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Multi-ancestry genome-wide association meta-analysis of Parkinson's disease

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Although over 90 independent risk variants have been identified for Parkinson's disease using genome-wide association studies, most studies have been performed in just one population at a time. Here we performed a large-scale multi-ancestry meta-analysis of Parkinson's disease with 49,049 cases, 18,785 proxy cases and 2,458,063 controls including individuals of European, East Asian, Latin American and African ancestry. In a meta-analysis, we identified 78 independent genome-wide significant loci, including 12 potentially novel loci (*MTF2, PIK3CA, ADD1, SYBU, IRS2, USP8, PIGL, FASN, MYLK2, USP25, EP300* and *PPP6R2*) and fine-mapped 6 putative causal variants at 6 known PD loci. By combining our results with publicly available eQTL data, we identified 25 putative risk genes in these novel loci whose expression is associated with PD risk. This work lays the groundwork for future efforts aimed at identifying PD loci in non-European populations.

Parkinson's disease (PD) is a neurodegenerative disease pathologically defined by Lewy body inclusions in the brain and the death of dopaminergic neurons in the midbrain. The identification of genetic risk factors is imperative for mitigating the global burden of PD, one of the fastest growing age-related neurodegenerative diseases. A large PD genome-wide association study (GWAS) meta-analysis uncovered 90 independent genetic risk variants in individuals of European ancestry¹. Similarly, large-scale PD GWAS meta-analyses of East Asian² and a single GWAS of Latin American³ individuals have each identified two risk loci that were not previously identified in Europeans. For PD, there

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Meta-analyses identify 66 known and 12 novel loci

In addition to results from previously described European¹, East Asian² and Latin American³ studies, we also used FinnGen and additional datasets for East Asian, Latin American and African cohorts from 23andMe. Inc (Table 1, Fig. 1 and Supplementary Table 1). In total, we included 49,049 PD cases, 18,618 proxy cases (first-degree relative with PD) and 2,458,063 neurologically-healthy controls. Genetic covariance intercepts from linkage disequilibrium (LD) score regression⁶ within ancestries were close to zero or near the 95% confidence interval, implying that there is no sample overlap between the cohorts (Supplementary Table 1). After the data were harmonized and mapped to genome build hg19, MAMAs were conducted using a random-effects model and meta-regression of multi-ethnic genetic association (MR-MEGA)⁷. The random-effects model had greater power to detect homogenous allelic effects⁷. MR-MEGA uses axes of genetic variation as covariates in its meta-regression analysis and had greater power to detect heterogeneous effects across the different cohorts.MR-MEGA also distinguishes ancestral heterogeneity (differences in effect estimates due to ancestry-level genetic variation) from residual heterogeneity using axes of genetic variation generated from the allele frequencies across the different cohorts.

Combining results from the random-effects model and MR-MEGA, we found 12 novel PD risk loci and 66 hits in known risk loci from single-ancestry GWAS (Table 2, Fig. 2 and Supplementary Tables 2-5) that met the Bonferroni-corrected alpha of 5×10^{-9} , a more stringent threshold chosen to account for the larger number of haplotypes resulting from the ancestrally diverse datasets⁸. Of the 78 risk loci identified, 69 were significant in the random-effects model, whereas 3 were only significant in MR-MEGA. Eight of the novel loci found by the random-effect method showed homogeneous effects across the four different ancestries. An additional novel locus (FASN) identified by the random-effect method showed homogeneous effects in all available populations, but note that this variant failed quality control in both East Asian datasets. The other three loci, identified exclusively in MR-MEGA, showed ancestrally heterogeneous effects. All three loci (IRS2, MYLK2 and USP25) showed evidence of significant ancestral heterogeneity ($P_{ANC-HET} < 0.05$) but no significant residual heterogeneity ($P_{\text{RES} \cdot \text{HET}} > 0.148$), supporting the idea that the signals are due to population structural differences rather than other confounding factors (Fig. 3). For the *IRS2* locus (lead SNP rs1078514, $P_{ANC+HET} = 5.3 \times$ 10⁻³) the Finnish cohort has an opposite effect direction compared to the meta-analysis effect estimate (Supplementary Fig. 4). Similarly, the MYLK2 locus has the African effect estimate most different from the meta-analysis effect estimate (lead SNP rs6060983, $P_{ANC-HET} = 0.035$), suggesting different effects between populations. Although this is a novel single-trait GWAS locus, its lead SNP was previously discovered as a potential pleiotropic locus in a multi-trait conditional/conjunctional false discovery rate (FDR) study between schizophrenia and PD⁹. Lastly, the USP25 locus had the most significant ancestral heterogeneity (lead SNP rs1736020, $P_{ANC-HET} = 4.74 \times 10^{-5}$) and its effects were specific to European and African cohorts, albeit in different directions. When looking at the nearest protein coding gene to each novel lead SNP and their probability of being loss-of-function intolerant (pLI) score, we found that 7 out of 12 genes had a pLI score of 0.99 or 1. Genes with low pLI scores were found both in loci with (MYLK2) and without (SYBU, PIGL and PPP6R2) significant ancestry heterogeneity.

Study	Ancestral population	Cases/proxy/controls
Nalls et al. ¹	European (EUR)	37,688/18,618/1,411,006
Foo et al. ²	East Asian (EAS)	6,724/0/24,851
LARGE-PD 3	Latin American (AMR)	807/0/690
FinnGen Release 4	European-Finnish (EUR)	1,587/0/94,096
23andMe—African	African (AFR)	288/0/193,985
23andMe—East Asian	East Asian (EAS)	322/0/151,905
23andMe—Latino	Latin American (AMR)	1,633/0/581,530
MAMA		49,049/18,618/2,458,063

PESCA v0.3 (ref. 10) was run for the main European and East Asian meta-analyses and all loci identified in the main analysis were explored (Supplementary Table 6). PESCA uses ancestry-matched LD estimates to infer whether the causal variants are population-specific or shared between two populations. Variants identified as shared between the populations may be more likely to be causal. In addition, we expect higher posterior probability (PP) for shared causal variants in the loci identified by MAMA, even if they have not previously been identified in the single-ancestry study. The lead SNP in the RIMS1 locus (rs12528068) had a high PP for being a shared causal variant (PP = 0.972) despite being significant in the European study¹ but not in the East Asian study². We also observed that the novel lead variants for MTF2 (rs35940311), PIK3CA (rs11918587), EP300 (rs4820434) and PPP6R2 (rs60708277) had higher PP estimates for being shared causal variants across both populations (PP_{shared} = 0.757, 0.214, 0.769, 0.946) than for being causal variants in a single population ($PP_{EUR} < 0.080$, PP_{FAS} < 0.001). However, it is important to note that the sample size discrepancy between the European and East Asian data impacts our power to detect population-specific causal variants at any of these loci.

We found 17 suggestive loci that failed to meet our stringent significance threshold but had $P < 5 \times 10^{-8}$ in a fixed-effects meta-analysis and $P < 1 \times 10^{-6}$ in the random-effects meta-analysis (Supplementary Table 4). Fourteen of these regions were novel loci. Two loci near *JAK1* and *HS1BP3* were exclusively found in the 23andMe Latin American and African cohorts. The lead SNPs (rs578139575 and rs73919910) for these loci are non-coding and very rare in European populations but are more common in Africans and Latin Americans (gnomAD v3.1.2 minor allele frequencies in EUR: 0.02%, 0.23%; AFR: 1.64%, 8.84%; AMR: 0.41%, 1.91%). If confirmed, these loci merit further studies in the African and Latin American populations.

Fine-mapping identifies six credible sets with single variants

Fine-mapping was also performed using MR-MEGA, which uses ancestry heterogeneity to increase fine-mapping resolution. We identified 23 loci that had fewer than 5 variants within the 95% credible set. Of these, MR-MEGA nominated a single putative causal variant with >95% PP in 6 loci: *TMEM163, TMEM175, SNCA, CAMK2D, HIP1R* and *LSM7* (Table 3 and Supplementary Tables 7 and 8). Our results affirmed previous results showing the *TMEM175* p.M393T coding variant as the likely causal variant¹¹. The putative variants *HIP1R* have strong evidence for regulome binding (RegulomeDB rank \leq 2). In particular the *HIP1R* variant rs10847864 is located in a transcription start site that is active in substantia nigra tissue (chromatin state windows: chr12:123326400.123327200) and astrocytes in the spinal cord and the brain (chromatin state windows: chr12:123326400.123326600). Outside of the credible sets containing a single variant, we identified missense variants in two genes: *FCGR2A* (p.H167R, PP = 0.145) and *SLC18B1* (p.S30P, PP = 0.780).



Goal: Interpret the meta-analysis results and identify potential targets and biological mechanisms



Fig. 1 | **MAMA study design.** Top panel: four ancestry groups used in the metaanalysis. Middle panel: MAMA and the two methods used. Random-effect (top) is better suited for risk variants with homogeneous effect direction across different ancestries, whereas MR-MEGA (bottom) can identify risk variants with heterogeneous effects due to population stratification introduced by ancestry

differences. The densely dashed lines indicate Bonferroni adjusted suggestive threshold of two-sided $P < 1 \times 10^{-6}$, and the loosely dashed lines indicate Bonferroni adjusted significant threshold of two-sided $P < 5 \times 10^{-9}$. Bottom panel: downstream analyses and their examples. Created with Biorender.com.

Gene set analysis finds enrichment in brain tissues

We used the Functional Mapping and Annotation (FUMA) software^{12,13} to functionally annotate the random-effect results. We generated a custom 1000 Genome reference panel that reflected the ancestry proportions of our dataset and ran multi-marker analysis of genomic annotation (MAGMA)¹⁴ for gene ontology, tissue level and single-cell expression data. We tested 16,992 gene ontology sets in MSigDB v7.0 (ref. 15) and used conditional analysis to discard redundant terms or identify gene sets that must be interpreted together. We found that 40 gene sets were significantly enriched with conditional analysis

identifying 13 gene sets that share their signals with at least one other gene set (Supplementary Table 9). This is a substantial increase from previous 10 gene sets in the European meta-analysis performed by Nalls and colleagues¹. Only two gene ontology terms that were significant in the Nalls et al. meta-analysis were also significant in the multi-ancestry results after multiple test correction: 'curated geneset: Ikeda MIR30 Targets Up' ($P_{\rm FDR} = 0.018$) and 'cellular component: vacuolar membrane' ($P_{\rm FDR} = 0.047$). In addition, ontology terms in immune system pathways (microglial cell proliferation, macrophage proliferation, natural killer T cell differentiation: $P_{\rm FDR} < 0.04$), mitochondria (response to mitochondrial depolarization: $P_{\rm FDR} = 0.028$), vesicles (vesicle uncoating,

rsID	Nearest coding gene	SMR nominated putative genes	CHR:BP:A1:A2	BETA(RE)	SE	(RE)	P(MR-MEGA)	P(ANC-HET)	P(RES-HET)	gnomAD EUR AF	gnomAD EAS AF	gnomAD AMR AF	gnomAD AFR AF	pLI
rs11164870	MTF2	CCDC18	1:93552187:C:G	0.054	0.009	.15×10 ⁻¹⁰	2.64 × 10 ⁻⁹	0.229	0.928	39.0%	35.1%	45.2%	85.0%	-
rs6806917	PIK3CA	KCNMB3	3:178861417:T:C	-0.070	0.011	.65 × 10 ⁻¹⁰	3.43×10 ⁻⁹	0.215	0.762	82.0%	89.9%	77.5%	57.8%	-
rs16843452	ADD1	ADD1, NOP14-AS1, NOP14	4:2849168:T:C	-0.068	0.012 4	1.11 × 10 ⁻⁹	3.19 × 10 ⁻⁷	0.747	0.687	18.5%	47.4%	18.2%	8.9%	0.99
rs6469271	SYBU	SYBU	8:110644774:T:C	-0.056	0.010	1.62 × 10 ^{-°}	2.04×10 ⁻⁷	0.590	0.954	77.5%	59.3%	74.7%	61.5%	0
rs1078514	IRS2	None	13:110463168:T:C	0.068	0.026 4	1.82 × 10 ⁻³	2.30×10 ⁻⁹	5.30×10 ⁻³	0.261	33.3%	39.2%	40.6%	10.7%	0.99
rs28648524	USP8	TRPM7	15:50787409:A:T	0.064	0.010	6.45×10 ⁻¹⁰	2.58 × 10 ⁻⁸	0.406	0.661	78.1%	53.7%	76.5%	79.8%	-
rs11650438	PIGL	ADORA2B, ZSWIM7, PIGL, TTC19, NCOR1, CENPV, TRPV2	17:16234260:A:G	0.050	0.009	93 × 10 ⁻⁹	1.46×10 ⁻⁷	0.528	0.288	46.9%	17.8%	48.5%	64.0%	0
rs4485435	FASN	None	17:80045086:C:G	0.082	0.014 2	61 × 10 ⁻⁹	N/A	N/A	N/A	17.3%	12.1%	34.8%	30.3%	-
rs6060983	MYLK2	None	20:30420924:T:C	0.069	0.037 0	0.0322	3.86×10^{-9}	0.035	0.149	69.3%	99.0%	71.8%	29.0%	0.23
rs1736020	USP25	None	21:16812552:A:C	0.006	0.005 0	.885	1.12×10^{-9}	4.74×10^{-5}	0.638	43.0%	18.6%	38.6%	13.2%	0.75
rs73174657	EP300	ZC3H7B, POLR3H, CSDC2, PMM1, RANGAP1, ME11, L3MBTL2, SLC25A17	22:41434158:A:G	-0.059	0.010	1.81 × 10 ⁻⁹	4.90 × 10 ⁻⁷	0.983	0.655	27.2%	6.3%	47.5%	14.2%	-
rs10775809	PPP6R2	PPP6R2	22:50808017:A:T	0.092	0.015 4	1.09 × 10 ⁻¹⁰	5.61×10^{-8}	0.943	0.903	10.1%	80.3%	80.1%	56.5%	0.16
MR-MEGA could noi effect in log odds ra P(RES-HET): <i>P</i> value	be run for the lea tio; SE, standard € for the two-sided v211 for the near	d SNP of the FASN locus, as it was missing in more than three cohorts: Foo error; P(RE), two-sided P value of association from random effect; P(MR-MF estacial heterepareity test, to fich-squared rear with rf = 3); grom AD fAnces est coding dene (score was unavailable for conom 20 x31 2); SMR summa	et al.², 23andMe East A GA): two-sided P value try] AF, A1 frequency re v-based Mendelian ran	sian and 23anc of association f ported for Euro domization: N/	dMe Latino. I from MR-ME peans (EUR) A. not availa	Vo P values we 3A (chi-square , East Asians (E ble, Bolded ar	re corrected for m d test with df = 4); AS), Amerindians a all significant P v	ultiple tests. CHR, P(ANC-HET), <i>P</i> val (AMR) and African alues (<i>P</i> <5 ×10 ⁻⁹ fo	chromosome; ue for the two-s s (AFR) by gnor r the two-sided	BP, base pair; sided ancestra mAD v3:1.2; pL	A1, effect alle al heterogene .1, probability ests. P<0.051	ele; A2, other a eity test (chi-sc of being loss- for the heteroo	llele; BETA(RE luared test wit of-function int aeneity tests).), allelic h df = 3); olerant
rs10775809 MR-MEGA could noi effect in log odds ra 2(RES-HET): P value score from anomAD	PPP6R2 be run for the les tio; SE, standard € for the two-sided v2:1.1 for the near	PPP6R2 d SNP of the FASN locus, as it was missing in more than three cohorts. Fo error, P(RE), two-sided <i>P</i> value of association from random effect; P(MRMR residual heterogeneity test (chi-squared test with df = 3); gnom AD [Ances rest coding pene [score was unaveilable for gnom AD V3.12); SNR, summa	22:50808017:A:T et al. ² , 23andMe East A GA): two-sided P value try] AF, A1 frequency re v-based Mendelian ran	0.092 sian and 23anc of association f ported for Euro domization; N/	0.015 2 Me Latino. I from MR-ME peans (EUR) A, not availa	I.09 × 10⁻¹⁰ Vo <i>P</i> values we 3A (chi-square , East Asians (E ble. Bolded ar	5.61 × 10 ⁻⁸ re corrected for m d test with df = 4), AS), Amerindians s all significant <i>P</i> v	0.943 Nultiple tests. CHR. P(ANC-HET), <i>P</i> val (AMR) and African alues (<i>P</i> <5 ×10 ⁻⁹ fo	0.900 chrom ue for s (AFR s (AFR	3 nosome; the two-s) by gnor wo-sided	3 10.1% nosome; BP, base pair; the two-sided ancestr) by gnomAD v3.1.2; pl wo-sided association t	3 10.1% 80.3% nosome; BP, base pair; A1, effect alle the two-sided ancestral heterogene by gnomAD v3.12; pLI, probability, we sided association tests. P-O05	3 10.1% 80.3% 80.1% nosome: BP, base pair, A1, effect allele, A2, other a the two-sided an osstral hereinogeneity rest (chi-so) by gnomAD v3.1.2; pL1, probability of being loss- wesided association tests, P <0.05 for the hereino for the hereino.	3 101% 80.3% 80.1% 56.5% nosome: BP, base pair; A1, effect allele; A2, other allele; BETA(RE the two-sided ancestral heterogeneity test (chi-squared test wit b) by gnomAD v31.2; pLI, probability of being loss-of-function int wesided association tests, P-0.05 for the heterogeneity tests).

phagolysosome assembly, regulation of autophagosome maturation: $P_{\text{EDP}} < 0.03$) and tau protein (tau protein kinase activity: $P_{\text{EDP}} = 0.034$) were significant. At the tissue level, the genes of interest were enriched in all brain cell types, as well as pituitary tissue (Supplementary Fig. 9), consistent with the results from Nalls et al.¹.

When analyzing single-cell RNA-sequencing data, there was no expression enrichment across 88 brain cell types in mouse brain data when cross-referenced with DropViz¹⁶ (Supplementary Fig. 10). There was also no enrichment of any specific cell types in the substantia nigra tissue in DropViz (Supplementary Fig. 10). However, in human midbrain data¹⁷, dopaminergic (DA1) and GABAergic (GABA) neurons were enriched (Supplementary Fig. 10).

eOTLs and SMR nominate 25 putative genes near novel loci

We also searched the GTEx v8 (ref. 18) brain tissue eQTLs and multi-ancestry eOTL meta-analysis of the brain¹⁹ to correlate novel loci with gene expression data (Supplementary Tables 10 and 11). To correlate potential putative genes with PD risk, we searched the significant-eQTL genes and genes near the loci with previously completed summary-based Mendelian randomization (SMR)²⁰ results in European-only data. When comparing the SNPs in novel loci with multi-ancestry brain eQTLs¹⁹, 28 genes were significant (Supplementary Fig. 8 and Supplementary Tables 10 and 11). SMR found 25 genes in four novel loci associated with PD risk (Table 2 and Supplementary Table 12). Interestingly, PPP6R2 and CENPV expression changes in substantia nigra were associated with PD risk. PPP6R2 encodes protein phosphatase 6 regulatory subunit 2, a regulatory protein for protein phosphatase 6 catalytic subunit (PPP6C), which is involved in the vesicle-mediated transport pathway. Centromere protein V (CENPV) is involved in centromere formation and cell division.

Discussion

This study is a large-scale GWAS meta-analysis of PD that incorporates multiple diverse ancestry populations. From the joint cohort analysis, we identified 66 independent risk loci near previously known PD risk regions and 12 potentially novel risk loci. Of the putative novel loci, nine had homogeneous effects and three had heterogeneous effects across the different cohorts. We found 17 additional suggestive loci using fixedeffects meta-analysis threshold at $P < 5 \times 10^{-8}$ and random-effects metaanalysis threshold at $P < 1 \times 10^{-6}$. We fine-mapped 23 loci by leveraging the diverse ancestry populations. We highlighted tissues and cell types associated with PD risk, which were consistent with previous findings¹. Finally we used SMR to nominate 25 putative genes near our novel loci.

Novel loci contained genes in pathways previously implicated in PD. The *MTF2* and *PPP6R2* loci contain the genes *TMED5* and *PPP6R2*. Protein TMED5 localizes to Golgi body²¹ and PPP6C, regulated by PPP6R2, is part of the vesicular transport pathways (https://reactome. org/content/detail/R-HSA-199977)²², both of which are implicated in PD pathogenesis²³⁻²⁸. eQTL and SMR analysis showed association between expression changes for PPP6R2 and CENPV in substantia nigra and PD risk. Because substantia nigra deterioration is a hallmark pathogenic feature of PD, PPP6R2 and CENPV merit additional investigation. Within a known locus, a new independent signal was found in RILPL2 (rs28659953). Protein RILPL2 interacts with LRRK2-phosphorylated Rab10 to block primary cilia generation²⁹. Genes JAK1 and HS1BP3 are in two suggestive loci that were found only in Latin American and African populations. JAK1 is one of the proteins in the Janus kinase family, which is a critical part of the JAK-STAT pathway and is implicated in cytokine and inflammatory signaling³⁰. JAK1 variants have been implicated in autoimmune diseases such as juvenile idiopathic arthritis and multiple sclerosis³¹. HS1BP3, also known as essential tremor 2 (ETM2), has been implicated in essential tremor³²⁻³⁴. Based on its sequence, ETM2 may modulate interleukin-2 signaling³⁵. If these loci are confirmed, they would further support the growing appreciation for the role of

Table 2 | Meta-analysis results of lead SNPs in the novel loci



Fig. 2 | **Manhattan plots of the meta-analysis results across 2,525,730 participants. a**, Random-effects model test. **b**, MR-MEGA meta-regression test (chi-squared test with df = 4). The *x* axis shows chromosome and base pair positions of each variant tested in the meta-analyses. The *y* axis shows the two-sided *P* value with no multiple-test correction in the $-\log_{10}$ scale. Orange

horizontal dashed line indicates the Bonferroni-adjusted significant threshold of $P < 5 \times 10^{-9}$. Gray horizontal dashed line indicates the truncation line, where all – $\log_{10} P$ values greater than 40 were truncated to 40 for visual clarity. Novel loci are highlighted in red and annotated with the nearest protein coding gene.

inflammation in PD³⁶. All of the potentially novel PD loci identified in this analysis will require additional replication and functional validation to elucidate their role in PD pathogenesis. Previous findings in European populations found that polygenic risk scores explained 16–36% of PD heritability¹. Although we did not perform similar tests incorporating our novel loci, they may explain additional heritable PD risk.

We found that 26 of the 66 detected known PD loci had nominally significant ancestral heterogeneity ($P_{ANC-HET} < 0.05$) and 10 remained significant after Bonferroni correction ($P_{ANC-HET} < 0.05/62$ MR-MEGA loci) (Fig. 3 and Supplementary Table 3). This heterogeneity may be caused by differences in effect sizes and allele frequencies between the different populations and thus should be studied as loci with potentially ancestrally divergent risk. 18 of the previous 92 known loci from single-ancestry GWASs did not overlap with any genome-wide significant loci in the multi-ancestry results at the significance threshold of 5×10^{-9} (Supplementary Table 13). However, our results do not necessarily invalidate these previous results. First, several of the cohorts have small sample sizes, which may increase the influence of sampling variation. Another reason may be due to the stringent genome-wide significance threshold of 5×10^{-9} . Although this is a large PD GWAS meta-analysis, the more stringent significance threshold further raises

the sample size needed to achieve equivalent statistical power. Of the 17 European loci identified, 3 were significant at the 5×10^{-8} threshold, and all 17 loci were at least nominally significant with the MR-MEGA method ($P_{MR-MEGA} < 5 \times 10^{-6}$). Lastly, variants may be more specific to the population in which they were first identified. 5 of the 17 variants had nominal ancestral heterogeneity ($P_{ANC-HET} < 0.05$). It is worth noting that there are large differences in statistical power across ancestries. Additional population-specific loci will likely reach significance when larger sample sizes are available for non-European datasets.

Our fine-mapping isolated several putative causal variants in previously discovered loci. *TMEM175*-rs34311866 has been previously identified as functionally relevant to PD risk³⁷, which is consistent with our fine-mapping results. Fine-mapped variants in *TMEM163*, *HIP1R* and *CAMK3D* were also found to be parts of active or strong transcription sites in substantia nigra tissues. Among the fine-mapped variants were two missense variants in *FCGR2A* and *SLC18B1*, albeit with a lower PP than the 7 singular putative variants that we highlighted in Table 3. *FCGR2A* is present in multiple immune-related ontology gene sets, further highlighting the potential role of the immune system in PD pathology. However, the function of *SLC18B1* is still unknown. Although the fine-mapping results provided by MR-MEGA are sufficient



Fig. 3 | **Heterogeneity upset plots. a**, Top variants per novel loci. **b**, Top variants per MR-MEGA identified locus with moderate to high heterogeneity ($l^2 > 30$). The top bar plot illustrates heterogeneity with dark blue indicating ancestry heterogeneity proportion and light blue indicating other residual heterogeneity proportion. The bottom plot shows the subcohort level beta values with blue

indicating positive and red indicating negative effect directions. Three variants with greater than 30% /² total heterogeneity were only identified in the MR-MEGA meta-analysis method, whereas little to no heterogeneity is observed in loci identified in random effect.

to identify putative causal variants for loci driven by one independent signal, multiple variants in a locus can contribute to complex traits. The additive and epistatic effects of multiple causal variants in a locus can be difficult to interpret when the effects associated with each independent signal are small. The gene ontology analysis found multiple pathways that may be relevant to PD pathology (Supplementary Table 9), including those related to mitochondria (response to mitochondrial depolarization) vesicles (vesicle uncoating, phagolysosome assembly, regulation of autophagosome maturation) tau protein (tau protein kinase activity)

Locus	Number of significant SNPs	Nominated variant	CHR:BP:A1:A2	Nearest gene	Known PD gene ±1 MB	Functional consequence	CADD	RDB
11	6	rs57891859	2:135464616:A:G	TMEM163	TMEM163	intronic	6.746	4
19	926	rs34311866	4:951947:C:T	TMEM175	TMEM175	exonic	11.09	NA
23	1483	rs356182	4:90626111:A:G	SNCA	SNCA	ncRNA intronic	8.962	NA
24	121	rs13117519	4:114369065:T:C	CAMK2D	CAMK2D	intergenic	1.216	За
45	1371	rs10847864	12:123326598:G:T	HIP1R	HIP1R	intronic	2.403	2b
60	1	rs55818311	19:2341047:C:T	SPPL2B	LSM7	ncRNA exonic	1.096	5

Table 3 | MR-MEGA fine-mapping results for loci with a single SNP within the 95% credible set

Known PD genes are either known PD risk genes (SNCA and TMEM175) or genes with the highest score in the nearest known PD locus by the PD GWAS Locus Browser³⁷. CHR, chromosome; BP, base pair; A1, effect allele; A2, other allele; CADD, combined annotation-dependent depletion score; RDB, regulomeDB score; ncRNA, non-coding RNA.

and immune cells (microglial cell/macrophage proliferation, and natural killer T cell differentiation)³⁶. Neither mitochondrial nor immune cell pathways were significant in the previous European-only meta-analysis. Novel signals from the multi-ancestry approach may have given enough power to highlight these ontology terms. Out of 10 ontology terms that were significant in the previous European-only meta-analysis¹, 4 terms were not tested due to version differences in MSigDB and only 2 of the remaining terms were significant. However, the other 4 terms were still nominally significant at P < 0.05. This may be due to genome-wide signals that were less significant due to their heterogeneity across the different populations.

Although this is a large multi-ancestry PD meta-analysis GWAS, the European population is still overrepresented. Around 80% of full PD cases are of European descent. Individuals of African descent were particularly underrepresented at just 0.5% of the effective PD cases. The discoveries in our study warrant future efforts to expand studies in more diverse populations. The Global Parkinson's Genetics Program (GP2) is partnering with institutions that care for underrepresented populations to generate data for these underserved communities all over the world⁵, and we will continue the ongoing analysis as more participants are genotyped. Just as the first PD GWASs failed to identify significant signals^{38,39}, we are confident that future diverse ancestry GWAS will produce impactful association results as sample sizes increase. Further efforts in multi-ancestry and non-European GWAS will identify loci that are more relevant to the global population and will continue to facilitate finemapping efforts to identify the genetic variants that drive these associations.

Online content

Any methods, additional references, Nature Portfolio reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information, details of author contributions and competing interests and statements of data and code availability are available at https://doi.org/10.1038/s41588-023-01584-8.

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Methods

Study design and cohort descriptions

We used a single joint meta-analysis study design to maximize statistical power⁴⁰. We used datasets representing four different ancestry groups: European, East Asian, Latin American and African. The meta-analysis included 49,049 PD cases, 18,618 PD proxy cases (participant with a parent with PD) and 2,458,063 neurologically normal controls (Table 1 and Supplementary Table 1). GWAS results of European¹, East Asian² and Latin American³ populations were previously reported. African dataset as well as the additional Latin American and East Asian PD GWAS summary statistics were provided by 23andMe. The Finnish PD GWAS summary statistics was acquired from FinnGen Release 4 (G6 PARKINSON EXMORE). For the FinnGen data, we chose the endpoint 'Parkinson's disease (more controls excluded)' (G6 PAR-KINSON EXMORE), which excludes control participants with psychiatric diseases or neurological diseases. Although some FinnGen GWAS results also include UK Biobank participants, our FinnGen data did not include any UK Biobank participants.

23andMe diverse ancestry data

All self-reported PD cases and controls from 23andMe provided informed consent and participated in the research online, under a protocol approved by the external AAHRPP-accredited institutional review board (IRB), Ethical & Independent Review Services (E&I Review). Participants were included in the analysis on the basis of consent status as checked at the time data analyses were initiated. The name of the IRB at the time of the approval was Ethical & Independent Review Services. Ethical & Independent Review Services was recently acquired, and its new name as of July 2022 is Salus IRB (https://www. versiticlinicaltrials.org/salusirb). Samples were genotyped on one of five genotyping platforms: V1 and V2, which are variants of Illumina HumanHap550+ BeadChip; V3, Illumina OmniExpress+ BeadChip; V4, Illumina custom array that includes SNPs overlapping V2 and V3 chips; or V5, Illumina Infinium Global Screening Array. For inclusion, samples needed a minimal call rate of 98.5%. Genotyped samples were then phased using either Finch or Eagle2 (ref. 41) (RRID:SCR 015991) and imputed using Minimac3 (RRID:SCR_009292) and a reference panel of 1000 Genomes Phase III⁴² (GRCh38) and UK10K data⁴³. For this study, samples were classified as African, East Asian or Latino using a genotype-based pipeline⁴⁴ consisting of a support vector machine and a hidden Markov model, followed by a logistic classifier to differentiate Latinos from African Americans. Unrelated individuals were included in the analysis, as determined via identity-by-descent (IBD). Variants were tested for association with PD status using logistic regression, adjusting for age, sex, the first five principal components and genotyping platform. Reported P values were from a likelihood ratio test.

MAMA

We performed MAMA of GWAS results using MR-MEGA v0.2 (ref. 7) and PLINK 1.9 (RRID:SCR_001757). MR-MEGA performs a meta-regression by generating axes of genetic variation for each cohort, which are then used as covariates in the meta-analysis to account for differences in population structure. Although MR-MEGA was able to generate four principal components as axes of genetic variation, three principal components visibly separated the super population ancestries and explained 98% of the population variance (Supplementary Fig. 7). Therefore, we used three principal components to minimize overfitting. MR-MEGA has reduced power to detect associations for variants with homogeneous effects across populations. It is therefore recommended to run MR-MEGA alongside another meta-analysis method. PLINK 1.9 was used to perform random-effect meta-analysis to detect homogenous allelic effects.

Before the analysis, all datasets were harmonized to genome build hg19 using CrossMap⁴⁵ (RRID:SCR_001173) and Python 3.7. All variants were filtered by imputation score ($r^2 > 0.3$) and minor allele frequency ≥ 0.001 . Only autosomal variants were kept in the final results as sex-chromosome data were not available for all ancestries. In total 20,590,839 variants met the inclusion criteria. However, MR-MEGA has a cohort-number requirement that varies based on the number of axes of variation. Therefore, 5,662,641 SNPs present in at least 6 of the 7 cohorts were analyzed in the MR-MEGA analysis. Bonferroni-adjusted alpha was set to a more stringent 5×10^{-9} for all MAMAs to account for the larger number of haplotypes resulting from the ancestrally diverse datasets⁸. Genomic inflations were measured for all cohorts and the meta-analysis. Inflation for cohorts with large discrepancy between the case and control numbers was normalized to 1,000 cases and 1,000 controls. All inflation was nominal and below 1.02 (Supplementary Figs. 1–3 and Supplementary Table 1). No genomic control was applied prior to meta-analysis.

We identified genomic risk loci within our meta-analysis results using Functional Mapping and Annotation (FUMA) v1.3.8 (refs. 11,12). In brief, FUMA first identifies independent significant SNPs in the GWAS results by clumping all significant variants with the r^2 threshold <0.6, and then a locus is defined by merging LD blocks of all independent significant SNPs within 250 kb of each other. Start and end of a locus is defined by identifying SNPs in LD with the independent significant SNPs ($r^2 \ge 0.6$) and defining a region that encompasses all SNPs within the locus. Lead SNPs within a locus are determined by further clumping the independent significant variants within the genomic locus ($r^2 \ge 0.1$). The 1000 Genome reference panel with all ancestries was used to calculate the r^2 .

To determine if any associated loci in the meta-analysis were not previously identified, all significant SNPs were compared to the 92 known PD risk variants found in the previous two major meta-analyses^{1,2}. Two variants identified in the Latin American admixture population³ could not be replicated, as the variants and their proxies were removed during quality control. If a genomic risk locus contained a significant hit in either population within 250 kb, then the locus was considered a known hit. Otherwise the locus was considered a novel hit. Forest plots and QQ plots were generated using python 3.7 with seaborn v0.11.2 and matplotlib v3.5.1. Manhattan plots were generated using gwaslab v3.3.11.

Fine-mapping

Fine-mapping was performed using MR-MEGA⁷, which approximates a single-SNP Bayes factor in favor of association. This is reported as the natural log of Bayes factor (lnBF) per SNP in the MR-MEGA meta-analysis summary statistics. SNPs were selected at meta-GWAS significance level ($P < 5 \times 10^{-9}$). PPs of driving the association signal at each locus were calculated from the Bayes factor as follows:

$$\pi_j = \frac{\Lambda_j}{\sum_{j=1}^n \Lambda_j},$$

where Λ_j is the Bayes factor of the *j*th SNP within a locus with *n* number of SNPs. Credible sets of fewer than 5 SNPs with sum PP (π_j) greater than 0.95 were accepted as putative causal variants. We excluded results located in the major histocompatibility complex region and the MAPT locus due to their complex LD structure.

Estimation of population-specific or shared causal variants at associated loci

Proportion of population-specific and shared causal variants (PESCA v0.3)¹⁰ was used to estimate whether causal variants at the loci identified in the meta-analysis were population-specific or shared between two populations. In brief, genome-wide heritability was estimated for the European and East Asian GWAS summary statistics using LD score regression^{6,46}. Summary statistics of both populations were intersected with common variants with the 1000 Genome reference panels provided by PESCA, which have already been LD pruned ($R^2 > 0.95$) and

low-frequency SNPs removed (minor allele frequency < 0.05). The intersected variants were further split according to independent LD regions from the European and East Asian populations. The genome-wide prior probabilities of population-specific and shared causal variants were calculated using default parameters or as otherwise recommended by PESCA; then the results were used to calculate the PP for each variant. When the lead SNP was unavailable in the results, proxy variants (R^2 > 0.8) were used to approximate the PP for each variant for East Asian and European ancestry using R 4.2.0 and LDlinkR v1.1.2 (ref. 47). Other cohorts were not included due to sample size constraints for this method.

Functional annotation and GSEA

Functional annotation of the discovery results utilizing publicly available annotation data was done using FUMA v1.3.8 (refs. 11,12). The summary statistics were annotated by ANNOVAR⁴⁸ (RRID:SCR_012821) through the FUMA platform. Our meta-analysis results were analyzed using MAGMA¹³ (RRID:SCR_001757) to check for enrichment in gene ontology terms and gene expression data from tissues in GTEx v8 (ref. 18). We tested 16,992 gene sets and gene ontology terms from MSigDB v7 (ref. 15) as well as single-cell RNA-sequencing expression data from mouse brain samples in DropViz¹⁶ and human ventral midbrain samples¹⁷. Test parameters were set to default. MAGMA gene analysis was run with a custom 1000 Genome reference panel that had a similar proportion of European, East Asian, Latin American and African participants as our main analysis. In short, we added all European participants and randomly selected participants from the East Asian, Latin American and African populations until the ancestry proportions of the reference panel were matching the effective sample size proportions of our study. The MAGMA gene analysis results were then analyzed using gene set analysis for ontology terms and gene-property analysis for tissue specificity. Results were adjusted for multiple tests using Benjamini-Hochberg FDR correction with the alpha of 0.05. The significant ontology terms were analyzed again in conditional analyses to identify and filter terms that share the same signals. Conditional analyses rerun the analyses with significant ontology terms as additional covariates. This can identify terms that lose significance when 'conditioned' on another, which may mean the terms share an underlying signal. When a term lost significance while the paired term retained nominal significance, the term that was no longer significant was discarded. When both terms lost significance, both were retained but highlighted with the comment that the pairs need to be interpreted together. Tissue level enrichment analysis was done using the pre-processed GTEx gene expression dataset provided by FUMA investigators. Single-cell expression enrichment analyses were performed by uploading the MAGMA gene analysis results to the FUMA cell-type analysis tool, which runs the MAGMA gene-property analysis with the chosen RNA-sequencing data. Additional pathway analyses of genes mapped by FUMA SNP2GENE were performed through GENE2FUNC with default parameters.

SNPs in the novel loci were searched in multi-ancestry brain eQTL meta-analysis results¹⁹ (under Synapse ID syn23204884). We used a *P*-value cutoff of 10⁻⁶ as previously described¹⁹. eQTL and GWAS comparison plots were generated using LocusCompareR⁴⁹. Multi-SNP SMR was used to test if DNA methylation and/or RNA expression of genes near the novel loci were associated with PD risk²⁰. The nearest genes from the lead SNPs, significant genes in MAMA brain eQTL results and significant genes in GTEx v8 brain tissue were chosen for SMR. In total, 44 genes near the novel loci were searched in a list of previously completed PD SMR results from European-only GWAS meta- analysis (https://www.ukbiobank.ac.uk/learnmore-about-uk-biobank/news/nightingale-health-and-uk-biobankannounces-major-initiative-to-analyse-half-a-million-bloodsamples-to-facilitate-global-medical-research)18,20,50-56. Only tissues in the central nervous system, digestive system and blood were used due to their relevance to PD pathology. Methylation

probes were annotated using the Bioconductor R package IlluminaHumanMethylation450kanno.ilmn12.hg19 v0.6.0 (https:// bioconductor.org/packages/release/data/annotation/html/ IlluminaHumanMethylation450kanno.ilmn12.hg19.html). The association signals were adjusted using FDR correction with the alpha of 0.05 and all signals with $P_{\text{HEDI}} < 0.05$ were removed due to heterogeneity.

Reporting summary

Further information on research design is available in the Nature Portfolio Reporting Summary linked to this article.

Data availability

GWAS summary statistics for Foo et al.² and Loesch et al.³ are available upon request to the respective authors. The UKBB genotype and phenotype data are available through the UKBB web portal https://www. ukbiobank.ac.uk/. FinnGen summary statistics are available through the FinnGen website https://www.finngen.fi/. GWAS summary statistics for 23andMe datasets (post-Chang and data included in Chang et al.⁵⁷ and Nalls et al.58) will be made available through 23 and Me to qualified researchers under an agreement with 23andMe that protects the privacy of the 23andMe participants. Please visit research.23andme.com/ collaborate/#publication for more information and to apply to access the data. An immediately accessible version of the multi-ancestry summary statistics is available on the Neurodegenerative Disease knowledge Portal (https://ndkp.hugeamp.org/) excluding Nalls et al.58, 23andMe post-Chang et al.57 and Web-Based Study of Parkinson's Disease (PDWBS) but including all analyzed SNPs. Same summary statistics are also available at AMP-PD (https://amp-pd.org/) under GP2 Tier 1 access and GWAS Catalog (https://www.ebi.ac.uk/gwas/) under accession code GCST90275127 (http://ftp.ebi.ac.uk/pub/databases/gwas/summary_statistics/GCST90275001-GCST90276000/ GCST90275127/). After applying with 23 and Me, the full summary statistics including all analyzed SNPs and samples in this GWAS meta-analysis will be accessible to the approved researcher(s). MSigDb is available at http://software.broadinstitute.org/gsea/msigdb/.GTEx is available at https://gtexportal.org/home/. Multi-ancestry brain eQTL data from Zeng et al.¹⁹ are available at https://hoffmg01.hpc.mssm.edu/brema/. eQTL/mQTL/caQTL data used for SMR outside of MetaBrain⁵⁰ and eQTLGen⁵² are available at https://yanglab.westlake.edu.cn/software/ smr/#DataResource. MetaBrain eQTL data are available at https://www. metabrain.nl/. eOTLGen data are available at https://www.eqtlgen. org/. pQTL data from Wingo et al.⁵⁴ are available upon request to the respective author. UK Biobank-Nightingale metabolomic data used for SMR are available at https://gwas.mrcieu.ac.uk/.

Code availability

The analysis pipeline code is available on GP2 github: (https://github. com/GP2code/GP2-Multiancestry-metaGWAS) and deposited on Zenodo (https://doi.org/10.5281/zenodo.8045547)⁵⁹.

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Author contributions

A.B.S., C.B., M.A.N., I.M. and J.N.F. conceived the project. C.B., M.A.N., I.M. and J.N.F. designed and supervised the project. K.H., J.N.F. and I.M. provided data. J.J.K., D.V., D.V.-O. and M.M.L. performed the experiment. J.L. and C.W.S. assisted with data visualization. H.I., H.L., M.B.M., E.-K.T., S.B.-C. and A.J.N. advised on the project. J.J.K. wrote the manuscript with input from all authors.

Competing interests

K.H. and members of the 23andMe Research Team are employed by and hold stock or stock options in 23andMe. M.A.N.'s participation in this project was part of a competitive contract awarded to Data Tecnica International by the NIH to support open science research; he also currently serves on the scientific advisory board for Clover Therapeutics and is an advisor to Neuron23. A.J.N. reports consultancy and personal fees from AstraZeneca, AbbVie, Profile, Roche, Biogen, UCB, Bial, Charco Neurotech, uMedeor, Alchemab and Britannia outside the submitted work. The other authors declare no competing interests.

Additional information

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Correspondence and requests for materials should be addressed to Jonggeol Jeffrey Kim, Cornelis Blauwendraat, Mike A. Nalls, Jia Nee Foo or Ignacio Mata.

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Reporting Summary

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		Our web collection on <u>statistics for biologists</u> contains articles on many of the points above.

Software and code

Policy information about availability of computer code

Data collection	On No software was used.				
Data analysis	Data harmonization was done on Python 3.7 and CrossMap. Meta-analyses were done in PLINK 1.9 and Meta-Regression of Multi-Ethnic Genetic Association (MR-MEGA). Putative burden analysis was done using population-specific/shared causal variants (PESCA). Gene-ontology and tissue enrichment tests were done in Functional Mapping and Annotation (FUMA) and Multi-marker Analysis of GenoMic Annotation (MAGMA). Variant level annotation was done using ANNOVAR through FUMA. Plots and other miscellaneous analyses were done in Python or R. Analysis scripts are available on Github: https://github.com/GP2code/GP2-Multiancestry-metaGWAS and deposited deposited on Zenodo under doi:10.5281/zenodo.8045547.				
	Programs and their respective versions: PLINK v1.9 MR-MEGA v0.2 FUMA v1.3.8 MAGMA v1.08 ANNOVAR-last updated on Dec 5 2016 PESCA v0.3 Python 3.7 R 4.2.0				
	R packages used: LDlinkR v1.1.217				

Python packages: gwaslab v3.3.11 seaborn v0.11.2 matplotlib v3.5.1

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GWAS summary statistics for Foo et al. 2020 and Loesch et al. 2020 are available upon request to the respective authors. The UKBB genotype and phenotype data are available through the UKBB web portal https://www.ukbiobank.ac.uk/. FinnGen summary statistics are available through the FinnGen website https:// www.finngen.fi/. GWAS summary statistics for 23andMe datasets (post-Chang and data included in Chang et al. 2017 and Nalls et al. 2014) will be made available through 23andMe to qualified researchers under an agreement with 23andMe that protects the privacy of the 23andMe participants. Please visit research.23andme.com/collaborate/#publication for more information and to apply to access the data. An immediately accessible version of the multi-ancestry summary statistics is available on the Neurodegenerative Disease knowledge Portal (https://ndkp.hugeamp.org/) excluding Nalls et al. 2014, 23andMe post-Chang et al. 2017 and Web-Based Study of Parkinson's Disease (PDWBS) but including all analyzed SNPs. Same summary statistics are also available at AMP-PD (https:// amp-pd.org/) under GP2 Tier 1 access and GWAS Catalog under accession code GCST90275127 (http://ftp.ebi.ac.uk/pub/databases/gwas/summary_statistics/ GCST90276000/GCST90275127/). After applying with 23andMe, the full summary statistics including all analyzed SNPs and samples in this GWAS meta-analysis will be accessible to the approved researcher(s). MSigDb is available at http://software.broadinstitue.org/gsea/msigdb/. GTEx is available at https:// gtexportal.org/home/. Multi-ancestry brain eQTL data from Zeng et al. 2020 are available at https://hoffmg01.hpc.mssm.edu/brema/. eQTL/mQTL/caQTL data used for SMR outside of MetaBrain and eQTLGen are available at https://yanglab.westlake.edu.cn/software/smr/#DataResource. MetaBrain eQTL data are available at https://gwas.mrcieu.ac.uk/.

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Policy information about studies with <u>human participants or human data</u>. See also policy information about <u>sex, gender (identity/presentation)</u>, <u>and sexual orientation</u> and <u>race, ethnicity and racism</u>.

Reporting on sex and gender	All data used in this analysis used biological sex as a covariate in their respective studies.				
Reporting on race, ethnicity, or other socially relevant groupings	We used data from four ancestrally-distinct populations: European, East Asian, African, and Latin American. Ancestry was defined as continent of ancestral origin defined by an individual's genotype information. We did not use any socially constructed variable such as race and ethnicity as a proxy for ancestry. As we used summary-level data as-is, we did not make specific ancestry determinations as the respective studies made the determinations using their individual-level data. For 23andMe data, ancestry was determined using a genotype-based pipeline consisting of a support vector machine and a hidden Markov model, followed by a logistic classifier to differentiate Latinos from African Americans.				
Population characteristics	Data included participants with Parkinson's disease and control participants. As we used data as-is, we do not describe any additional population characteristics and they can be found in their respective manuscripts				
Recruitment	We did not recruit participants for this study. However, data from 23andMe participants were collected through self-report which may influence the results via self-report bias.				
Ethics oversight	All self-reported PD cases and controls from 23andMe provided informed consent and participated in the research online, under a protocol approved by the external AAHRPP-accredited IRB, Ethical & Independent Review Services (E&I Review). Participants were included in the analysis on the basis of consent status as checked at the time data analyses were initiated. The name of the IRB at the time of the approval was Ethical & Independent Review Services. Ethical & Independent Review Services was recently acquired, and its new name as of July 2022 is Salus IRB (https://www.versiticlinicaltrials.org/salusirb).				

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Sample size	We collected datasets from as many datasets as available. As a meta-analysis of 7 different studies/analyses, this work is the largest Parkinson's disease Genome-Wide meta-analysis to-date with 49049 cases, 18618 proxy-cases, and 2458063 controls.
Data exclusions	For 23andMe datasets, only data that met the quality control criteria and were unrelated were included in the analysis. For inclusion, samples needed a minimal call rate of 98.5%. Only data from unrelated participants were used to minimize bias from relatedness. We used data from other studies as-is, but each respective studies performed their own quality-control procedures, including removing related participants and low genotype quality samples.
Replication	We used a single joint meta-analysis study design to maximize statistical power. No replication samples were available as all available datasets were used in the meta-analysis.
Randomization	The experimental groups were divided into Parkinson's disease cases and controls. All data used were adjusted for age, sex, population structure-principal components to account for population stratification.
Blinding	Blinding is not relevant to study as this is a meta-analysis of observational genetic studies and not a randomized experiment.

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