

# IUPAC-IUGS recommendation on the half-lives of $^{147}\text{Sm}$ and $^{146}\text{Sm}$

I.M. Villa<sup>1,2,4,5</sup>, N.E. Holden<sup>1,3,6</sup>, A. Possolo<sup>1,3,7</sup>, R.B. Ickert<sup>1,2,8</sup>, D.B. Hibbert<sup>1,3,9</sup>, P.R. Renne<sup>1,2,10,11</sup>

<sup>1</sup>Joint IUPAC-IUGS Task Group on isotope data in geosciences; <sup>2</sup>International Union of Geological Sciences; <sup>3</sup>International Union of Pure and Applied Chemistry; <sup>4</sup>Institut für Geologie, Universität Bern, CH-3012 Bern, Switzerland; e-mail igor@geo.unibe.ch; <sup>5</sup>Centro Universitario Datazioni e Archeometria, Università di Milano Bicocca, I-20126 Milano, Italy; <sup>6</sup>National Nuclear Data Center, Brookhaven National Laboratory, Upton, NY 11973, USA; <sup>7</sup>Statistical Engineering Division, National Institute of Standards & Technology, Gaithersburg, MD 20877, USA; <sup>8</sup>Department of Earth, Atmospheric, and Planetary Sciences, Purdue University, West Lafayette, IN, 47907, USA; <sup>9</sup>School of Chemistry, University of New South Wales, Sydney, NSW 2052, Australia; <sup>10</sup>Berkeley Geochronology Center, 2455 Ridge Road, Berkeley, CA 94720, USA; <sup>11</sup>Dept. of Earth and Planetary Science, University of California at Berkeley, Berkeley, CA 94720.

## Abstract

The IUPAC-IUGS joint Task Group “Isotopes in Geosciences” recommends a value of  $(106.25 \pm 0.38)$  Ga for the half-life of  $^{147}\text{Sm}$ , and a corresponding decay constant  $\lambda_{147} = (6.524 \pm 0.024) \times 10^{-12} \text{ a}^{-1}$ , both with a coverage factor  $k = 2$ . For the extinct radionuclide  $^{146}\text{Sm}$  two very different half-lives are used in the scientific community (*c.* 68 and 103 Ma), to such a degree that no consensus value can be endorsed at present by the Task Group. Pending dedicated re-investigations it is recommended that papers using the  $^{146}\text{Sm}$  decay to quantify the cosmo/geological evolution of (extra)terrestrial samples perform a twin set of calculations using both proposed half-lives.

## Introduction

The two isotopes of Sm most relevant for geochronological applications are  $^{147}\text{Sm}$  and  $^{146}\text{Sm}$  (Lugmair et al., 1975a, 1975b; Lugmair and Marti, 1977, 1978). Both are  $\alpha$  emitters, which makes the discrimination of true disintegration events from noise straightforward. However, analytical challenges stem from the half-lives of these two isotopes:  $^{147}\text{Sm}$  has a half-life,  $t_{1/2}$ , in excess of 100 Ga, resulting in a very low count rate, whereas  $^{146}\text{Sm}$  has  $t_{1/2} \leq 0.12$  Ga and is extinct in the Solar System, so that it must be synthesized before it can be counted. The  $\alpha$  counting literature of the last 60 years includes 14 papers dealing with  $^{147}\text{Sm}$  and four with  $^{146}\text{Sm}$ . While all estimates agree on the order of magnitude, the individual half-life estimates differ by up to 30 %, despite one order of magnitude smaller published uncertainty estimates.

This situation requires a decision on how to distinguish acceptable measurements from incorrect ones. A first approach is to apply Gaussian statistics (e.g. Rajput and MacMahon, 1992) on the entire data-base from the literature, and then optimize the dispersion (quantified by the reduced chi square,  $\chi^2_n$ , defined as the variance-weighted residuals normalized by the number of degrees of freedom, in the case of a weighted average the number of measurements minus one) by progressively removing measurements until a  $\chi^2_n$  value acceptably near the expectation value of 1 is achieved. This statistical procedure identifies outliers, irrespective of the existence of experimental flaws. The  $\chi^2_n$  is identical to the MSWD parameter (mean square of weighted deviates: McIntyre et al., 1966), widely used in geochronology. The requirement that  $\chi^2_n \leq 1$  is explained by the value of the  $p$  parameter, the “probability of fit”. If  $\chi^2_n > 1$ , the dispersion of the measured points from the calculated ones (be it by linear regression or by averaging) is larger than the repeatability of the measurement; the data are then said to be overdispersed. This implies that the dispersion of values remains uncorrected for systematic errors, which include, possibly alongside other contributions, “incomplete knowledge of certain physical phenomena” (JCGM, 2008, §D.4). If  $\chi^2_n > 1$ , this also means that the likelihood,  $p$ , that the calculation accurately reflects the process that caused the distribution of the measured points becomes progressively smaller. The more typical examples

52 involve treating bimodal distributions as if they were unimodal, and linearly modeling a two-stage  
53 geological process as if it were single-stage. Since the likelihood that the function used for the  
54 calculation correctly predicts the observed dispersion is  $p$ , the chance that the calculation is  
55 incorrect is by definition  $1-p$ ; for a  $\chi^2_n$  so high that  $p = 0.05$  the chance of misidentifying the  
56 process controlling the distribution of the data points is 95 %.

57 An entirely different approach is to identify flawed experiments affected by systematic errors. To  
58 this end one aims to distinguish and evaluate the repeatability of measurements, typically via Type  
59 A evaluations (JCGM, 2012, entries 2.15 and 2.26), and the presence of systematic errors, typically  
60 via Type B evaluations "based on experience" (JCGM, 2012, entries 2.17 and 2.26). The IUPAC-  
61 IUGS Task Group has carried out Type B "expert" evaluations of the measurement protocols, and  
62 this is reflected in the present recommendation.

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### $^{147}\text{Sm}$

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66 In the last 60 years a total of 14 counting experiments and one theoretical determination of the  
67  $^{147}\text{Sm}$  half-life have been published. What emerges is that one technique, liquid scintillation  
68 counting (LSC), gives four concordant measurement results ( $106.4 \pm 0.8$  Ga, with  $\chi^2_n = 0.50$ ). Two  
69 other techniques appear at first sight to show a bimodal distribution, the ionization chamber and the  
70 silicon surface barrier. This bimodality suggests that one should look for a substantive explanation  
71 for this artifact rather than gloss over it by merely averaging all the measured values. In Appendix  
72 A, one exemplary dataset illustrates the paradoxical, and incorrect, conclusions of a pure Type A  
73 evaluation, in contrast to the Type B evaluation performed by the experimenters themselves.

74 The very authors of some experiments that have been shown to suffer from "substantive,  
75 documented" inaccuracy (Koepke et al, 2017a, §3.P1) had already set the corresponding results  
76 aside: (a) the Helsinki team (Valli et al., 1965) determined a half-life of  $108 \pm 2$  Ga, superseding the  
77 earlier ionization chamber measurements by the same team (Karras and Nurmia, 1960; Graeffe and  
78 Nurmia, 1961); (b) Gupta and MacFarlane (1970) repeated, refined and replaced the earlier  
79 ionization chamber experiment by MacFarlane and Kohman (1961); (c) the silicon surface barrier  
80 measurement by Kinoshita et al. (2003) was later ignored by Kinoshita et al. (2012) who preferred  
81 using the 106 Ga half-life obtained in other experiments. These authors took great care to repeat  
82 their experiments eliminating factors that would lead to systematic errors after learning from past  
83 mistakes. It would be unfair (and statistically incorrect) to still include the superseded outliers into  
84 an unsupervised average by arbitrarily inflating the originally stated Type A uncertainty so as to  
85 transform a bimodal distribution with peaks around 116 Ga and 106 Ga (the 116 Ga peak being  
86 caused by incorrect experimental protocols) into a broader, unimodal Gaussian distribution, whose  
87 average and standard deviation offer no accurate information. The same argument applies to the use  
88 of Delrin plastic holders (see Appendix A), which were not a "random effect" to be propagated into  
89 an increased Dark Uncertainty, but pure and simply a flawed design, to which the argument made  
90 by Koepke et al. (2017b, p. 61, lines 7-8) must be applied. The remaining 10 measurements are  
91 listed in Table 1, grouped by experimental approaches. As the techniques are very different, any  
92 systematic error affecting one of them could not affect the other ones.

93 In addition to the counting experiments, Tavares and Terranova (2018) derived the half-life of  
94  $^{147}\text{Sm}$  from a first-principles calculation based on the quantum tunneling through a modelled  
95 potential barrier for the nucleus (Table 1). Any calculation relative to a nucleus with  $> 100$  nucleons  
96 requires approximated models of multibody interactions; the accuracy of the model assumptions is  
97 not charted in all cases. What is certain is that a potential inaccuracy of the nuclear model would  
98 produce a bias completely unrelated to those that could have affected experimental studies. In the  
99 case of  $^{147}\text{Sm}$ , the agreement with the experimental data is a very strong indication that no major  
100 inaccuracies affected the assumptions used for the calculation.

101 The fact that the results in Table 1 are mutually consistent is strong evidence that none of them is  
102 significantly biased. Indeed, an unintended benefit of the selection performed by the experimental  
103 groups themselves is that the statistical evaluation is that of a unimodal distribution (the reduced  $\chi^2_n$

104 = 0.28; Cochran's Q-test (Koepeke et al., 2017a) yields a p-value of 0.99). Therefore, the  
 105 conventional weighted average is appropriate, and yields a well-defined consensus value of  $106.61$   
 106  $\pm 0.99$  Ga ( $k=2$ , for approximate 95 % coverage).

107  
 108 **Table 1.** Post-1958 direct determinations of the  $^{147}\text{Sm}$  half-life. Only the 10 experimental papers not  
 109 renounced by the original authors in later publications are considered. Uncertainties are taken from the  
 110 original publication, with a coverage factor  $k = 1$ . LSC, liquid scintillation counter; IoC, ionization chamber;  
 111 SSB, silicon surface barrier; Pro, proportional counter; Cr39, CR-39 detector. The four papers marked with  
 112 an asterisk were those averaged by Lugmair and Marti (1978) to derive the half-life currently in use by the  
 113 geochronological community,  $106 \pm 1.6$  Ga. The original result by Martins et al. (1992) was corrected as  
 114 explained by Begemann et al. (2001, p. 117). One theoretical calculation, marked as Calc, is also shown.  
 115

Method	reference	half-life (Ga)
LSC	Beard and Kelly (1958)	$106 \pm 4$
LSC	Wright et al. (1961) *	$105 \pm 2$
LSC	Donhoffer (1964) *	$104 \pm 3$
LSC	Kossert et al. (2009)	$107.0 \pm 0.9$
IoC	Valli et al. (1965) *	$108 \pm 2$
IoC	Gupta and MacFarlane (1970) *	$106 \pm 2$
IoC	Wilsenach et al. (2017)	$107.4 \pm 1.9$
SSB	Su et al. (2010)	$106.6 \pm 0.8$
Pro	Al Bataina and Janecke (1987)	$105 \pm 4$
Cr39	Martins et al. (1992) **	$106 \pm 4$
Calc	Tavares and Terranova (2018)	$108 \pm 3$

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 119 **Table 2.** Determinations of the  $^{147}\text{Sm}$  half-life by cosmo- and geochronological intercomparison.  
 120 Comparisons based on meteoritic samples and terrestrial samples are labelled "met" and "terr",  
 121 respectively. Only samples with acceptably low intra-sample dispersion, i.e.  $\chi^2_n < 1$ , are listed. The  
 122 uncertainties, with a coverage factor  $k = 2$ , are either taken from the original publication, or recalculated  
 123 from the original data and the U-Pb or Pb-Pb age of the same rock, propagating the uncertainty on the U  
 124 half-lives (Villa et al., 2016).  
 125

Rock	reference for Sm-Nd age	reference U-Pb / Pb-Pb age	half-life (Ga)
terr	De Paolo and Wasserburg (1979)	Wall et al. (2016)	$106.3 \pm 0.9$
met	Jacobsen and Wasserburg (1984)	Tissot et al. (2017)	$105.8 \pm 0.8$
met	Lugmair and Galer (1992)	Amelin (2008)	$106.1 \pm 0.8$
terr	Amelin and Semenov (1996)	Amelin et al. (1995)	$106.4 \pm 0.8$

126  
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 128 Cosmo- and geochronological intercomparisons are a half-life determination totally independent  
 129 from alpha counting. Begemann et al. (2001) had concluded that geological intercomparisons are  
 130 consistent with a  $(106.0 \pm 1.6)$  Ga half-life (Lugmair and Marti, 1978) but not with the ca. 116 Ga  
 131 half-life proposed in a part of the literature. The 106.0 Ga half-life is currently used by the  
 132 geochronological community. In this evaluation we treat intercomparisons separately from counting  
 133 experiments, as the metrological traceability only applies to the analytical part of the Sm-Nd and U-  
 134 Pb (or Pb-Pb) age determination, but not to establishing the equivalence of the two age  
 135 determinations. The equivalence is inherently a Type B evaluation (JCGM, 2012, § 2.29), as the  
 136 intercomparison relies on the mineral phases being cogenetic, formed at the same time (or in a short  
 137 time interval), and remaining closed to isotope exchange after formation; the evidence of such ideal  
 138 "point-like" geological histories can only be established after their measurement. As some of the

139 samples dated by Sm-Nd in the literature were also dated with other decay systems, the ideality of  
 140 the point-like behavior can be better assessed. In the present evaluation we filtered the available  
 141 cosmo- and geochronological intercomparisons by the requirement that the data define an  
 142 overdetermined internal isochron having a statistically acceptable dispersion, i.e.  $\chi^2_n \leq 1$ , for all  
 143 reported isotopic systems. This rigour is justified by the profound difference between "estimating  
 144 the age of a sample" and "establishing a sample as an acceptable calibrator reference material for  
 145 universal use". In the former case, most practitioners will accept some data overdispersion, even if  
 146  $\chi^2_n > 1$  means that the likelihood that the result is accurate drops below 50 % (in regional geology  
 147 studies it is sometimes tolerated that overdispersed isochrons having a near 0 % likelihood of  
 148 accurately describing the distribution of data-points as due to a single-stage process be considered  
 149 fit-for-purpose to the extent that they provide a first estimate of the approximate age of an otherwise  
 150 undated sample). When the purpose of a measurement is dating a single sample, any inaccurate  
 151 assumption only affects the accuracy of that sample. On the contrary, the selection of a reference  
 152 material (JCGM, 2012, § 5.13) must be much more strict, as any systematic inaccuracy cascades on  
 153 the accuracy on all measurements that rely on it (e.g. Potts, 2010). A geological intercomparison  
 154 using natural rock samples as reference materials is only as good as the least homogeneous and the  
 155 least pristine of the rocks in the comparison chain; therefore, the requirements for absence of any  
 156 secondary modification must be much more strict than for general dating purposes.

157 Table 2 lists the intercomparisons from the literature that satisfy these *a priori* criteria. The  
 158 weighted average of the four intercomparisons listed in Table 2 is  $t_{1/2} = 106.22 \pm 0.38$  Ga,  $\chi^2_n = 0.4$ .  
 159 In summary, the weighted average of the determinations of the  $^{147}\text{Sm}$  half-life listed in Tables 1-2 is  
 160  $(106.25 \pm 0.38)$  Ga, with a coverage factor  $k = 2$ . The corresponding decay constant is  $\lambda_{147} = (6.524$   
 161  $\pm 0.024) \times 10^{-12} \text{ a}^{-1}$ . The uncertainty on the half-life here recommended is about half of that of the  
 162 most precise age of individual samples. In practical terms, the uncertainty is not completely  
 163 negligible but is unlikely to become the dominant source of composite uncertainty with the present  
 164 generation of analytical equipment. If and when the technical improvements will make it possible to  
 165 reduce the measurement uncertainty consistently below that of the present recommendation, there  
 166 will be scope for a renewed evaluation that will have to be based on the new, improved age  
 167 intercomparisons - provided that  $\chi^2_n$  is always  $< 1$  for the selected overdetermined isochron  
 168 calculations.

## 146Sm

174 The number of publications reporting  $^{146}\text{Sm}$  half-life measurements is quite small (Table 3). The  
 175 most relevant problem is that the distribution of estimated half-life values is bimodal, whereby the  
 176 two most discrepant half-lives estimates differ by a factor of 1.5. Two counting experiments give  
 177 compatible results in the uncertainty interval (with a coverage factor  $k = 2$ ) between 57 and 81 Ma,  
 178 whereas two counting experiments agree in an uncertainty interval (with a coverage factor  $k = 2$ )  
 179 between 93 and 112.2 Ma. This discrepancy needs to be addressed in detail.

181 Table 3. Determinations of the  $^{146}\text{Sm}$  half-life, with a coverage factor  $k = 1$ . Four counting experiments are  
 182 shown, with uncertainties taken from the original publications. The entry marked by an asterisk is the half-  
 183 life calculation by Qian and Ren (2014), whose uncertainty estimate was not provided by its authors.

reference	half-life (Ma)
Nurmia et al. (1964)	$74 \pm 15$
Friedman et al. (1966)	$102.6 \pm 4.8$
Meissner et al. (1987)	$103.1 \pm 4.5$
Kinoshita et al. (2012)	$68 \pm 7$
Qian and Ren (2014)*	69

185 The quest for systematic artifacts is especially demanding. All papers state that they have estimated  
186 and taken into account all possible biases, but this obviously cannot be true for both contrasting  
187 modes of the bimodal distribution of half-life estimates. None of the early counting experiments  
188 actually provides a comprehensive description of all possible systematic artefacts. While limiting  
189 the descriptions of experimental details was commonplace in printed journals until the age of  
190 Electronic Supplements, it makes the evaluation of legacy papers nearly impossible.

191 From the point of view of counting, one can note that the calculated half-life is inversely  
192 proportional to the number of detected events. Counting artifacts can include both false positives  
193 (detecting a signal caused by an interference and not corresponding to a genuine  $^{146}\text{Sm}$  decay) and  
194 false negatives (not detecting a decay). Detecting a spurious 2.55 MeV signal is presumably  
195 unlikely; however, other possible artefacts can bias a counting experiment more severely than  
196 counting alone. Issues such as purity, stoichiometry and precise isotopic composition of the  
197 artificial  $^{146}\text{Sm}$  source have not been adequately described by any of the four counting experiments  
198 listed in Table 3. Similarly to the recent "philosophical" shift that in U-Pb dating the non-isotopic  
199 information on petrogenesis is at least as important than mass spectrometer precision, the  
200 information on material purity and isotopic composition is at least as important as counting.

201 Friedman et al. (1966) followed two different experimental designs, so as to remove all artifacts that  
202 they considered: in one design, they determined the  $^{145}\text{Sm}/^{146}\text{Sm}$  ratio, and in the other they  
203 "doped"  $^{146}\text{Sm}$  with natural Sm. Meissner et al. (1987) were primarily interested in the systematics  
204 of alpha reduced widths and energies in a comparison of all nuclides with 84 neutrons; the half-life  
205 measurements of some of these nuclides were a by-product of the main goal of their experiments.  
206 Seeing as their estimate of the  $^{146}\text{Sm}$  half-life was of the right order of magnitude and compared  
207 favorably with the two preceding literature results, they did not further discuss the possibility of  
208 artifacts that could have biased the half-life of just one among the nuclides that they were  
209 addressing, namely  $^{146}\text{Sm}$ . Kinoshita et al. (2012) evaluated several possible causes for systematic  
210 error and concluded that none had affected their results significantly. They were first in addressing  
211 the possibility of isobaric interference of  $^{146}\text{Nd}$  impurities during the mass spectrometric  
212 determination of the amount of  $^{146}\text{Sm}$ . However, they did not document the stoichiometric purity of  
213 their Sm solution, nor did they determine its Sm isotopic composition by up-to-date analytical  
214 protocols.

215 Irradiated Sm targets were also used by Nurmia et al. (1964), who produced  $^{146}\text{Sm}$  by  $^3\text{He}$   
216 irradiation of  $^{147}\text{Sm}$ , and by Meissner et al. (1987), who produced  $^{146}\text{Sm}$  via  $^{146}\text{Eu}$  by  $^2\text{H}$  irradiation  
217 of  $^{147}\text{Sm}$ . Friedman et al. (1966) irradiated a Nd target to generate  $^{146}\text{Sm}$ . In their experiment the  
218  $^{146}\text{Sm}$  half-life is independent from the Sm(dope)/Sm(sample) ratio. This implies that the supposed  
219 contaminant  $^{146}\text{Nd}$  was not an impurity contained in the added Sm dopant, but probably was the  
220 result of ion exchange resins providing a less complete removal of the irradiation target than the  
221 100.0000 % achieved by AMS (Kinoshita et al., 2012). The counting experiments should, in  
222 principle (Begemann et al. 2001), allow metrological traceability to a reference material having a  
223 known decay rate and a known mass fraction of the radioactive nuclide. Despite the foregoing  
224 explanation of possible discrepancies among the counting experiments, their very low number  
225 frustrates the attempt of a reliable evaluation of uncertainty.

226 The fifth entry in Table 3 is a theoretical calculation by Qian and Ren (2014) using a density-  
227 dependent cluster model. As these authors did not attempt to quantify the possible bias inherent in  
228 the model and did not provide an uncertainty estimate, their calculation cannot be compared with  
229 the counting experiments and cannot be evaluated in the present paper.

230 A number of studies proposed to evaluate the  $^{146}\text{Sm}$  half-life by Early Solar System cosmo- and  
231 geochronological intercomparisons. A difficulty with such intercomparisons is that they all depend  
232 on very extensive assumptions on the ideality of single-stage evolution of the analyzed systems and  
233 on the uniform incorporation of one uniform  $^{146}\text{Sm}/^{147}\text{Sm}$  ratio in the entire Solar System, none of  
234 which is amenable to a metrological tracing: there is neither a reference material for an  
235 extraterrestrial object with a single-stage point-like history, nor for its  $^{146}\text{Sm}$  mass fraction, and  
236 therefore no propagation of uncertainty down the traceability chain. Mass spectrometric analyses

237 *per se* are traceable to standard reference materials (as are the response functions of the gamma  
238 counters of the experiments in Table 3), but assumptions regarding the "point-like" history of early  
239 Solar System samples are not (just as the missing experimental details of the experiments in Table 3  
240 are not). Traceability ensures that the measurement is accurate and that any other laboratory, when  
241 repeating the measurement under the same conditions, will reproduce the same number within the  
242 stated uncertainty. What traceability to an agreed reference material cannot guarantee is that the  
243 points in an isochron diagram pertain to a sample in petrologic equilibrium, nor can it guarantee that  
244 the isochron reflects the sample's single-stage formation age. A detailed discussion of results based  
245 on early Solar System samples is given in Appendix B.

246 There appears to be an unbridged gap between one part of the scientific community (e.g. Audi et al.,  
247 2017, p. 81), which only endorses the alpha-counting data (short half-life) measured by Kinoshita et  
248 al. (2012) and calculated by Qian and Ren (2014), versus another part of the community, which  
249 mostly accepts the long half-life measured by Meissner et al. (1987). At this time the  
250 inconsistencies evident in Table 3 cannot be pinpointed with certainty to a specific, documented  
251 systematic error. In the face of this dilemma, the Task Group must refrain from making a  
252 recommendation that satisfies both camps. If one seeks a consensus interval from the present data  
253 using the DerSimonian-Laird procedure as implemented in the NIST Consensus Builder (Koepke et  
254 al., 2017a, § 5.2), choosing the version of the uncertainty analysis that uses an adjustment (Knapp-  
255 Hartung) for the fact that the number of measurement results being combined is very small, and  
256 excluding the calculation by Qian and Ren (2014) because it is not qualified with an uncertainty  
257 evaluation, the resulting 95 % coverage interval ranges from 59 Ma to 119 Ma. This interval is too  
258 wide to be of much use; its width means that it is meaningless to try to blend such a heterogeneous  
259 set of measurement results into a single consensus value.

260 It is hoped that this review, by laying open all issues on undocumented systematic effects, will  
261 stimulate future work to improve the non-isotopic characterization (for the cosmochronological  
262 intercomparison) and non-counting characterization (for the counting experiments). A consensus on  
263 the half-life might only be reached if new, dedicated experiments will take into account the  
264 systematic errors identified over recent years. We recommend that until then all  
265 cosmo/geochronological reconstructions follow the interpretive approach by Sanborn et al. (2015,  
266 p. 92-93), who avoided firmly choosing between the  $68 \pm 7$  Ma and the  $103 \pm 4$  Ma half-life for  
267  $^{146}\text{Sm}$  and performed their calculations in double, comparing the results obtained using both  
268 conflicting assumptions.

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## 376 377 378 **Appendix A**

379  
380 Wilsenach et al (2017, their Table 3) describe two different types of holders on which the thin  
 381  $^{147}\text{Sm}$  film is evaporated: metallic silicon and a plastic, Delrin. The five Delrin holder runs, taken  
 382 alone, define a trimodal distribution. On these five samples, a Rajput-MacMahon-type statistical  
 383 test confirms the inconsistency of the data set with a unimodal distribution, as  $\chi^2_n = 11$ . Removing  
 384 successive points gave the best results for experiment SM004, with an average half-life of  $117.1 \pm$   
 385  $0.4$  Ga and  $\chi^2_n = 0.14$ . The paradox is that this result is incorrect: pure Type A evaluations only take  
 386 into account repeatability, irrespective of accuracy. However, the Type A approach was not pursued  
 387 by Wilsenach et al. (2017), who instead correctly used a Type B argument. They realized that the  
 388 isolating Delrin holders cause space charge artifacts, and performed later experiments with metallic  
 389 Si holders. This modified design yielded a less precise but much more accurate half-life of  $107.4 \pm$   
 390  $1.9$  Ga.

## 391 392 **Appendix B**

393  
394 Most meteorites are contextless samples of parent planets whose geological history is unknown;  
 395 even for meteorites from Moon, Mars and Vesta, the extraterrestrial bodies about whose evolution  
 396 we have initial understanding, the data fall short of an uncontroversial reconstruction of the local  
 397 processes as would be required for using one of them as a metrologically acceptable standard  
 398 calibrator.

399 The ages of several meteorite classes have been estimated by different short-lived chronometers,  
 400 such as e.g.  $^{26}\text{Al}$ - $^{26}\text{Mg}$ ,  $^{53}\text{Mn}$ - $^{53}\text{Cr}$  and  $^{146}\text{Sm}$ - $^{142}\text{Nd}$ , which give a reasonably consistent chronology  
 401 of the first 5-10 Ma of Early Solar System events. The most useful samples for determining the  
 402  $^{146}\text{Sm}$  half-life are however not the meteorites but rocks younger than 4.56 Ga by at least 0.2-0.3  
 403 Ga, because they provide a greater leverage on the rate of  $^{146}\text{Sm}$  decay. Such samples are lunar  
 404 highlands rocks and terrestrial Early Archean rocks. The post-accretion evolution of both Earth and  
 405 Moon is very far from being known by unanimous consensus, let alone known with such accuracy  
 406 to be useful as a metrological calibrator for the decay constant of  $^{146}\text{Sm}$ . If samples with a relatively  
 407 late formation age are chosen, their prehistory (including the presence of relicts and isotopic  
 408 disequilibrium) and subsequent retrogression become the limiting factor for accuracy.

409 As discussed in the section on the  $^{147}\text{Sm}$  half-life, we require that all internal isochrons respect  $\chi^2_n \leq$   
 410  $1$  in order to be acceptable as half-life calibrators. This is a necessary requirement to ensure that the  
 411 likelihood that an isochron represents the true single-stage evolution of the sample be greater than  
 412  $50\%$ ; if  $\chi^2_n > 1$  this likelihood drops rapidly towards zero.

413 For meteorites there is evidence (the very title of the paper by Sanborn et al. (2015) explicitly  
 414 alludes to it) that the parent planets of the meteorites in our collections had a prolonged history of  
 415 primary and secondary events. These authors propose a secondary Nd isotopic disturbance (Sanborn  
 416 et al., 2015, § 4.1) of the angrite parent planet around 4.51 Ga (Sanborn et al., 2015, p. 91), at a time  
 417 the shorter-lived systems had long become extinct. It is therefore not obvious that the very short-

418 lived systems would record the proposed 4.51 Ga disturbance. Moreover, since neither the  
419 geochemical mobility of parent elements Al, Hf and Mn is a proxy for that of parent element Sm,  
420 nor that of daughter elements Mg, W and Cr faithfully mirrors that of Nd, it is not rigorous to use  
421 the very short-lived radioactive Al-Mg, Hf-W and Mn-Cr systems to deduce the pristineness of the  
422 Sm-Nd pair. The issue is complicated even more by the suggestion by Tissot et al. (2017, p. 610-  
423 615), who list the possibilities of subtle ( $\leq 0.1\%$ ) analytical artifacts in the U-Pb, Al-Mg and Mn-  
424 Cr chronometers in meteorites, implying that contemporaneity of samples and isotopic equilibrium  
425 in the early Solar System may be probable but are not certain. Indeed, all  $^{146}\text{Sm}$ - $^{142}\text{Nd}$  isochrons on  
426 the angrites analyzed by Sanborn et al. (2015, p. 84-88) were obtained after eliminating olivine  
427 and/or plagioclase to achieve  $\chi^2_n$  values  $< 20$  (but still mostly  $> 1$ ).

428 Younger lunar samples were dated by the Carnegie-Livermore collaboration (Carlson et al., 2014;  
429 Marks et al., 2014, and abstracts quoted therein). Their primary interest was estimating the age of  
430 the magma ocean on the Moon (Carlson et al., 2014); the ages obtained in their survey are discussed  
431 by Carlson et al. (2014, p. 11) in the following verbatim quote:

432 *"As with many studies of lunar highland samples, this study finds discordant ages for different isotopic*  
433 *systems when applied to exactly the same mineral and whole rock separates. Choosing which age best*  
434 *represents the crystallization age of the rock depends on applying subjective criteria. In reality, none*  
435 *of the ages need to reflect accurate crystallization ages as the lack of concordance from the different*  
436 *systems is a clear indication of disturbance of the radioisotope systems."*

437 This supports an open-system behavior (like that documented by Sanborn et al., 2015) and raises  
438 questions if the isochrons in the abstracts cited by Marks et al. (2014, Fig. 4) reflect a single-stage  
439 evolution of  $^{142}\text{Nd}$  over time, i.e. only mirror the half-life of  $^{146}\text{Sm}$ . The observation that angritic  
440 achondrite D'Orbigny exhibits

441 *"surprisingly disturbed  $^{147}\text{Sm}$ - $^{143}\text{Nd}$  systematics supported by the unexpectedly young  $^{146}\text{Sm}$ - $^{142}\text{Nd}$*   
442 *"age" of this meteorite in comparison with its otherwise good behavior for most other radiometric*  
443 *systems"* (R.W. Carlson, written communication, 2020)

444 can be viewed as a general question if the entirety of its constituent mineral phases are pristine. If  
445 even in the achondrite D'Orbigny the mineral phases that provide robust radiometric information for  
446 the Mn-Cr system (Glavin et al., 2004) coexist with mineral phases carrying a disturbed Sm-Nd  
447 system, it becomes difficult to argue "beyond reasonable doubt" that such disturbances of the Sm-  
448 Nd system were absent in meteorites whose other radiometric systems are not pristine. In the search  
449 for a "point-like" calibrator for the  $^{146}\text{Sm}$  half-life, it must be explicitly stressed that the default  
450 assumption must never be that a given mineral chronometer behaves ideally. On the contrary, the  
451 burden of proof must lie on the proponent of a calibrator, who needs to actively demonstrate the  
452 degree of ideality of the proposed chronometric system.

453 The open-system behavior documented by Sanborn et al. (2015) manifests itself in a variability of  
454 the subjectively chosen ages for at least two of the four lunar samples analyzed by the Carnegie-  
455 Livermore collaboration. If one uses the  $^{147}\text{Sm}$ - $^{143}\text{Nd}$  age for lunar rock 77215 determined by  
456 Carlson et al. (2014),  $4.42 \pm 0.07$  Ga, instead of that by Marks et al. (2014), 4.29 Ga, the resulting  
457  $^{146}\text{Sm}/^{147}\text{Sm}$  ratio is compatible with the 68 Ma half-life but not the 103 Ma half-life for  $^{146}\text{Sm}$ .  
458 Troctolite 76535 (for which a dozen mutually incompatible ages have been produced over the  
459 years), has a  $^{207}\text{Pb}/^{206}\text{Pb}$  age of 4.38 Ga, the Rb-Sr age with the Villa et al. (2015) half-life of  $^{87}\text{Rb}$   
460 is  $4.33 \pm 0.05$  Ga, and the  $^{146}\text{Sm}$ - $^{142}\text{Nd}$  isochron translates to an age of  $4.38 \pm 0.02$  Ga when  
461 assuming a  $^{146}\text{Sm}$  half-life of 68 Ma.

462 Both the paper by Carlson et al. (2014) and the abstracts cited by Marks et al. (2014) often calculate  
463 internal isochrons that exclude the fine sieve fraction due to its isotopic disequilibrium with the  
464 other minerals. While excluding a particular sieve fraction can produce an approximate age that is  
465 fit for the intended purpose of these papers, namely estimating the age of lunar magma ocean  
466 formation, the fact remains that the sample is in petrological disequilibrium. In this case, it is not  
467 correct to assume that sieving can get rid of all alteration products. Alteration occurs at the sub- $\mu\text{m}$   
468 scale and it is expected that even the coarse grains selected by handpicking do contain some  
469 alteration products, as indeed shown by the overdispersion of the calculated isochrons and  
470 acknowledged in the verbatim quote from the Carlson et al. (2014) paper cited above. This means

471 that none of the lunar rocks in Fig. 4 of Marks et al. (2014) is fit-for-use as a standard reference  
472 material on which to pin the calibration of the  $^{146}\text{Sm}$  half-life.

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### 477 **Figure Captions**

478

479 Figure 1 -  $^{147}\text{Sm}$  half-life counting measurements of the last 60 years, as reported in Table 1. The  
480 red dots represent the measured values, and the vertical blue line segments depict the measured  
481 values plus or minus one associated standard uncertainty. The horizontal dark green line indicates  
482 the consensus value, and the height of the light green rectangle centered on it represents the  
483 consensus value plus or minus one associated standard uncertainty.

484

485 Figure 2 - All  $^{146}\text{Sm}$  half-life counting measurements, as reported in Table 3. The red dots represent  
486 the measured values, and the vertical blue line segments depict the measured values plus or minus  
487 one associated standard uncertainty.