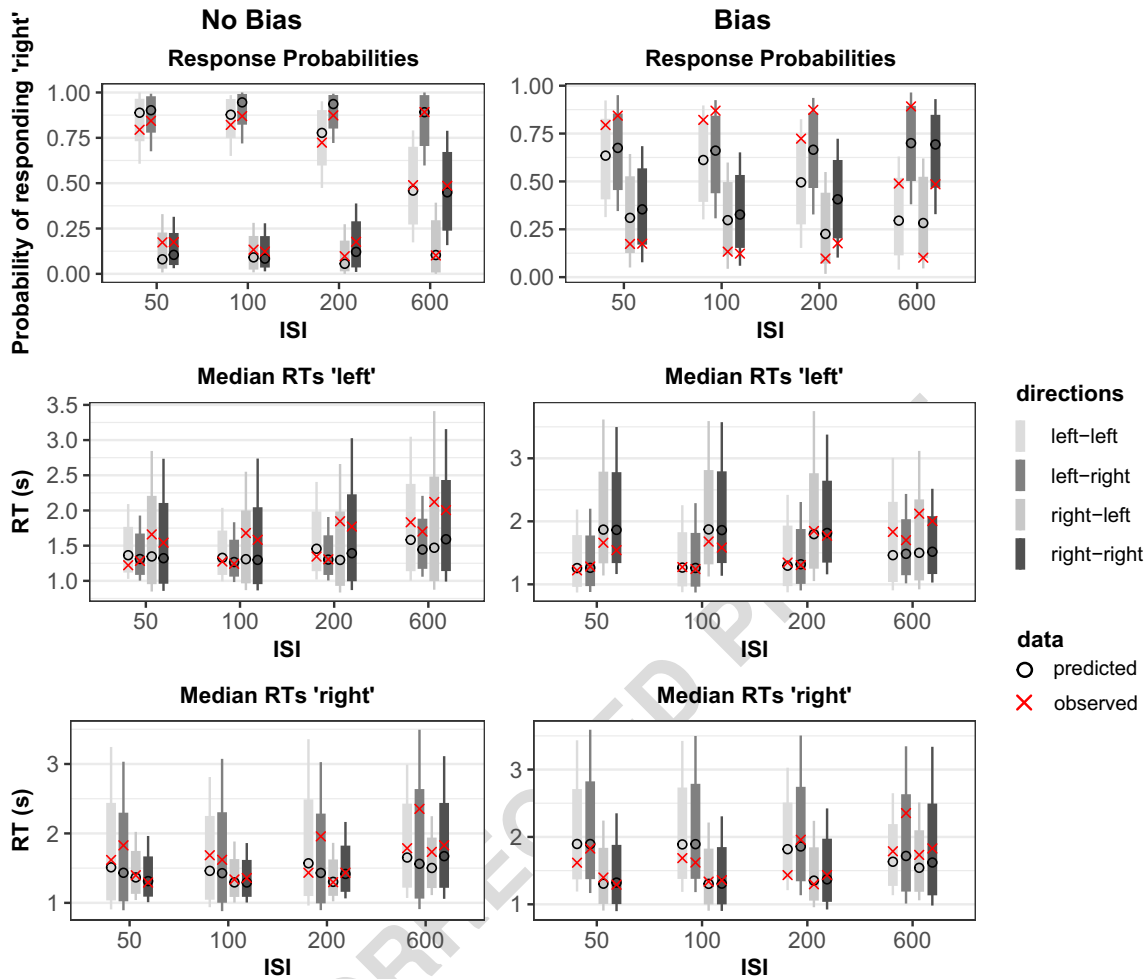
539 **Motion stimuli** The cue stimuli were rotations about an
 540 earth-vertical axis (right/left) with single-cycle sinusoi-
 541 dal acceleration and a frequency of 5 Hz [$a(t) = A$
 542 $\sin(2\pi ft) = A \sin(2\pi t/T)$]. The acceleration amplitude was
 543 set to 33 deg/s^2 ($v_{\max} = 2.10 \text{ deg/s}$, $\Delta p = 0.021 \text{ deg}$).
 544 The target stimuli consisted of translations in the y-axis
 545 (right/left) with the same acceleration profile as the cue
 546 stimuli. The acceleration amplitude was set to 0.18 m/s^2

($v_{\max} = 0.011 \text{ m/s}$, $\Delta p = 0.0011 \text{ m}$). The linear accelera- 547
 548 tion was always along the participants' y-axis.

Experimental design Within each participant, cue congruence 549
 550 (two levels: congruent, incongruent) and ISI (four levels: 50, 551
 552 100, 200, and 600 ms) were manipulated. In congruent trials a 552
 553 yaw rotation was followed by a translation along the interaural 553
 554 y-axis to the corresponding side. In incongruent trials a yaw 554
 555 rotation was followed by an interaural y-axis translation to the 555
 556 opposite side. Again, participants' responses and RTs were 556
 557 recorded.

Experimental procedure The experimental procedure was 557
 558 similar to that of Experiment 1, differing only in the type of 558

Experiment 2



Q5 **Fig. 4** Comparison of the posterior predictive distributions and the observed data for Experiments 1 (left panels) and 2 (right panels), with the drift depending on the congruence of the directions of cue and target motion and the ISI, no shift in the starting point, and correctness as the response variable. The median predictions for response probabilities and median RTs averaged across all participants are plotted as black circles. The bold gray bars represent 80% credible intervals, computed from the

.1 and .9 quantiles. The thinner gray bars represent 95% intervals, computed from the .025 and .975 quantiles. Together, these indicate the uncertainty of the models' predictions, with larger intervals indicating more uncertainty. The means of the observed response probabilities and the medians of the observed RT data averaged across all participants are shown as crosses. The different ISIs are presented on the x-axis

559 motion sequences presented. In this experiment, 192 trials
560 were randomly presented (96 congruent, 96 incongruent).

561 **Data analysis** Participants' response accuracy was analyzed
562 with a multilevel Bayesian logistic regression similar to that
563 in Experiment 1, but with the factor cue congruence consisting
564 of only two levels (congruent, incongruent). The details of the
565 joint analysis of accuracy and RTs (DDM) can be found in the
566 Method section of Experiment 1.

567 **Results**

568 **Accuracy** The logistic regression model (Fig. 1, Exp. 2 panel)
569 revealed that the motion discrimination performance was below
570 chance level in the congruent conditions for ISIs of 50,

100, and 200 ms, indicating a misperception of the spatial
571 direction of the translation. For an ISI of 600 ms, the discrimi-
572 nation performance was around chance level. Direction discrimi-
573 nation performance was above chance level when a yaw
574 rotation preceded an incongruent motion, meaning that the
575 spatial direction of the target stimulus was correctly perceived.
576 Just as in Experiment 1, the response patterns indicate that the
577 direction of the rotation led to a misperception of the direction
578 of the translation when the directions were congruent—that is,
579 a self-motion aftereffect—except for an ISI of 600 ms. 580

581 **Reaction times** The analysis of RTs indicated that participants
582 were slightly faster on correct (estimate = 1.21 s, 95% CI =
583 [1.15, 1.27]) than on wrong (estimate = 1.31 s, 95% CI =
584 [1.24, 1.40]) responses when the cue and target motions were

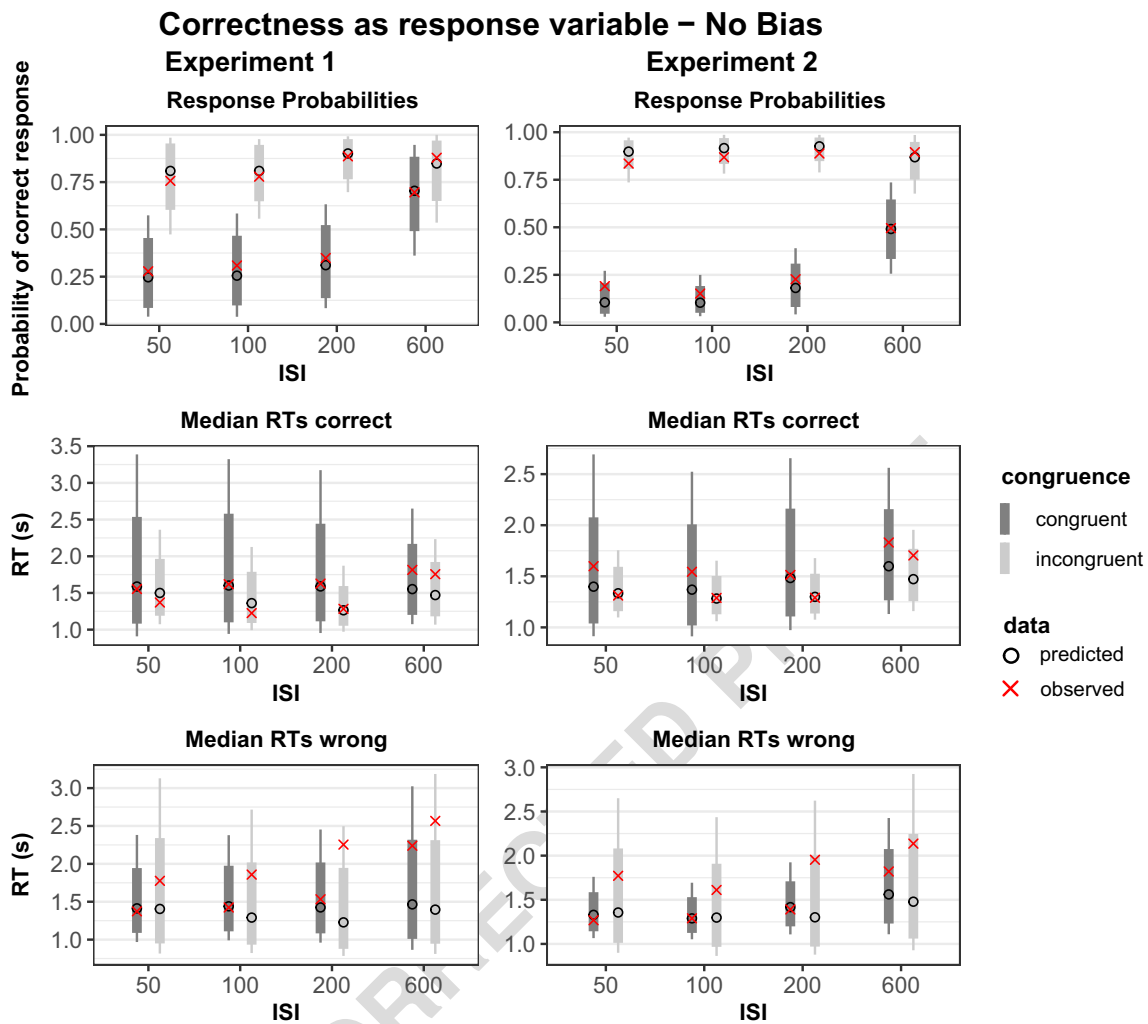


Fig. 5 Estimates for the effects of congruence and ISI on the drift rate of the drift diffusion models in Experiments 1 and 2. The estimates for Experiment 1 are represented by squares, and the estimates for Experiment 2 by triangles. Estimates for congruent conditions are in a

cool color, whereas estimates for incongruent conditions are in a warm color. Vertical lines around the estimates show the 95% credible interval for each estimate

Q6

585 congruent at an ISI of 50 ms. For the same ISI, participants
 586 were meaningfully faster for correct (estimate = 1.18 s, 95%
 587 CI = [1.13, 1.23]) than for wrong (estimate = 1.60 s, 95%
 588 CI = [1.47, 1.74]) responses when the cue and target stimuli were
 589 incongruent. Additionally, RTs for wrong responses were
 590 clearly faster in the congruent than in the incongruent condi-
 591 tion. This difference was absent for the RTs of correct re-
 592 sponses. The slowest RTs for wrong responses were observed
 593 at an ISI of 600 ms over all conditions (estimate = 2.03 s, 95%
 594 CI = [1.85, 2.23]). Parameter estimates of the means and the
 595 95% CIs for all conditions are depicted in Fig. 2.

596 **Drift diffusion model** The LOOIC values for the model with
 597 the perceived direction as response and the drift-rate param-
 598 eter depending of the direction of the first motion (no-bias
 599 model), the model with the starting point depending on the
 600 direction of the first motion (bias model), and their LOOIC

601 difference are presented in Table 2 (Exp. 2). As in Experiment
 602 1, the lower LOOIC value indicates that the model fit for the
 603 model with the drift-rate parameter depending on the direction
 604 of the first motion (no-bias model) was better.

605 A comparison of the posterior predictive distributions and
 606 the observed data for both models in Experiment 2 is present-
 607 ed in Fig. 7; in the left panels are comparison results for the
 608 model that included cue and target motion as well as ISI as
 609 predictors of the drift rate (no-bias model), and the right panels
 610 are results for the model that included target motion and ISI as
 611 predictors for the drift rate, with the starting point depend-
 612 ing on the cue motion in combination with ISI (bias model). As in
 613 Fig. 3, the first row in Fig. 7 shows comparisons of the pre-
 614 dicted and observed median response probabilities of
 615 responding “right” for the no-bias and bias models across all
 616 participants. This comparison shows that the predicted and
 617 observed response probabilities of responding “right” are

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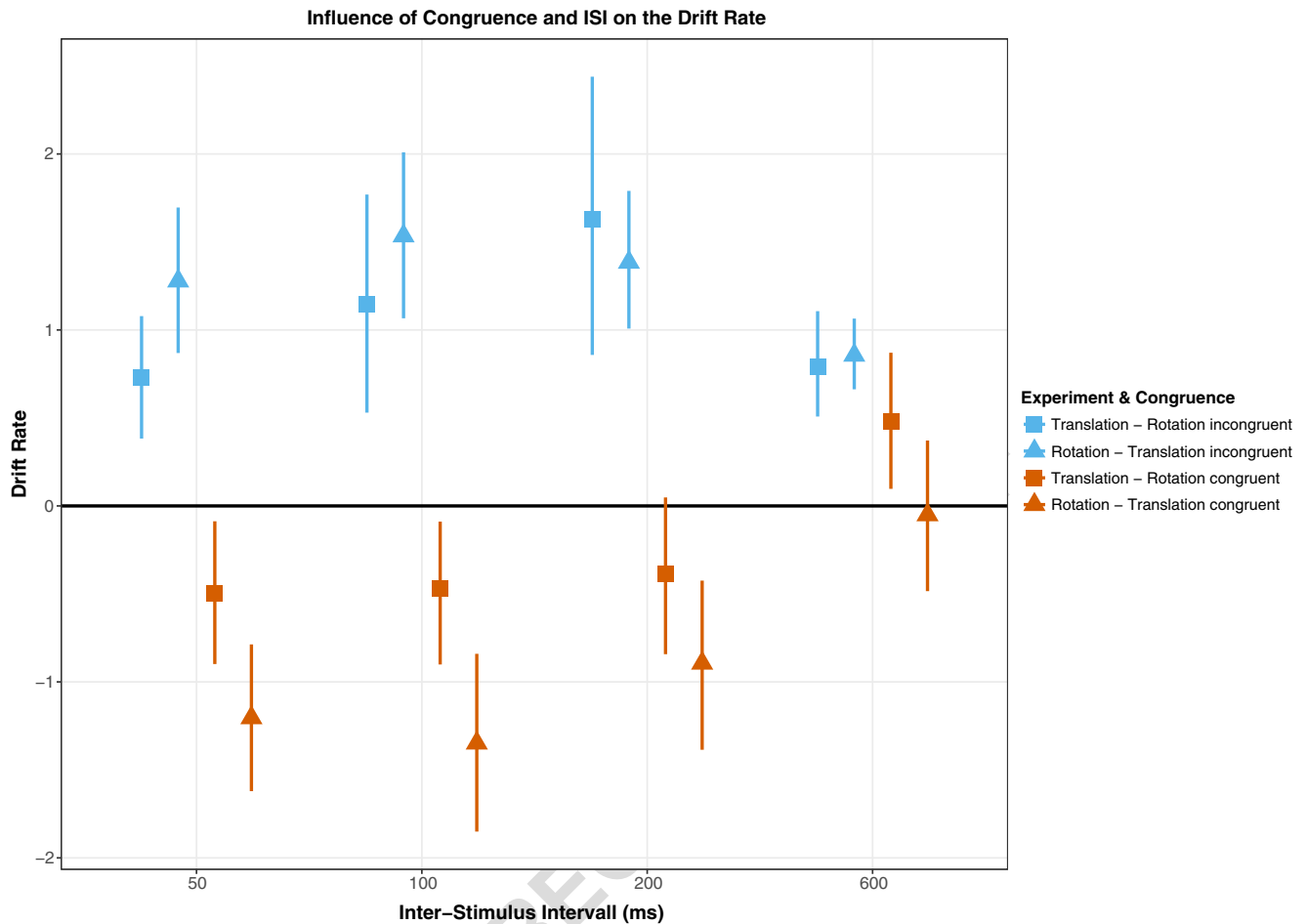


Fig. 6 An illustration of the self-motion aftereffects and the underlying perceptual decision-making processes in Experiment 1. In this example, left/right translations are used as the cues and yaw rotations as the targets. The reverse pairing yielded the same results (Exp. 2). Thought bubbles indicate the perceived self-motion direction. Checkmarks represent correct perceptions of self-motion direction, and Xs represent misperceptions. The accumulation of sensory evidence is depicted in the right graph for each cue–target combination. A decision is made when the evidence accumulation process reaches a threshold (left or right). The sensory evidence accumulation always starts at the same

starting point, equidistant from both decision thresholds—that is, there is no response bias. In contrast, the sign of the drift rate is switched, depending on the congruency of cue and target motions. In the upper half, the cue and target motions are incongruent. For those conditions, evidence for the actual physical stimulus is accumulated, leading to a correct perceptual decision. In the lower half, the cue and target motions are spatially congruent. In these conditions, sensory evidence is accumulated for the opposite target direction, leading to an incorrect decision

618 much closer in the no-bias than in the bias model. Moreover, 619 the 80% (fat gray bars) and 95% (thin gray lines) credible 620 intervals reflecting the posterior predictive distributions are 621 quite narrow in the no-bias model. These observations indicate 622 that the model with no bias describes the observed response 623 probabilities substantially better than the bias model. 624 Moreover, the ranges of the predictions for the response 625 probabilities made by the no-bias model are narrower, indicating 626 the precision of the predictions.

627 Again, the second and third rows in Fig. 7 show the predicted 628 and observed median RTs for “left” and “right” responses. 629 The intervals of the predicted median RTs are equally 630 wide for both models. The distance of the observed and predicted 631 median RTs over all conditions is slightly closer in the

bias model (right panels). However, the no-bias model is also 632 able to describe the observed median RTs. 633

634 Overall, the model for which the drift rate depends on the 635 cue and target motions as well as the ISI (no-bias model) 636 describes the observed data better, especially the response 637 probabilities. Again, this conclusion is also supported by that 638 model’s lower LOOIC. Therefore, the model with no shifted 639 starting point was preferred.

640 As in Experiment 1, we decided not to distinguish between 641 the explicit directions of the cue and target motions and to 642 summarize the directions in the factor cue congruence. 643 Responses were coded in terms of accuracy. Thus, we estimated 644 an additional model with fixed effects of congruence and 645 ISI together with random effects of participant on the drift rate.

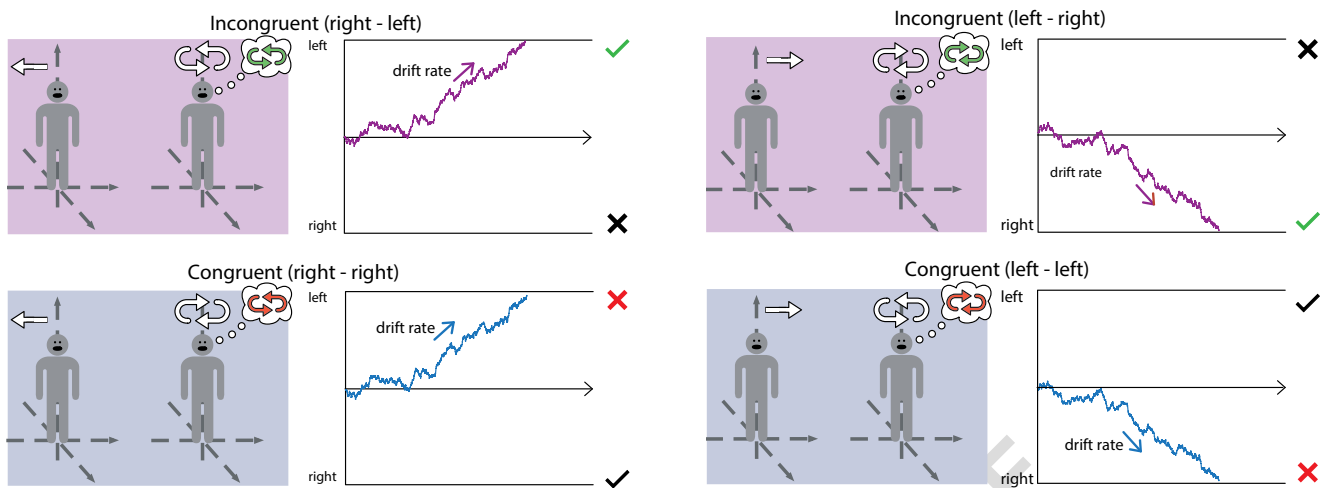


Fig. 7 Comparison of the posterior predictive distributions and the observed data from Experiment 2, for the model with the drift rate depending on cue and target motion but no shift in the starting point (on the left; no-bias model) and for the model with the drift rate depending on target motion and ISI and with the starting point depending on the cue motion in combination with ISI (on the right; bias model). The median predictions for response probabilities and median RTs averaged across all participants are plotted as black circles. The

bold gray bars represent 80% credible intervals, computed from the .1 and .9 quantiles. The thinner gray bars represent 95% intervals, computed from the .025 and .975 quantiles. Together, these indicate the uncertainty of the models' predictions, with larger intervals indicating more uncertainty. The means of the observed response probabilities and the medians of the observed RT data, averaged across all participants, are shown as crosses. The different ISIs are presented on the x-axis

Q9

646 For the starting point, we estimated a by-participant random
647 intercept model.

648 The goodness of fit for this model based on posterior pre-
649 dictive checks is presented in Fig. 4, right panels. The model
650 predicts participants' responses well, even better than in
651 Experiment 1, since the observed median response proba-
652 bilities (crosses) are both close to the predicted response proba-
653 bilities (black circles) and within the 80% and 95% intervals
654 of the predictive posterior distribution. The intervals of the
655 posterior predictive distributions are very narrow, even
656 narrower than in Experiment 1. The median RTs for correct
657 responses match the observed median RTs, as well. Again, the
658 model has difficulties describing the RTs for wrong responses,
659 especially in the incongruent conditions, where the observed
660 median RTs are at the upper end of the 80% and 95% intervals
661 of the posterior predictive distributions. This is probably due
662 to the fact that incorrect responses were rare in these
663 conditions.

664 The accuracy-coded DDM revealed that evidence for the
665 wrong response was accumulated when the cue and target
666 motions were spatially congruent for ISIs of 50, 100, and
667 200 ms. The parameter estimate for congruent cues and an
668 ISI of 600 ms indicates no systematic evidence accumulation
669 for either response (Fig. 5), and this is also demonstrated by
670 the fact that participants' accuracy was at chance level (Fig. 1).

671 **Discussion**

672 The results of the logistic regression clearly indicate self-
673 motion aftereffects for short ISIs, with participants

misperceiving the direction of translation in the congruent
674 trials. In Experiment 2, we have demonstrated that
675 intravestibular self-motion aftereffects are also present when
676 a yaw rotation is followed by a translation in the same spatial
677 direction. Thus, intravestibular self-motion aftereffects are bi-
678 directional: The processing of otolith stimulation is altered
679 depending on the direction of a previous stimulation of the
680 semicircular canals. Otolith and canal receptors function dif-
681 ferently, and yet we found a striking similarity in both combi-
682 nations (otolith preceded by canal stimulation in Exp. 2, and
683 the opposite order in Exp. 1). The RT data in Experiment 2
684 also show that participants generally responded faster on cor-
685 rect trials. However, this was less the case in the congruent
686 conditions for ISIs of 50, 100, and 200 ms—that is, when self-
687 motion aftereffects were more pronounced. In contrast to
688 Experiment 1, the RTs were slightly faster for correct than
689 for incorrect trials. 690

674 Q10

691 **Experiment 3**

692 **Method**

693 In the third experiment, we wanted to test whether self-motion
694 aftereffects were coupled with the intensity of the cue and
695 target motions. For rotations about the earth-vertical axis and
696 for horizontal y-axis translations, we chose the same intensi-
697 ties as in Experiment 1; however, the order of motion types
698 was the same as in Experiment 2. Thus, in proportion to the

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699 detection threshold, the cue stimulus was weaker than the
700 target stimulus, in contrast to Experiments 1 and 2.

701 **Participants** For Experiment 3, 12 new participants were re-
702 cruited. One participant pressed the left button on 92% of all
703 trials and was excluded from the study. The remaining 11
704 participants (nine female, two male; mean age 25, range 21–
705 38 years) were included in the further analysis. All participants
706 were right-handed according to a German version of the hand-
707 edness questionnaire by Chapman and Chapman (1987).
708 None of the participants reported a history of relevant neuro-
709 logical, vestibular, or attentional disorders. The study was ap-
710 proved by the ethics committee of the University of Bern, and
711 all participants gave written informed consent prior to the
712 experiment in accordance with the Declaration of Helsinki.

713 **Motion stimuli** The cue stimuli consisted of yaw rotations
714 about an earth-vertical axis (left/right). They consisted of sin-
715 gular cycles of sinusoidal acceleration at a frequency of 5 Hz.
716 The acceleration amplitude was 24 deg/s^2 ($v_{\text{max}} = 1.53 \text{ deg/s}$,
717 $\Delta p = 0.0153 \text{ deg}$). The target stimuli were translations along
718 the y -axis (left/right). They consisted of the same acceleration
719 profile as the cue stimuli. Their frequency was again 5 Hz, and
720 the acceleration amplitude was 0.25 m/s^2 ($v_{\text{max}} = 0.016 \text{ m/s}$,
721 $\Delta p = 0.0016 \text{ m}$).

722 **Experimental design and procedure** The design and experi-
723 mental procedure were identical to those of Experiment 2, but
724 using the motion intensities of Experiment 1.

725 **Data analysis** The statistical model for the accuracy data in
726 Experiment 3 did not differ from the analysis of Experiment 2.

727 Results

728 **Accuracy** The logistic regression model showed that the mo-
729 tion discrimination performance was around chance level (for
730 ISIs of 50 and 100 ms), or even above chance level (for ISIs
731 200 and 600 ms), when the prior stimulus was spatially con-
732 gruent with the subsequent motion (e.g., a leftward yaw rota-
733 tion followed by a left translation). Direction discrimination
734 performance for the second motion was higher when a yaw
735 rotation was followed by an incongruent translation. In this
736 experiment, there were no clear self-motion aftereffects (see
737 also Fig. 1, Exp. 3 panel).

738 Discussion

739 In Experiment 3 we wanted to investigate whether the same
740 self-motion aftereffect could be observed if the relative
741 strengths of the cue and target motions were reversed,
742 resulting in a cue that was relatively weaker than the target.
743 Although we observed that accuracy was impaired and that

participants performed at chance level, this manipulation did
not result in self-motion aftereffects.

General discussion

In Experiment 1, we showed that translations influence the
ability to discriminate the direction of yaw rotations, depend-
ing on the spatial congruency of the translations and rotations.
These same response patterns were also observed in
Experiment 2 when a yaw rotation preceded a translation.
These results are striking, in that participants systematically
misjudged strong motion stimuli well above the detection
threshold as a function of the preceding motion stimulus.
The RT data from Experiment 1 further suggest that in the
congruent conditions participants were confident about cor-
rectly detecting the direction of the target stimulus, as they
responded slightly faster on incorrect trials. Usually, incorrect
responses are accompanied by slower RTs, as was shown for
the 600-ms ISI and in the incongruent conditions in
Experiments 1 and 2, as well as in the neutral condition in
Experiment 1 for all ISIs. The application of a DDM to the
results of these two experiments revealed that erroneous self-
motion perception is explained better by an altered accumula-
tion of vestibular sensory evidence than by a biased starting
point. This implies that the altered perceptual decision-making
process and the resulting misperceptions are likely due to al-
tered processing of vestibular information. It is possible that
decision-making circuits dynamically accumulate evidence
for the wrong direction, because they receive less or no input
from the correct motion direction because its sensitivity has
been altered.

These results contradict the predictions made by a mecha-
nism that relies on attentional shifts by means of cueing. In
fact, from previous studies in other modalities (see Spence,
2010, for a review) one could have expected facilitating ef-
fects for spatially congruent trials with short ISIs and an IOR
for spatially congruent trials with long ISIs. In the present
study, participants showed a consistent response pattern that
indicates an actual misperception and not just longer response
times as one could predict within the framework of the spatial
cueing paradigm. Also, the comparison of the results from
Experiments 1 and 3 contradicts spatial cueing. Reversal of
the relative magnitude of cue and target intensity is not sup-
posed to alter the results as the directional information given
by the cue remains unchanged. However, the results from
Experiment 3 differ completely from those of Experiment 1,
and therefore, a mere shift of spatial attention can be ruled out.
This also suggests that sensory evidence accumulation is not
independent of the intensity of the first motion stimulus.

The present study has demonstrated intravestibular self-
motion aftereffects (from the otoliths to the SCCs, and vice
versa) for the first time. This self-motion aftereffect is

794 interesting for several reasons. The cue and target stimuli in
795 the present experiments relied on different sensory inputs—
796 that is, to the otoliths and SCCs. Previous research has shown
797 that otolith and SCC signals converge at an early stage of
798 processing in the vestibular nuclei (Carriot et al., 2015;
799 Dickman & Angelaki, 2002). Albeit linear and translational
800 angular movements are fundamentally different, we speculate
801 that, at this stage, a common spatial representation of the di-
802 rectional information implied by the rotations and translations
803 might be obtained by integrating rotational and translational
804 sensory afferents. The application of DDMs to the perceptual
805 decision-making process in self-motion aftereffects provided
806 confirmatory evidence. In fact, the drift rate changing as a
807 function of the congruence of the cue and target suggests that
808 the cue motion alters sensory evidence accumulation during
809 real-time vestibular processing. Because the cue and target
810 stimuli relied on different sensory systems, the evidence ac-
811 cumulation process in self-motion aftereffects can be
812 pinpointed to a stage at which the SCC and otolith signals
813 have already converged. A conceivable alternative would
814 have been a strategic bias that operated prior to processing
815 of the sensory evidence. However, this explanation can be
816 ruled out by the results of the DDM.

817 The observed phenomenon in Experiments 1 and 2 resem-
818 bles the description of MAEs in the vestibular modality
819 (Coniglio & Crane, 2014; Crane, 2012a, 2012b). Yet, the par-
820 adigm used in this study clearly differs from what is normally
821 used to investigate (vestibular) MAEs. Classically, the target
822 motion is adapted to the participants' previous responses and
823 the adaptor is presented longer. In contrast, in the present
824 experiments strong target stimuli were presented. Although
825 participants are capable of correctly perceiving the direction
826 of self-motion when it is preceded by a neutral or incongruent
827 motion, they consistently misperceive the direction if the pre-
828 ceding motion was congruent, despite the fact that the motion
829 intensity was well above the direction detection threshold
830 (Grabherr et al., 2008; Valko et al., 2012).

831 The temporal intervals between the two sequences of mo-
832 tion are a crucial factor for self-motion aftereffects. In the
833 present study, the ISIs were selected on the basis of the liter-
834 ature on the spatial-cueing paradigm, ranging from 50 to 600
835 ms. Again, as compared to the literature on vestibular MAEs,
836 these ISIs are rather short. For example, Crane (2012a) used
837 ISIs ranging from 500 to 3,000 ms and only found consistent
838 MAEs starting at 1,000 ms. Similarly, Coniglio and Crane
839 (2014) observed weaker MAEs with increasing ISIs. In con-
840 trast to those studies, here the self-motion aftereffects were
841 only observed at ISIs of 50, 100, and 200 ms. The short du-
842 ration of the cue stimuli in the present experiments marks the
843 most crucial difference from classical MAE studies: In fact,
844 whereas MAEs possibly arise due to adaptation at a neural
845 level, it is rather unlikely that cue stimuli of 200 ms could lead
846 to adaptation. Hence, although the outcomes of the present

847 experiments may resemble MAEs, the underlying mechanism
848 of this self-motion aftereffect is most likely not neural
849 adaptation.

850 The present data suggest that self-motion aftereffects arise
851 as a function of the interaction of stimulus duration, stimulus
852 intensity, and ISI. Important data to back up this claim stem
853 from Experiment 3 in the present study. In fact, in Experiment
854 3 the intensity relative to the threshold was weaker in the cue
855 than in the target motion. Thus, in comparison to Experiment
856 2, the intensities for the cue and target were reversed. This
857 subtle change resulted in virtually no self-motion aftereffect
858 in Experiment 3, although direction discrimination perfor-
859 mance was still worse in the congruent than in the incongruent
860 condition. Notably, previous studies on perceptual aftereffects
861 for prolonged asymmetric passive vestibular stimulation by
862 means of different velocities for different rotation directions
863 have reported asymmetric perceptual responses (Panichi et al.,
864 2011; Pettorossi et al., 2013): Where perceptual responses
865 were enhanced for fast rotations, reduced responses were ob-
866 served for slower rotations. These results underline the impor-
867 tance of the intensity of a cue stimulus in the perception of a
868 following vestibular stimulus.

869 The intravestibular self-motion aftereffects suggest that a
870 representation of spatial direction is obtained by combining
871 information from both otoliths and SCCs. The activation of
872 this representation through translations or rotations could lead
873 to an altered sensitivity in self-motion perception. This is in
874 accordance with a study by Nooij, Nesti, Bühlhoff, and Pretto
875 (2016), who found that the linear addition of translational and
876 rotational components of the sensory input does not suffice to
877 explain the perceived motion; instead, the components must
878 be combined in a more complex manner. Interestingly, two
879 other studies looking at the intravestibular interaction between
880 otoliths and SCCs found that translation detection thresholds
881 were increased with concurrent rotation (Crane, 2016;
882 MacNeilage et al., 2010). However, yaw rotation detection
883 thresholds were barely influenced by simultaneous transla-
884 tions. This seemingly unidirectional relationship underlines
885 the rather complex intravestibular interaction of the otoliths
886 and SCCs. These results are in line with our findings indicat-
887 ing that yaw rotations as the cue stimuli produced stronger
888 self-motion aftereffects.

889 Our results suggest that altered sensitivity produces self-
890 motion aftereffects at very short ISIs, whereas at an ISI of
891 600 ms the effect disappears. The results from Experiment 3
892 also imply that the amount of sensitivity alteration depends on
893 the intensity of the vestibular stimulation. Importantly, altered
894 sensitivity is in line with altered evidence accumulation.
895 Future studies aiming at better understanding the mechanisms
896 underlying self-motion aftereffects should systematically
897 study the interaction of the ISI, duration, and intensity of cue
898 and target motions, to delineate the limits of the self-motion
899 aftereffect. In particular, the notion of attention in the context

900 of passive self-motion needs to be refined (cf. Figliozzi et al.,
 901 2005). Attention can be viewed as a gain control mechanism
 902 that modulates sensory processing (Hillyard, Vogel, & Luck,
 903 1998). Brown, Friston, and Bestmann (2011) have discussed
 904 this mechanism in the context of motor planning and active
 905 inference, and it would be beneficial to investigate the poten-
 906 tial role of attentional modulation as a gain control mechanism
 907 in the context of evidence accumulation and self-motion
 908 aftereffects.

909 Another important aspect is the influence of proprio-
 910 ception. Participants in typical self-motion aftereffect
 911 studies are seated and fixated on a motion platform. Q
 912 Accelerations and decelerations not only stimulate the
 913 vestibular system, but also generate proprioceptive cues.
 914 The processing of those proprioceptive signals could inter-
 915 fere with the vestibular afferent signals of the second
 916 motion and lead to incorrect information processing,
 917 which would result in altered sensory evidence accumula-
 918 tion. Thus, the reported self-motion aftereffects might oc-
 919 cur at a processing stage at which vestibular information
 920 has been combined with proprioceptive information. The
 921 misperception of the direction in spatially congruent con-
 922 ditions could foster exploratory behavior similar to the
 923 phenomenon of spontaneous alternation behavior
 924 (Vecera, Rothbart, & Posner, 1991). Spontaneous alterna-
 925 tion behavior is a memory-driven tendency to avoid pre-
 926 viously selected spatial locations in order to explore the
 927 environment. We speculate that the observed altered evi-
 928 dence accumulation process might lead to a similar behav-
 929 ior. However, it remains to be investigated whether and
 930 how these two phenomena interact.

931 In conclusion, we have shown intravestibular self-motion
 932 aftereffects and suggest that this effect may be rooted in al-
 933 tered evidence accumulation in the perceptual decision-
 934 making process at the level of a common representation of
 935 spatial direction from the SCCs and otoliths. Higher-level de-
 936 cision-making processes have been rather neglected in vestib-
 937 ular psychophysics for a long time. However, in the last few
 938 years several studies have acknowledged the importance of
 939 such higher-level aspects in self-motion perception and have
 940 started to investigate the underlying processes in vestibular
 941 (Clark et al., 2018; Ellis, Klaus, & Mast, 2017; Merfeld,
 942 Clark, Lu, & Karmali, 2016; Wertheim, Mesland, & Bles,
 943 2001) and multisensory (Drugowitsch, DeAngelis, Angelaki,
 944 & Pouget, 2015; Drugowitsch, DeAngelis, Klier, Angelaki, &
 945 Pouget, 2014; Lim, Wang, & Merfeld, 2017) perceptual deci-
 946 sion-making. Yet, they have important theoretical and practi-
 947 cal implications regarding the complex nature of biological
 948 self-motion perception. In daily life, we are constantly ex-
 949 posed to stimuli targeting the otoliths and semicircular canals,
 950 both simultaneously and in rapid succession. Here we showed
 951 the necessity to further investigate such sequences of motions,
 952 since they can alter perceptual decision-making processes.

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AUTHOR QUERIES

AUTHOR PLEASE ANSWER ALL QUERIES.

- Q1. Note reordering of Tables 1&2, per order of treatment in paper. 
- Q2. Please check rephrase of models, here and in Fig. 3, as well as in Exp. 2 and Fig. 7. 
- Q3. The few color refs. to the figs. (in the main text and captions) are removed or rephrased to also work with grayscale print versions. 
- Q4. Bias model description rephrased for (I hope) clarity. Description of regression structure in captions was difficult to parse, so please check paraphrases here and in Fig. 7 carefully. 
- Q5. Note that Figs. 4–7 are reordered, also per order of treatment in paper. Please check that figs. still correspond to correct captions. 
- Q6. Do rephrased colors work with grayscale fig.? 
- Q7. Rephrased for clarity. OK? 
- Q8. Please check rephrase of models. 
- Q9. Rephrase OK? 
- Q10. Changed from “yaw rotation” per Exp. 2 Method, and following sentence here. OK? 
- Q11. Sentence elaborated for clarity. OK? 

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