

Impact of Antiretroviral Therapy on Tuberculosis Incidence Among HIV-Positive Patients in High-Income Countries

The HIV-CAUSAL Collaboration^a

Background. The lower tuberculosis incidence reported in human immunodeficiency virus (HIV)-positive individuals receiving combined antiretroviral therapy (cART) is difficult to interpret causally. Furthermore, the role of unmasking immune reconstitution inflammatory syndrome (IRIS) is unclear. We aim to estimate the effect of cART on tuberculosis incidence in HIV-positive individuals in high-income countries.

Methods. The HIV-CAUSAL Collaboration consisted of 12 cohorts from the United States and Europe of HIV-positive, ART-naive, AIDS-free individuals aged ≥ 18 years with baseline CD4 cell count and HIV RNA levels followed up from 1996 through 2007. We estimated hazard ratios (HRs) for cART versus no cART, adjusted for time-varying CD4 cell count and HIV RNA level via inverse probability weighting.

Results. Of 65 121 individuals, 712 developed tuberculosis over 28 months of median follow-up (incidence, 3.0 cases per 1000 person-years). The HR for tuberculosis for cART versus no cART was 0.56 (95% confidence interval [CI], 0.44–0.72) overall, 1.04 (95% CI, 0.64–1.68) for individuals aged >50 years, and 1.46 (95% CI, 0.70–3.04) for people with a CD4 cell count of <50 cells/ μL . Compared with people who had not started cART, HRs differed by time since cART initiation: 1.36 (95% CI, 0.98–1.89) for initiation <3 months ago and 0.44 (95% CI, 0.34–0.58) for initiation ≥ 3 months ago. Compared with people who had not initiated cART, HRs <3 months after cART initiation were 0.67 (95% CI, 0.38–1.18), 1.51 (95% CI, 0.98–2.31), and 3.20 (95% CI, 1.34–7.60) for people <35 , 35–50, and >50 years old, respectively, and 2.30 (95% CI, 1.03–5.14) for people with a CD4 cell count of <50 cells/ μL .

Conclusions. Tuberculosis incidence decreased after cART initiation but not among people >50 years old or with CD4 cell counts of <50 cells/ μL . Despite an overall decrease in tuberculosis incidence, the increased rate during 3 months of ART suggests unmasking IRIS.

Tuberculosis is the most common AIDS-defining condition worldwide [1, 2]. The use of combined antiretroviral therapy (cART) against human immunodeficiency virus (HIV) is one of the strategies recommended by the World Health Organization (WHO) to prevent tuberculosis [1–3]. However, cART initiation may lead to immune reconstitution inflammatory syndrome (IRIS), a poorly understood immunological phenomenon [4, 5]

that is particularly frequent in patients infected by *Mycobacterium tuberculosis*. IRIS may lead to either the paradoxical worsening of previously treated tuberculosis or the “unmasking” of subclinical tuberculosis for up to 3 months following cART initiation [4, 5]. Inflammatory signs of unmasked tuberculosis may range from mild to severe, the latter being IRIS in sensu stricto [5].

A randomized clinical trial in Haiti [6] and observational studies from low-income [7, 8], middle-income [9–11], and high-income [12–20] settings have found lower tuberculosis incidence among HIV-positive patients receiving cART. However, none of these studies differentiated between short- and long-term effects of cART, most included prevalent cART users, and all but 2 [8, 20] did not appropriately adjust for measured time-dependent confounders.

We estimated the effect of cART initiation on tuberculosis incidence in HIV-positive individuals in

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a large collaboration from Europe and the United States. To overcome some of the limitations of previous studies, we restricted our analyses to cART-naive patients and used inverse probability weighting to appropriately adjust for measured confounders. We present these estimates stratified by time since cART initiation, to explore short-term increases in tuberculosis incidence compatible with unmasking IRIS. For comparison purposes, we estimated the effect of cART on *Pneumocystis jirovecii* pneumonia (PCP), a condition for which cART has been shown to be protective [12] and for which IRIS is infrequent [5].

METHODS

Study Population

The HIV-CAUSAL Collaboration has been described elsewhere [21]. Briefly, the collaboration includes 12 prospective cohort studies from 6 European countries and the United States: UK CHIC (United Kingdom), UK Register of HIV Seroconverters (United Kingdom), ATHENA (the Netherlands), FHDH-ANRS CO4 (France), ANRS PRIMO (France), ANRS-SEROCO (France), SHCS (Switzerland), GEMES (Spain), PISCIS (Spain), CoRIS/CoRIS-MD (Spain), AMACS (Greece), and VACS-VC (United States). All cohorts are based on data collected for clinical purposes within healthcare systems with universal access to care.

The HIV-CAUSAL Collaboration includes 71 743 HIV-positive antiretroviral-naive individuals aged ≥ 18 years. Analyses were restricted to those who (1) had adequate baseline data to adjust for confounding (having CD4 cell count and HIV RNA level measurements within 6 months of each other) and (2) had no diagnosis of an AIDS-defining illness (including tuberculosis) at baseline. In all, 188 patients were excluded because of condition 1 and 6434 because of condition 2, leaving 65 121 eligible patients.

An individual's baseline was defined as the first time during 1996–2007 when all of the above criteria were met. An individual's follow-up ended at tuberculosis diagnosis, death, 12 months after the most recent laboratory measurement, or cohort-specific administrative censoring (from December 2003 through September 2007), whichever occurred earliest. Our analyses were restricted to HIV-positive individuals who did not have a tuberculosis diagnosis and did not start cART during the first month after baseline.

Initiation of cART was defined as the date on which an individual initiated treatment with ≥ 3 antiretroviral drugs, 2 ritonavir-boosted protease inhibitors, or 1 nonnucleoside reverse-transcriptase inhibitor plus 1 boosted protease inhibitor. Alternative definitions of cART did not affect the overall effect estimate. PCP prophylaxis was generally initiated when the CD4 cell count dropped to < 200 cells/ μL , in accordance with international guidelines.

The primary outcome was tuberculosis diagnosis (pulmonary and extrapulmonary combined), although separate analyses for pulmonary tuberculosis and extrapulmonary tuberculosis were also conducted. Diagnostic procedures for tuberculosis in participating cohorts reflect standard clinical practice in Europe and the United States, with most diagnoses reported to be definite (generally defined as isolation of *M. tuberculosis* from a culture of any of the following clinical samples: sputum, bronchial washings, pleural fluid, lymph node, bone marrow aspirate, or cerebrospinal fluid) rather than presumptive (generally defined as clinical or radiological signs suggestive of tuberculosis, exclusion of other causes, and clinical response to tuberculosis treatment). Skin tests for *M. tuberculosis* infection were not systematically performed and isoniazid preventive treatment was not routinely implemented, reflecting clinical practice in most of Europe and the United States.

Statistical Methods

Incidence rates of tuberculosis were calculated as tuberculosis diagnoses per 1000 person-years. We estimated the average hazard ratio (HR) of tuberculosis for cART initiation versus noninitiation via a pooled logistic model for risk of tuberculosis at month $m + 1$ that included a time-varying indicator for ever use of cART through month m , month of follow-up m (restricted cubic splines with 5 knots), and the following baseline covariates: CD4 cell count (< 50 , 50–99, 100–199, 200–349, 350–499, or ≥ 500 cells/ μL), HIV RNA level ($< 10\,000$, 10 000–100 000, or $> 100\,000$ copies/mL), sex, transmission group (heterosexual, men who have sex with men, injection drug users, or other/unknown), calendar year (1996–1998, 1999–2000, or 2001–2007), age (< 35 , 35–50, or > 50 years), geographic origin (Western countries, sub-Saharan Africa, other countries, or unknown), time since HIV infection diagnosis (< 3 or ≥ 3 months), and cohort. To estimate the HR as a function of time since cART initiation, we fit a separate model in which the indicator for cART use was replaced by a time-varying treatment variable with 3 categories: no cART initiation, < 3 months after cART initiation, or ≥ 3 months after cART initiation. This cutoff was based in the consensus clinical case definition for tuberculosis-associated IRIS issued by the International Network for the Study of HIV-associated IRIS [22].

Because cART is more likely to be initiated in individuals with a low CD4 cell count and a high viral RNA level, estimates from the above regression models have to be adjusted for this time-dependent confounding by indication. To do so, one could add the time-varying confounders CD4 cell count and HIV RNA level as covariates in the logistic regression model. However, this standard approach may introduce bias because the confounders are affected by prior cART use and are on the causal pathway between cART and tuberculosis. We therefore used

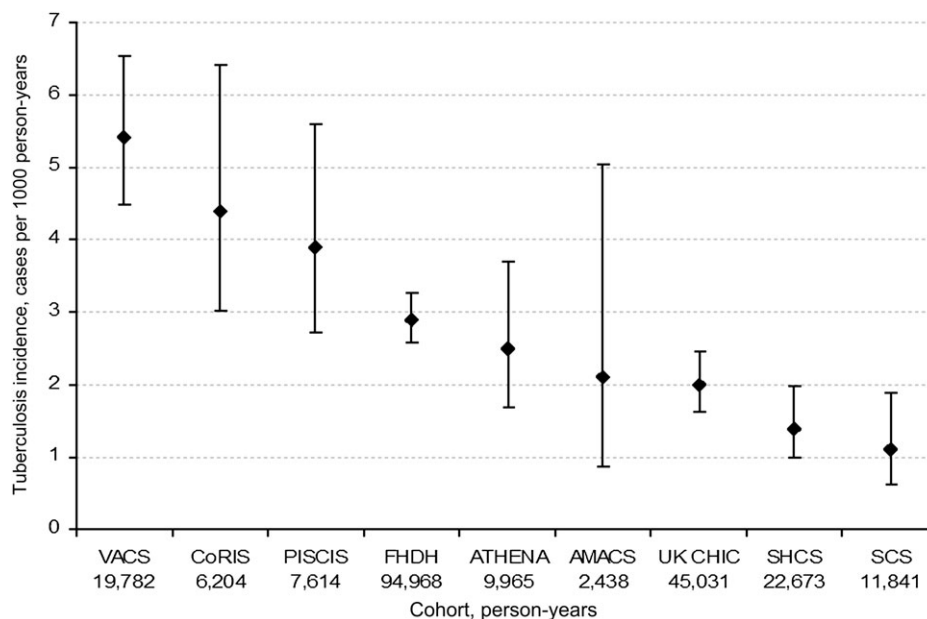


Figure 1. Tuberculosis incidence rates by cohort, HIV-CAUSAL Collaboration, 1996–2007. The countries (no. of eligible individuals) associated with each cohort are as follows: VACS, United States (7281; definite tuberculosis cases only are included); CoRIS, Spain (3198); PISCIS, Spain (2988); FHDH, France (26 089); ATHENA, Netherlands (4316); AMACS, Greece (750); UK CHIC, United Kingdom (13 183); SHCS, Switzerland (4510); and SCS, seroconverters from the UK Register of HIV Seroconverters in the United Kingdom, PRIMO/SEROCO in France, and GEMES in Spain (2806).

inverse probability weighting to adjust for measured time-dependent confounders that are affected by prior cART use. Formally, under the assumption that all time-varying predictors of both cART and tuberculosis were included in the analyses, our weighted model estimates the parameters of a marginal structural Cox model [23].

Each patient in the above logistic models received a time-varying weight inversely proportional to the probability of having their own observed history of cART initiation, as described elsewhere [24]. To estimate each patient's probability of cART initiation in each month, we fit a pooled logistic model that included the covariates listed above for the outcome model plus the most recent measurement of the following time-varying covariates: CD4 cell count (restricted cubic splines with 6 knots), HIV RNA level (3 categories), AIDS (yes or no), and time since last laboratory measurement (5 categories). To better adjust for cohort-specific factors, the models also included separate product terms between cohort and time of follow-up, time-varying CD4 cell count, and HIV RNA level. Inverse probability weights were also computed to adjust for potential selection bias due to censoring by infrequent measurement. Both the cART initiation and censoring weights were stabilized and their product used to fit the weighted regression model. To avoid undue influence of outliers on the variance of the estimates, we truncated weights at a maximum value of 10. Truncation did not materially change the point estimates. The estimated weights

used in the analyses had a mean of 1.04. We computed 95% confidence intervals (CIs) for the log HR of tuberculosis by adding and subtracting 1.96 times the standard error, where the standard error was the square root of a variance estimator that accounts for the weight-estimation procedure.

Several sensitivity analyses were also performed: (1) in addition to censoring individuals at 12 months without a laboratory measurement, we censored at 18 and 24 months after the last measurement; (2) the start of follow-up was delayed by 3 months to exclude early tuberculosis diagnoses (and thus ensure that only incident tuberculosis cases were included); (3) the weights were re-estimated by adding the rates of CD4 cell count and HIV RNA level change to the models; (4) the weights were re-estimated by lagging the CD4 cell count and HIV RNA level 7 or 14 days to ensure that cART initiation was predicted using prior laboratory measurements; and (5) additional inverse probability weights for censoring by death were estimated.

We fit stratum-specific weighted models to conduct subgroup analyses according to the following baseline characteristics: sex, age, HIV RNA level, CD4 cell count, transmission group, geographical origin, and year. For comparison purposes, we repeated the main analyses after replacing the outcome tuberculosis by PCP. We chose PCP because it is a frequent condition for which cART has been shown to be very effective and for which IRIS is infrequent. All analyses were conducted with SAS, version 9.2.

Table 1. Characteristics of Study Participants and Hazard Ratio of Tuberculosis for Combined Antiretroviral Therapy Initiation Versus No Initiation, HIV-CAUSAL Collaboration, 1996–2007

Baseline Characteristic	Individuals, no.	Median Follow-up Time, mo (IQR)	Person-Years, no.	Tuberculosis Cases, no.	Tuberculosis Incidence, Cases Per 1000 Person-Years	cART Initiators, no. (%)	HR of Tuberculosis ^a (95% CI)
CD4 cell count, cells/μL							
<50	2288	27 (11–59)	7610	64	8.4	1809 (79)	1.46 (0.70–3.04)
50–99	2081	31 (13–64)	7425	41	5.5	1712 (82)	0.68 (0.28–1.63)
100–199	6029	30 (13–62)	21 298	104	4.9	4777 (79)	0.39 (0.22–0.71)
200–349	15 387	28 (13–58)	52 974	199	3.8	10 340 (67)	0.59 (0.37–0.94)
350–499	16 759	28 (13–57)	57 662	148	2.6	8133 (49)	0.40 (0.26–0.63)
≥500	22 577	26 (12–54)	73 546	156	2.1	6972 (31)	0.68 (0.43–1.08)
HIV RNA level, copies/mL							
<10 000	20 479	25 (11–51)	64 387	145	2.3	6877 (34)	0.80 (0.48–1.35)
10 000–100 000	29 870	28 (13–58)	102 717	357	3.5	16 252 (54)	0.55 (0.40–0.75)
>100 000	14 772	31 (14–63)	53 411	210	3.9	10 614 (72)	0.51 (0.31–0.82)
Age, years							
<35	28 522	29 (13–62)	103 618	248	2.4	14 316 (50)	0.33 (0.21–0.51)
35–50	30 497	27 (12–54)	96 774	368	3.8	15 759 (52)	0.72 (0.53–1.00)
>50	6102	28 (12–56)	20 123	96	4.8	3668 (60)	1.04 (0.64–1.68)
Sex							
Female	16 157	27 (12–55)	53 252	182	3.4	8526 (53)	0.39 (0.24–0.63)
Male	48 964	28 (13–58)	167 263	530	3.2	25 217 (52)	0.62 (0.46–0.83)
Transmission group							
Heterosexual	21 397	28 (13–56)	71 923	284	3.9	11 718 (55)	0.52 (0.36–0.77)
Men who have sex with men	24 827	31 (14–63)	91 643	95	1.0	12 591 (51)	0.29 (0.13–0.61)
Injection drug users	7175	24 (11–53)	23 623	78	3.3	3314 (46)	0.39 (0.18–0.86)
Other/unknown	11 722	22 (11–47)	33 326	255	7.7	6120 (53)	0.89 (0.63–1.27)
Geographical origin							
Western countries	31 491	30 (14–62)	114 999	206	1.8	16 235 (52)	0.57 (0.35–0.91)
Sub-Saharan Africa	8072	25 (12–50)	23 799	169	7.1	4477 (55)	0.50 (0.30–0.83)
Other countries	4516	22 (11–45)	12 264	51	4.2	2136 (47)	0.81 (0.36–1.81)
Unknown	21 042	27 (12–57)	69 454	286	4.1	10 895 (52)	0.61 (0.42–0.89)
Calendar period							
1996–1998	19 400	41 (16–106)	95 665	258	2.7	11 109 (57)	0.62 (0.40–0.98)
1999–2000	8987	39 (15–83)	36 192	134	3.7	5296 (59)	0.53 (0.31–0.88)
2001–2007	36 734	23 (11–42)	88 658	320	3.6	17 338 (47)	0.52 (0.37–0.74)
Overall	65 121	28 (13–57)	220 515	712	3.2	33 743 (52)	0.56 (0.44–0.72)

Abbreviations: cART, combined antiretroviral therapy; CI, confidence interval; HIV, human immunodeficiency virus; HR, hazard ratio; IQR, interquartile range.

^a Estimated from a weighted pooled logistic model that included a time-varying indicator for ever use of cART, month of follow-up (restricted cubic splines with 5 knots), and the baseline covariates listed in the table. The inverse probability weights were a function of the baseline covariates and most recent measurement of the following time-varying covariates: CD4 cell count (restricted cubic splines with 6 knots), HIV RNA level (3 categories), AIDS (yes or no), and time since last laboratory measurement (5 categories).

RESULTS

Our study included 65 121 HIV-positive individuals who met eligibility criteria from 1996 through 2007. The median follow-up was 28 months (interquartile range, 13–57 months). During follow-up, 712 individuals received a diagnosis of tuberculosis (493 had pulmonary and 296 had extrapulmonary tuberculosis) and 2275 individuals died. The overall

tuberculosis incidence was 3.0 cases per 1000 person-years and varied by cohort (Figure 1).

Tuberculosis incidence decreased with CD4 cell count and increased with viral load and age. Rates were similar for men and women, men who have sex with men had the lowest rates, and people of sub-Saharan African origin had about 4 times the rate of people born in Western countries (Table 1). Overall, the HR of tuberculosis for cART initiation versus no initiation

Table 2. Hazard Ratio of Tuberculosis for Combined Antiretroviral Therapy Initiation Versus No Initiation by Time Since Initiation, HIV-CAUSAL Collaboration, 1996–2007

Baseline Characteristic	No cART Initiation		Time Since cART Initiation, <3 mo		Time Since cART Initiation, ≥3 Mo	
	Cases, no.	HR ^a	Cases, no.	HR ^a (95% CI)	Cases, no.	HR ^a (95% CI)
CD4 cell count, cells/μL						
<50	14	1 (ref)	15	2.30 (1.03–5.14)	35	1.02 (0.44–2.36)
50–99	10	1 (ref)	8	1.17 (0.37–3.70)	23	0.46 (0.22–0.99)
100–199	37	1 (ref)	14	0.94 (0.47–1.86)	53	0.28 (0.15–0.54)
200–349	88	1 (ref)	21	1.50 (0.84–2.69)	90	0.45 (0.27–0.75)
≥350	195	1 (ref)	10	0.70 (0.30–1.62)	99	0.49 (0.35–0.70)
HIV RNA level, copies/mL						
<10 000	92	1 (ref)	8	2.05 (0.89–4.74)	45	0.68 (0.39–1.17)
10 000–100 000	173	1 (ref)	26	1.27 (0.76–2.14)	158	0.46 (0.33–0.64)
>100 000	79	1 (ref)	34	1.10 (0.70–1.74)	97	0.36 (0.20–0.62)
Age, years						
<35	131	1 (ref)	16	0.67 (0.38–1.18)	101	0.29 (0.19–0.47)
35–50	170	1 (ref)	38	1.51 (0.98–2.31)	160	0.59 (0.42–0.83)
>50	43	1 (ref)	14	3.20 (1.34–7.60)	39	0.59 (0.34–1.04)
Sex						
Female	96	1 (ref)	11	0.65 (.33, 1.25)	75	0.35 (0.21–0.59)
Male	248	1 (ref)	57	1.64 (1.13–2.37)	225	0.47 (0.33–0.65)
Transmission group						
Heterosexual	138	1 (ref)	30	1.35 (0.84–2.17)	116	0.39 (0.26–0.60)
Men who have sex with men	48	1 (ref)	12	1.19 (0.46–3.12)	35	0.20 (0.09–0.44)
Injection drug users	41	1 (ref)	7	1.23 (0.45–3.37)	30	0.27 (0.13–0.60)
Other/unknown	117	1 (ref)	19	1.48 (0.85–2.56)	119	0.79 (0.54–1.15)
Geographical origin						
Western countries	100	1 (ref)	21	1.63 (0.86–3.10)	85	0.42 (0.25–0.71)
Sub-Saharan Africa	84	1 (ref)	17	1.10 (0.55–2.17)	68	0.39 (0.22–0.71)
Other countries	26	1 (ref)	6	1.70 (0.68–4.25)	19	0.57 (0.23–1.42)
Unknown	134	1 (ref)	24	1.36 (0.83–2.22)	128	0.52 (0.35–0.76)
Calendar period						
1996–1998	117	1 (ref)	19	1.46 (0.81–2.64)	122	0.53 (0.34–0.85)
1999–2000	57	1 (ref)	13	1.69 (0.77–3.71)	64	0.37 (0.22–0.64)
2001–2007	170	1 (ref)	36	1.08 (0.69–1.69)	114	0.41 (0.28–0.68)
Overall	344	1 (ref)	68	1.36 (0.98–1.89)	300	0.44 (0.34–0.58)

Abbreviations: cART, combined antiretroviral therapy; CI, confidence interval; HIV, human immunodeficiency virus; HR, hazard ratio; ref, reference value.

^a Estimated from a weighted pooled logistic model that included a time-varying treatment variable with 3 categories (no cART initiation, <3 months after cART initiation, or ≥3 months after cART initiation), month of follow-up (restricted cubic splines with 5 knots), and the baseline covariates listed in the table. The inverse probability weights were a function of the baseline covariates and most recent measurement of the following time-varying covariates: CD4 cell count (restricted cubic splines with 6 knots), HIV RNA level (3 categories), AIDS (yes or no), and time since last laboratory measurement (5 categories).

was 0.56 (95% CI, 0.44–0.72). All HRs were <1 except for people whose baseline CD4 cell count was <50 cells/μL (HR, 1.46; 95% CI, 0.70–3.04) and for those aged >50 years (HR, 1.04; 95% CI, 0.64–1.68). The overall HR of pulmonary tuberculosis was 0.58 (95% CI, 0.44–0.77) and that of extrapulmonary tuberculosis was 0.45 (95% CI, 0.29–0.68).

The HRs for tuberculosis differed by time since cART initiation: 1.36 (95% CI, 0.98–1.89) <3 months after initiation and 0.44 (95% CI, 0.34–0.58) ≥3 months after initiation (Table 2). These estimates were similar for pulmonary or extrapulmonary tuberculosis: the HR for pulmonary tuberculosis was 1.31 (95%

CI, 0.87–1.96) for <3 months and 0.47 (95% CI, 0.35–0.63) for ≥3 months, and the HR for extrapulmonary tuberculosis was 1.31 for <3 months (95% CI, 0.84–2.05) and 0.34 for ≥3 months (95% CI, 0.22–0.53). The corresponding estimates were similar when we excluded the cohort with the highest tuberculosis incidence (ie, VACS): 1.33 (95% CI, 0.91–1.94) for <3 months and 0.35 (95% CI, 0.24–0.49) for ≥3 months. The elevated tuberculosis risk <3 months after cART initiation was found in men with a baseline CD4 cell count of <50 cells/μL and HIV RNA load of <10 000 copies/mL. This early risk increased with age. Tuberculosis incidence was lower ≥3 months after

Table 3. Hazard Ratio of *Pneumocystis jirovecii* Pneumonia for Combined Antiretroviral Therapy Initiation Versus No Initiation by Time Since Initiation and Selected Characteristics, HIV-CAUSAL Collaboration, 1996–2007

Baseline Characteristic	No cART Initiation		<3 Months Since cART Initiation		≥3 Months After cART Initiation	
	No. of Cases	HR ^a	No. of Cases	HR ^a (95% CI)	No. of Cases	HR ^a (95% CI)
CD4 cell count, cells/μL						
<50	57	1 (ref)	9	0.32 (0.15–0.67)	38	0.30 (0.17–0.51)
50–99	26	1 (ref)	5	0.33 (0.12–0.92)	31	0.18 (0.06–0.40)
100–199	43	1 (ref)	7	0.60 (0.22–1.64)	40	0.19 (0.10–0.39)
200–349	59	1 (ref)	8	0.62 (0.25–1.58)	46	0.06 (0.03–0.14)
≥350	121	1 (ref)	4	0.22 (0.07–0.65)	45	0.16 (0.09–0.28)
HIV RNA level, copies/mL						
<10 000	44	1 (ref)	3	0.81 (0.25–2.67)	23	0.31 (0.13–0.75)
10 000–100,000	126	1 (ref)	13	0.64 (0.31–1.31)	74	0.07 (0.04–0.13)
>100 000	136	1 (ref)	17	0.37 (0.20–0.67)	103	0.15 (0.09–0.25)
Age, years						
<35	96	1 (ref)	9	0.43 (0.20–0.95)	72	0.15 (0.08–0.29)
35–50	166	1 (ref)	21	0.47 (0.26–0.84)	99	0.08 (0.05–0.14)
>50	44	1 (ref)	3	0.51 (0.14–1.93)	29	0.13 (0.05–0.32)
Sex						
Female	49	1 (ref)	6	0.50 (0.17–1.46)	27	0.04 (0.02–0.10)
Male	257	1 (ref)	27	0.48 (0.29–0.77)	173	0.12 (0.08–0.19)
Overall	306	1 (ref)	33	0.48 (0.31–0.75)	200	0.10 (0.07–0.16)

Abbreviations: cART, combined antiretroviral therapy; CI, confidence interval; HIV, human immunodeficiency virus; HR, hazard ratio; ref, reference value.

^a Estimated from a weighted pooled logistic model that included a time-varying treatment variable with 3 categories (no cART initiation, <3 months after cART initiation, or ≥3 months after cART initiation), month of follow-up (restricted cubic splines with 5 knots), and the baseline covariates. The inverse probability weights were a function of the baseline covariates and most recent measurement of the following time-varying covariates: CD4 cell count (restricted cubic splines with 6 knots), HIV RNA level (3 categories), AIDS (yes or no), and time since last laboratory measurement (5 categories).

cART initiation, compared with untreated individuals, in all subgroups except individuals with baseline CD4 cell count of <50 cells/μL (HR, 1.02; 95% CI, 0.44–2.36). Sensitivity analyses did not materially affect the estimates.

During follow-up, 539 individuals received a diagnosis of PCP; the overall PCP incidence was 2.4 cases per 1000 person-years. The overall HR of PCP was 0.13 (95% CI, 0.09–0.20) for cART initiation versus no initiation; the HR varied by time since initiation within 3 months (HR, 0.48; 95% CI, 0.31–0.75) and after 3 months (HR, 0.10; 95% CI, 0.07–0.16), but not by age and CD4 cell count (Table 3). PCP prophylaxis was received by 24% of individuals; further adjustment for prophylaxis therapy did not materially change the results.

The probability of cART initiation was inversely related to CD4 cell count and directly related to HIV RNA levels. Compared with CD4 cell count of ≥500 cells/μL, the HR for cART initiation was 1.11, 2.7, 8.3, and 14.3 for CD4 cell count in the range 350–499, 200–349, 100–199, and <100 cells/μL, respectively. Compared with an HIV RNA level of <10 000 copies/mL, the HR for cART initiation was 1.4 and 2.2 for HIV RNA levels of 10 000–100 000 and >100 000 copies/mL, respectively. Because CD4 cell count and HIV RNA level are affected by prior cART use, conventional methods for confounding adjustment are expected to result in biased estimates.

As an illustration of this bias, Figure 2 shows the HR of tuberculosis from unweighted Cox models that adjusted for baseline covariates only (HR, 0.81; 95% CI, 0.67–0.97) and for both baseline and time-dependent covariates (HR, 1.03; 95% CI, 0.86–1.24).

DISCUSSION

This study shows a reduction of 44% in tuberculosis incidence among HIV-positive individuals who start cART in high-income countries. However, we found no evidence of reduction among those >50 years old or with CD4 cell counts <50 cells/μL. Despite the net lower risk of tuberculosis after cART initiation, tuberculosis incidence increased by 36% <3 months after cART initiation. This increased early risk of tuberculosis was higher in people with baseline CD4 cell counts of <50 cells/μL, in men, and in individuals ≥35 years old. An early increase in the risk of PCP among people receiving cART was not observed in any subgroup.

The reduction in tuberculosis incidence attributable to cART found in our study is closer to the null than most previously reported estimates from observational studies in other settings [7, 12–15, 19], possibly because of the selection criteria we imposed to avoid inadvertent classification of prevalent tuberculosis as incident tuberculosis. Individuals who received

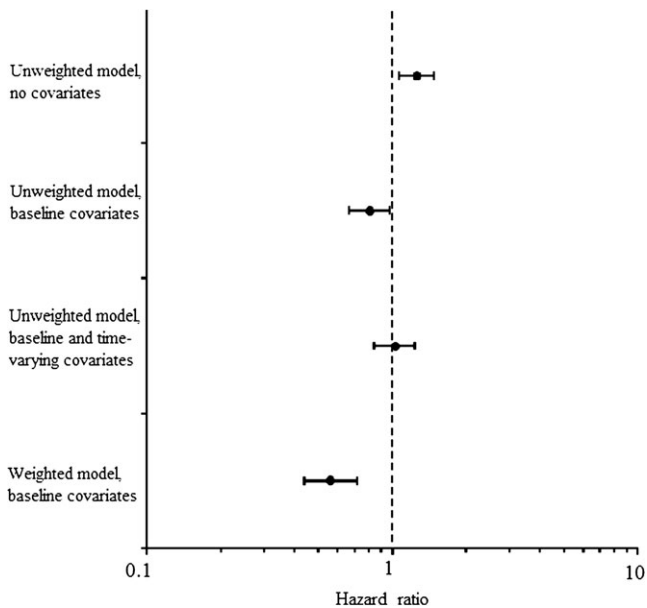


Figure 2. Hazard ratios of tuberculosis from conventional (unweighted) and inverse probability weighted models, HIV-CAUSAL Collaboration, 1996–2007.

a diagnosis of tuberculosis (206 cases) within the first month of follow-up were excluded to prevent the inclusion of patients with prevalent tuberculosis cases who would have otherwise been assigned to the untreated group, overestimating the beneficial effect of cART. Indeed, including these cases resulted in HRs of tuberculosis of 1.02 (95% CI, 0.79–1.32) <3 months after cART initiation and 0.43 (95% CI, 0.34–0.54) \geq 3 months after cART initiation, compared with no cART initiation.

Our study did not find a lower risk of tuberculosis after cART initiation among people with a baseline CD4 cell count of ≤ 50 cells/ μ L and those who were aged >50 years. This could be due to a higher tuberculosis incidence in the first 3 months following cART or to impaired immunological responses to *M. tuberculosis* in very immunosuppressed patients who start cART. We cannot distinguish between incident tuberculosis, unmasking tuberculosis IRIS, and unmasking tuberculosis without IRIS because individual tuberculosis diagnoses could not be assessed to confirm which individuals presented exaggerated inflammatory responses. However, the epidemiological pattern suggests that unmasking IRIS may be playing a role [4, 5, 25, 26]. A meta-analysis [5] found that the incidence of tuberculosis-associated paradoxical worsening IRIS is highest in individuals with a CD4 cell count <50 cells/ μ L, but data on unmasking IRIS are scarce. In contrast, a randomized trial in Haiti did not find an increased tuberculosis incidence after cART initiation, but all participants were screened for *M. tuberculosis* infection and started treatment with isoniazid if clinically indicated [6]. Lawn et al [27] have recently reported

that an intensive screening strategy before cART initiation markedly decreased the incidence of tuberculosis and IRIS.

After the first three months, overall tuberculosis incidence was 56% lower in the people on cART than in people not on cART, although no decrease was observed in individuals with baseline CD4 cell counts <50 cells/ μ L. Our findings suggest that very immunosuppressed individuals cannot lower their overall risk of tuberculosis by taking cART for ≥ 3 months. These patients remain in a state of relevant immunodeficiency in spite of cART for longer periods and are therefore at higher risk of reactivating latent tuberculosis beyond 3 months after starting cART. Further follow-up of these cohorts will show whether the risk of tuberculosis does eventually drop with longer exposure to cART.

Tuberculosis incidence soon after cART initiation increased with age, especially in those aged >50 years. Although the prevalence of *M. tuberculosis* infection, and thus the risk of reactivating disease, increases with age [28], the large increase in tuberculosis incidence among individuals receiving cART was unexpected. One possible explanation is the existence of an age-associated IRIS: HIV infection induces a premature aging of the immune system, and people >50 years of age have poorer immunological responses to cART, which may contribute to the reactivation of latent tuberculosis [29].

The reduction in tuberculosis incidence attributable to cART may vary depending on the background risk of *M. tuberculosis* transmission, as people may develop tuberculosis following recent infection and/or reinfection in settings with high transmission rates [28, 30]. We do not have information on the background risk of tuberculosis transmission by cohort, and country tuberculosis rates are unlikely to be a good marker. However, despite different background risks of *M. tuberculosis* infection, the effect estimate of cART on tuberculosis incidence obtained from South Africa is similar to ours [8].

For comparison purposes we repeated the same analyses for PCP, a condition for which IRIS is uncommon [5]. We found a lower PCP incidence at all times after cART. The decreased risk of PCP during the first 3 months of cART contrasts with the increased risk of tuberculosis during the same period and supports the role of unmasking IRIS in the latter condition.

Our study has several limitations. First, cART was not randomly assigned, and thus the possibility of residual confounding cannot be ruled out. In particular, treatment initiation might have been prompted by unrecorded symptoms (eg, weight loss due to undiagnosed tuberculosis), which might partly explain the high risk of tuberculosis <3 months after cART initiation. To reduce this potential confounding, we excluded individuals with tuberculosis at or within 1 month of baseline. Although symptoms of undiagnosed tuberculosis might have prompted cART initiation in some individuals, the same applies to PCP symptoms because both conditions may have insidious presentations.

Our findings have implications for future research, clinical practice, and public health practice. Impaired immunological responses to *M. tuberculosis* in people starting cART at very low CD4 cell counts and those >50 years of age warrant further investigation. These data support the WHO STOP TB partnership statement that multiple interventions are needed to control tuberculosis in HIV-positive people. For cART to be useful in tuberculosis control, prompt HIV infection testing is needed so cART can be started when clinically indicated rather than at advanced immunodeficiency. Diagnosis and treatment of *M. tuberculosis* infection, as national and international guidelines recommend, should be reinforced. Even in countries with low tuberculosis transmission, intensified case finding remains essential; physicians should be aware of the potentially increased risk of developing tuberculosis after cART initiation.

Supplementary Data

Supplementary materials are available at *Clinical Infectious Diseases* online (http://www.oxfordjournals.org/our_journals/cid/). Supplementary materials consist of data provided by the author that are published to benefit the reader. The posted materials are not copyedited. The contents of all supplementary data are the sole responsibility of the authors. Questions or messages regarding errors should be addressed to the author.

Notes

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References

1. Dye C, Watt CJ, Bleed DM, Hosseini SM, Raviglione MC. Evolution of tuberculosis control and prospects for reducing tuberculosis incidence, prevalence, and deaths globally. *JAMA* **2005**; 293:2767–75.
2. STOP TB Partnership. The global plan to stop TB 2006–2015. Available at: <http://www.stoptb.org/>. Accessed September 2011.
3. Lawn SD, Wood R, De Cock KM, Kranzer K, Lewis JJ, Churchyard GJ. Antiretrovirals and isoniazid preventive therapy in the prevention of HIV-associated tuberculosis in settings with limited health-care resources. *Lancet Infect Dis* **2010**; 10:489–98.
4. Lawn SD, Bekker LG, Miller RF. Immune reconstitution disease associated with mycobacterial infections in HIV-infected individuals receiving antiretrovirals. *Lancet Infect Dis* **2005**; 5:361–73.
5. Müller M, Wandel S, Colebunders R, Attia S, Furrer H, Egger M. IeDEA Southern and Central Africa. Immune reconstitution inflammatory syndrome in patients starting antiretroviral therapy for HIV infection: a systematic review and meta-analysis. *Lancet Infect Dis* **2010**; 10:251–61.
6. Severe P, Juste MA, Ambrose A, et al. Early versus standard antiretroviral therapy for HIV-infected adults in Haiti. *N Engl J Med* **2010**; 363:257–65.
7. Badri M, Wilson D, Wood R. Effect of highly active antiretroviral therapy on incidence of tuberculosis in South Africa: a cohort study. *Lancet* **2002**; 359:2059–64.
8. Fairall LR, Bachmann MO, Louwagie GM, et al. Effectiveness of antiretroviral treatment in a South African program: a cohort study. *Arch Intern Med* **2008**; 168:86–93.
9. Miranda A, Morgan M, Jamal L, et al. Impact of antiretroviral therapy on the incidence of tuberculosis: the Brazilian experience, 1995–2001. *PLoS One* **2007**; 2:e826.
10. Golub JE, Saraceni V, Cavalcante SC, et al. The impact of antiretroviral therapy and isoniazid preventive therapy on tuberculosis incidence in HIV-infected patients in Rio de Janeiro, Brazil. *AIDS* **2007**; 21:1441–8.
11. Santoro-Lopes G, Felix de Pinho A, Harrison LH, Schechter M. Reduced risk of tuberculosis among Brazilian patients with advanced human immunodeficiency virus infection treated with highly active antiretroviral therapy. *Clin Infect Dis* **2002**; 34:543–6.
12. Ledergerber B, Egger M, Erard V, et al. AIDS-related opportunistic illnesses occurring after initiation of potent antiretroviral therapy: the Swiss HIV Cohort Study. *JAMA* **1999**; 282:2220–6.
13. Kirk O, Gatell JM, Mocroft A, et al. Infections with *Mycobacterium tuberculosis* and *Mycobacterium avium* among HIV-infected patients after the introduction of highly active antiretroviral therapy. EuroSIDA Study Group JD. *Am J Respir Crit Care Med* **2000**; 162:865–72.
14. Girardi E, Antonucci G, Vanacore P, et al. Impact of combination antiretroviral therapy on the risk of tuberculosis among persons with HIV infection. *AIDS* **2000**; 14:1985–91.
15. Moreno S, Jarrin I, Iribarren JA, et al. Incidence and risk factors for tuberculosis in HIV-positive subjects by HAART status. *Int J Tuberc Lung Dis* **2008**; 12:1393–400.
16. Abgrall S, Del Giudice P, Melica G, Costagliola D. FHDH-ANRS CO4. HIV-associated tuberculosis and immigration in a high-income country: incidence trends and risk factors in recent years. *AIDS* **2010**; 24:763–71.
17. Grant AD, Bansi L, Ainsworth J, et al. Tuberculosis among people with HIV infection in the United Kingdom: opportunities for prevention? *AIDS* **2009**; 23:2507–15.
18. Elzi L, Schlegel M, Weber R, et al. Reducing tuberculosis incidence by tuberculin skin testing, preventive treatment, and antiretroviral

- therapy in an area of low tuberculosis transmission. *Clin Infect Dis* **2007**; 44:94–102.
19. Jones JL, Hanson DL, Dworkin MS, et al. HIV-associated tuberculosis in the era of highly active antiretroviral therapy. The Adult/Adolescent Spectrum of HIV Disease Group. *Int J Tuberc Lung Dis* **2000**; 4:1026–31.
 20. Pettit AC, Jenkins CA, Stinnette SE, et al. Tuberculosis risk before and after highly active antiretroviral therapy initiation: does HAART increase the short-term TB risk in a low incidence TB setting? *J Acquir Immune Defic Syndr* **2011**; 57:305–10.
 21. HIV-CAUSAL Collaboration. The effect of combined antiretroviral therapy on the overall mortality of HIV-infected individuals. *AIDS* **2010**; 24:123–37.
 22. Meintjes G, Lawn SD, Scano F, et al. for International Network for the Study of HIV-associated IRIS. Tuberculosis-associated immune reconstitution inflammatory syndrome: case definitions for use in resource-limited settings. *Lancet Infect Dis* **2008**; 8:516–23.
 23. Hernán MA, Hernández-Díaz S, Robins JM. A structural approach to selection bias. *Epidemiology* **2004**; 15:615–25.
 24. Robins JM, Hernán MA, Brumback B. Marginal structural models and causal inference in epidemiology. *Epidemiology* **2000**; 11:550–60.
 25. Lawn S, Wood R, Wilkinson R. Changing concepts of “latent tuberculosis infection” in patients living with HIV infection. *Clin Dev Immunol* **2011**; 2011:980594. Epub 2010.
 26. Manabe YC, Breen R, Perti T, Girardi E, Sterling TR. Unmasked tuberculosis and tuberculosis immune reconstitution inflammatory disease: a disease spectrum after initiation of antiretroviral therapy. *J Infect Dis* **2009**; 199:437–44.
 27. Lawn S, Kranzer K, Edwards D, McNally M, Bekker L, Wood R. Tuberculosis during the first year of antiretroviral therapy in a South African cohort using an intensive pretreatment screening strategy. *AIDS* **2010**; 24:1323–8.
 28. Styblo K. Epidemiology of tuberculosis. Selected papers. *KNCV*, **1991**; 24:9–111.
 29. Martin J, Volberding P. HIV and premature aging: a field still in its infancy. *Ann Intern Med* **2010**; 153:477–9.
 30. Alland D, Kalkut GE, Moss AR, et al. Transmission of tuberculosis in New York City. *N Engl J Med* **1994**; 330:1710–6.

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